Political Discretion and Antitrust Policy: Evidence from the Assassination of President McKinley

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Abstract: We study the importance of discretion in antitrust enforcement by analyzing the response of asset prices to the assassination of President McKinley in 1901. During his term in office the largest wave of merger activity in American history occurred, and McKinley did not attempt to enforce antitrust laws against those firms. By contrast, his vice president, Theodore Roosevelt, was known to be interested in controlling anticompetitive behavior. We find that firms with greater vulnerability to antitrust enforcement saw greater declines in their abnormal returns following McKinley's assassination. Roosevelt's accession caused one of the most significant changes in antitrust of the Gilded Age—not from new legislation, but from a change in the approach taken to the enforcement of existing law. Our results highlight the importance of discretion in antitrust enforcement.

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Over the past 40 years, the level of concentration within many sectors of the American economy has increased substantially. This has been associated with a decline in competition and labor's share of income, and a slowdown in aggregate output (Barkai, 2016; Autor et al., 2017; De Loecker and Eeckhout, 2017; Grullon et al., 2017). The failure of antitrust authorities to restrain the rise in concentration has provoked concerns that existing antitrust statutes may no longer offer regulators adequate tools for policing anticompetitive behavior. Yet government agencies often hold significant discretion over regulatory enforcement, and some have argued that stronger enforcement of existing statutes could have gone a long way towards reigning in anticompetitive forces (Baker et al., 2018). Disentangling the effects of the substance of existing statutes from the efforts exerted to enforce them is quite difficult, in no small part because enforcement efforts are typically not easy to measure or even observe. We exploit an extraordinary episode from the Gilded Age in which the enforcement of antitrust statutes was suddenly strengthened, and show that the discretion of top political leaders over antitrust enforcement can have meaningful consequences for the economy.

There is arguably no period in American history more important for understanding the forces of economic concentration and their interaction with the political process than the Gilded Age. From 1895 to 1904, the United States witnessed its largest-ever wave of business consolidations, as thousands of competing firms in many industries merged to form giant 'trusts' (Nelson, 1959; Lamoreaux, 1985). William McKinley, who was elected president in 1896, was generally friendly towards business interests, and did not attempt to use the Sherman Antitrust Act of 1890 to challenge any of those mergers. His assassination by an anarchist in September 1901 presents a unique opportunity to study the effects of a change in the president's attitude towards enforcement of antitrust laws at a time when other institutions remained unchanged. In contrast to McKinley, Theodore Roosevelt, who succeeded him as president, had been openly critical of big business. The sudden accession of a well-known Progressive reformer to the presidency likely shifted expectations regarding the aggressiveness with which antitrust laws would be enforced.

We use the stock market's reaction to the McKinley assassination to measure the expected impact of this change in preferences over antitrust enforcement. The quasi-random nature of the assassination enables us to estimate the market's reaction in a way that election outcomes, which were generally well anticipated, do not (see Rhode and Strumpf, 2004). The assassination did not coincide with any changes to the composition of Congress or the courts, and even the attorney general remained unchanged. But a president who wanted vigorous antitrust unexpectedly replaced one who had not. In response to the shooting of McKinley, the value of NYSE-traded firms fell by an average of 6.2 percent. To put this magnitude in perspective, the stock market declined by only 1.6 percent on average over the six other presidential assassinations and nearly successful assassination attempts we have identified. Importantly, the change in aggressiveness with which antitrust laws were expected to be enforced meant that firms that had engaged in mergers prior to the assassination were more likely to be vulnerable. We find that following McKinley's shooting, firms involved

¹Though the recent rise in concentration is well documented, there is an active ongoing debate on the degree to which market power has changed in recent decades. See, among others, De Loecker and Eeckhout (2017), Gutiérrez and Philippon (2017a), Traina (2018), and Hall (2018).

²For example, an influential piece in the *New York Times* argued that "The century-old antitrust laws do not seem up to the task. Perhaps it is time to create laws for a new Gilded Age and provide regulators the power to determine if megadeals are truly good for America" ("Changing Old Antitrust Thinking for a New Gilded Age," 22 July 2014.) A 2016 report by The Roosevelt Institute, a think tank, also called for new antitrust laws (Abernathy et al., 2016).

in recent mergers saw declines in their abnormal returns that were 1.5 to 2 percentage points greater than those of other firms. These results suggest that investors expected a change in antitrust enforcement, and that they anticipated that these new policies would have meaningful impacts on affected firms.

A possible source of concern regarding our analysis is that the effects of the change from McKinley to Roosevelt may have been confounded with the effects of a presidential assassination. The fact that an anarchist shot the president, for example, may have been perceived as a sign of rising political instability. Yet the experience with McKinley offers a unique opportunity to address this concern. President McKinley initially survived the shooting, and three days later his doctors announced that they expected him to make a "full recovery." When that prognosis was announced, the losses experienced following his shooting were largely reversed. Then, seven days following the shooting, when it was suddenly announced that McKinley was in fact near death, the market reversed again. Since the effects from political unrest should have been reflected in prices on the day of the shooting, this latter decline in stock prices suggests that investors instead reacted to expected policy changes that would result from Roosevelt becoming president.

An alternative interpretation of our results could be that investors expected a change in the law, rather than a change in enforcement. The historical context of our event suggests that this was unlikely. At the time of the assassination, there was relatively little support in Congress for new antitrust legislation. Between 1881 and the assassination of McKinley, in fact, 45 separate pieces of antitrust legislation had been proposed in Congress, most in the 1890s (Mitchell, 2007: 122). Nearly all of them had been defeated in Congress, usually stalled in committee. Roosevelt's accession to the presidency did not change the composition of Congress, which ultimately controlled the fate of any new legislation. Indeed Roosevelt repeatedly sought new antitrust legislation during his presidency, and was met with little success (Crane, 2008).

An additional source of concern could be that the transition from McKinley to Roosevelt may have been regarded as harmful to particular firms for reasons unrelated to antitrust enforcement. For example, Roosevelt may have been perceived as less friendly to corporations affiliated with major donors to McKinley's campaign. It is also possible that Roosevelt's policy agenda may have been expected to differ from McKinley's on issues beyond antitrust enforcement. Yet our results are robust to controlling for affiliations with McKinley's donors, and we show that the two most plausible policy differences between Roosevelt and McKinley—on tariffs and labor relations—are unlikely to be the main drivers of our results. Our results are also robust to the inclusion of controls for a broad range of other firm characteristics, and to the use of non-parametric event study tests, as in Dube et al. (2011).

Once in office, Roosevelt began to purse a more aggressive approach to antitrust. To further validate our interpretation of the market's reaction to the assassination as being driven by expectations of stricter antitrust enforcement, we use an event study methodology to analyze the stock market's response to the initiation of Roosevelt's first antitrust suit. This suit argued that a recent merger of competing railroads, which had been combined by J.P. Morgan into an enormous holding company known as Northern Securities, violated the Sherman Act. By attempting to apply the Sherman Act in this way for the first time, the suit signaled a dramatic shift in antitrust enforcement. Plans for this suit were kept secret, which enables us to observe the market's assessment of the expected change in antitrust doctrine in response to its announcement. Our analysis indicates that the firms whose shares performed worse in response to bad news regarding McKinley's

health also suffered differentially low abnormal returns following the announcement of the suit. This provides additional evidence that concerns related to antitrust enforcement, rather than other expected changes in policy or personnel, were likely to have been responsible for the fluctuations in firm values following the assassination.

The analysis of stock market returns only allows us to study the effects expected by investors over a short horizon. To provide insights into the longer-run effects of stricter enforcement of antitrust statutes, we construct a panel of accounting data from 1895 to 1905, and investigate whether the firms we designated as most likely to be affected by Roosevelt's policy changes were in fact differentially impacted in the years after Roosevelt's accession. Although noisy, our results rule out a significant contraction in investment rates, and in fact are suggestive of an increase. Thus, using a very different setting and strategy, our findings are consistent with (Gutiérrez and Philippon, 2017b), who find a positive relationship between competition and corporate investment in recent decades. We also find that firms that were more likely to have been affected by more aggressive antitrust enforcement saw a relative in profitability that was equivalent to 17 percent of the mean profit rate for 1900. Although we cannot use this difference-in-difference approach to assess the welfare impact of the change in policies, our findings at least suggest that stronger enforcement of antitrust laws may have had strong and persistent redistributive consequences.

Recent years have witnessed growing interest in reforms intended to reinvigorate antitrust (see Shapiro, 2018, and the references cited therein), and our results suggest that stronger efforts to enforce existing laws could produce meaningful changes. The transition from McKinley to Roosevelt had a significant impact on firm valuations, and resulted in a large number of antitrust suits by the federal government that almost certainly would not have been initiated under McKinley. The structure of antitrust enforcement is much more institutionalized today than it was in 1901 (Crane, 2011), and recent presidential administrations have exhibited a high degree of continuity in their approaches to the issue (Crane, 2012). Nonetheless, scholars interested in designing strategies to address the growth of economic concentration should not neglect the role of enforcement efforts. One of the most significant changes in antitrust enforcement of the Gilded Age resulted not from new legislation, but from a change in the approach taken to the enforcement of existing law when Roosevelt became president. This finding is significant because transformative new legislation may be more difficult to implement than discretionary changes in enforcement.

Our paper has broader implications for the role of discretion and top political leadership in regulation. A relatively new literature in economics has sought to analyze the importance of the enforcement of regulations on the economy, particularly on financial markets. To obtain identification, these studies have primarily focused on the resource capacity or preferences of more decentralized political agents. For example, Ponticelli and Alencar (2016), and Brown et al. (2017) show that the quality of court enforcement affects firms' access to finance, whereas Charoenwong et al. (2019) exploit a mandated shift from federal to state-securities regulators to show that investment advising quality suffers when enforcers are more resource constrained. We add to this literature by focusing on the importance of enforcement discretion of the top political leader—the President—and by providing evidence in an area—antitrust policy—that is once again at the forefront of the

³In contrast, Feinstein et al. (2019) show that the party of state attorneys general does not seem to affect the enforcement of mortgage lending laws, suggesting that party identification may not always determine enforcement preferences.

political and economic debate.

Our paper also relates to the literature on the effects of policy uncertainty or political risk on economic activity and equity markets (Baker et al., 2016; Pástor and Veronesi, 2012). Although that work has primarily focused on aggregate sources of risk, Hassan et al. (2017) show that exposure to political risk varies considerably across firms, depending on how likely they are to be the target of regulatory efforts. Our study demonstrates that top leadership discretion over the enforcement of existing regulations may be an important source of political risk for firms.

This paper also contributes to an older literature that has utilized stock prices to evaluate the effects of antitrust regulation on American companies. These studies often find modest to negligible effects of antitrust enforcement (Binder, 1988; Bittlingmayer, 1993), merger activity (Eckbo and Wier, 1985), and forced dissolution of trusts (Burns, 1977). These studies, however, focus mostly on events that were likely to have been anticipated by the market, and were therefore largely priced in during the period under analysis. In contrast, we study unanticipated events and find larger negative effects of antitrust enforcement on potential targets.

Finally, our results also shed light on the longstanding debate among historians regarding the McKinley and Roosevelt presidencies. Whereas McKinley is traditionally viewed as a puppet of plutocrats (Josephson, 1934), a revisionist view has argued that he had shifted toward a more reformist orientation by the time of his assassination and would have acted against the trusts had he served out his second term (Phillips, 2003; Morgan, 2003). Likewise Roosevelt is often portrayed as a bold reformer (Morris, 2001; 2002), whereas others have argued that he was in fact quite conservative (Kolko, 1967), and that he "might not perhaps have been a progressive at all if it were not for the necessity of fending off more radical threats" (Hofstadter, 1955). The large stock price movements documented in this paper around McKinley's assassination imply that the market perceived Roosevelt to be quite different from McKinley in his stance toward business.

1 Historical Background: McKinley, Roosevelt, and the Assassination

1.1 McKinley and Antitrust

William McKinley's rise to the presidency in 1896 was largely due to the efforts of political entrepreneur Mark Hanna, who engineered McKinley's campaigns for Governor of Ohio, Congress, and President of the United States. A businessman with interests in banking, newspapers, and several other sectors, Hanna became active in politics as a representative of "the business interest" in the Republican Party (Croly, 1912: 145).

Hanna's success as the chair of the Republican National Committee was due in part to his abilities as a fundraiser. He appealed directly to the wealthiest industrialists and financiers for contributions, and quickly amassed unprecedented sums to finance the 1896 presidential campaign (Pollock, 1926). The political plat-

⁴In contrast, Prager (1992) finds that various decisions related to the *Northern Securities* case had a significant legal precedent effect on the value of other railroads from 1901 to 1905. Mullin et al. (1995) find large positive effects on the value of downstream customers of US Steel following the initiation of the (unsuccessful) suit for its dissolution in 1911.

form of the Democratic Party also aided Hanna's fundraising efforts. Whereas the Republicans endorsed the gold standard, the Democrats nominated William Jennings Bryan, a populist who advocated for free coinage of silver at an overvalued rate and for the regulation of trusts. Hanna obtained significant contributions from the financial and industrial interests that were most threatened by the prospect of a Bryan presidency, such as Standard Oil and J.P. Morgan, and raised at least \$3.5 million (Eichengreen et al., 2017), whereas the Democrats' campaign budget was just over \$400,000 (Croly, 1912: 220).⁵

During McKinley's presidency, the greatest wave of industrial mergers in American history took place. These were primarily horizontal combinations—the merger of many of the competitors in an industry into a single firm—undertaken at least in part to pursue anticompetitive objectives (Lamoreaux 1985, 2019). Between 1898 and 1902, the peak years of the movement, more than 2,600 firms were absorbed into mergers, and the total capitalization of the resulting firms was more than \$6 billion (Nelson, 1959). Although not as large, a significant wave of railroad mergers also occurred during McKinley's presidency, particularly in 1900 and 1901.

These developments led to calls for an aggressive political response to what was called the trust problem. Economists such as Richard T. Ely argued that the "evils" of the new monopolies went beyond high prices, and linked them to increased concentration of wealth and economic power. Writing somewhat later, John Bates Clark (1904: 1) stated that there was "no more startling or disquieting tendency of recent times" than mergers that monopolized industries. Although there was broad debate regarding the causes of the merger wave and appropriate solutions to the problems it created, there was a general consensus that the federal antitrust law, the Sherman Act of 1890, had had little impact (Ely 1900: 243).

In fact the Sherman Act may have actually helped stimulate the merger wave (Bittlingmayer, 1985). In its *E. C. Knight* decision of 1895, the Supreme Court held that since manufacturing itself was not commerce, the monopolization of manufacturing capacity through the merger of competitors was not subject to the federal government's constitutional power to regulate interstate commerce.⁸ This apparently meant that competing firms would run afoul of the law if they formed a cartel, but would not if they simply merged. In its later *Trans-Missouri* and *Joint Traffic Association* decisions, the Supreme Court held that the Sherman Act applied to railroads and that cartels among railroads were illegal.⁹ Although the *Knight* decision did not address railroad mergers, these decisions against railroad cartels likely contributed to the subsequent merger activity in that sector.

Although it may appear that the Supreme Court's narrow interpretation of the Sherman Act produced the merger wave, it is likely that the quality of the efforts made to enforce the act—to select cases to pursue,

⁵As a constant share of GDP, the Republican budget in the 1896 campaign was equivalent to \$4.2 billion in 2016, much more than the estimated \$2.65 billion spent in total on the 2016 presidential campaign. In contrast, the Republicans had raised only \$1.6 million for the presidential campaign of 1892 (C. N. Bliss, Jr., in Senate Subcommittee on Campaign Finance (1913: vol 1, p. 204); data for the 2016 campaign reported in www.cbsnews.com/news/election-2016s-price-tag-6-8-billion/, accessed 12 February 2018).

⁶During those years, nominal GDP averaged less than \$20 billion.

⁷Ely (1900: Ch. 6). See also Jenks (1901).

⁸United States v. E. C. Knight, 156 U.S. 1 (1895).

⁹United States v. Trans-Missouri Freight Association, 166 U.S. 290 (1897); United States v. Joint Traffic Association, 171 U.S. 505 (1898).

and to argue those cases—were ultimately responsible.¹⁰ McKinley's advisers were closely aligned with financiers and industrialists, and "antitrust enforcement reached a low-water mark equaled during no other period" under his presidency (Thorelli, 1955: 405). Although his attorneys general argued that the *E.C. Knight* decision tied their hands (Letwin, 1965: 137-42), the facts of that case were particularly unfavorable to the application of the Sherman Act.¹¹ By selecting a stronger case and pursuing a different legal strategy, it would likely have been possible to apply the Sherman Act to mergers, but the McKinley Administration never even attempted to do so.¹² A presidential administration hoping to restrain anticompetitive behavior would have at least tested this possibility; that McKinley's attorneys general never did reflects a policy preference that was likely understood by corporate promoters at the time.

1.2 Vice President Roosevelt

McKinley's Vice President, Theodore Roosevelt, was quite different. As Governor of New York, he enacted important Progressive reforms, including a corporate franchise tax, and successfully opposed the reappointment of a corrupt "machine politician" as the state's regulator of insurance companies (Roosevelt, 1920: 285-304). His annual message as governor in 1900 focused on the trust problem, and detailed many abuses perpetrated by large corporations, including "unfair competition," "raising of prices," and the "crushing out of competitors who do no act improperly." As a state governor, he did not specifically comment on federal antitrust enforcement, but he did call for corporation laws to include stricter disclosure requirements, to help with the design and enforcement of business regulations (Roosevelt, 1926).

Following the battle over the choice of insurance commissioner, representatives of New York's insurance companies urged Senator Thomas C. Platt, boss of New York's Republican machine, to find a way to remove him from office. Since Roosevelt was an extremely popular politician, challenging his renomination for governor would have been folly. Instead, a plan was devised to add him to the national ticket, to replace the recently deceased Vice President Hobart. Although neither Hanna nor McKinley publicly opposed the nomination, behind the scenes Hanna tried to build support for other candidates.¹³ The support Roosevelt received at the Republican convention ultimately proved impossible to resist; Hanna quietly withdrew his

¹⁰Kovacic and Shapiro (2000: 45). The Sherman Act was enforced by the U.S. Attorneys and the Attorney General. In 1903, the office of Assistant Attorney General was created, and this office was placed in charge of antitrust enforcement. See Crane (2011).

¹¹McCurdy (1979: 328) argues that "The Justice Department hardly could have chosen a weaker case." Among the many problems with the case was that the firms being acquired were all located in one state, Pennsylvania, and fell within the jurisdiction of that state's corporation laws, which could have been used to block the merger. The Fuller court was reluctant to act in ways that might supersede the state's own jurisdiction over its corporations, and in fact expected the attorney general of Pennsylvania to initiate a suit to invalidate the merger (McCurdy, 1979: 308).

¹²In the Knight case, the government argued that the purchase of the stock of competitors by the American Sugar Refining Company constituted a restraint of trade, which would violate the terms of the Sherman Act. But Justice Fuller's decision noted that government did not actually document any restraints on trade arising from the purchase, and implied that if the government had actually done so, the outcome would have been different (Letwin, 1965: 165). In addition, the Supreme Court's subsequent *Trans-Missouri*, *Joint Traffic* and *Addyston Pipe* decisions broadened the potential options for pursuing cases against anticompetitive mergers.

¹³In a possibly apocryphal statement, Hanna is said to have exclaimed: "Don't any of you realize that there's only one life between that madman and the Presidency?" (Leech 1959: 537). Morgan (2003: 376) presents a slightly different version of this quotation.

objections, and Roosevelt himself accepted.

Privately, Roosevelt had expressed concerns regarding the McKinley Administration's extreme forbearance with respect to antitrust enforcement, and its electoral consequences for the Republican Party. In public, however, he aggressively defended the McKinley Administration's achievements.

1.3 The Assassination and Roosevelt's Accession to the Presidency

On Friday September 6, 1901, President McKinley was shot twice by an anarchist while attending the Pan-American Exhibition in Buffalo, New York. The best qualified surgeon available to treat him was a gynecologist with no experience with bullet wounds. The surgery was only partially successful, and one of the bullets was not found (Rauchway, 2003).

The shooting occurred at around 4 PM, just as the stock market closed. That evening, in the hotels where traders gathered after hours, "the air was filled with rumor and speculation" regarding the shooting and its consequences for financial markets. ¹⁴ Journalists asked prominent figures for comments regarding the possible consequences of Roosevelt becoming president. Many praised Roosevelt's essential "ability and integrity" and predicted that becoming president would make him "conservative and cautious," perhaps in an attempt to calm investors. ¹⁵ Yet some commentators were openly (if tactfully) pessimistic, and reported a "belief that Roosevelt is somewhat adversely inclined towards corporations," and that under a Roosevelt presidency, "all manufacturing and financial interests will suffer, and, of course, railroads will be seriously affected." One prominent commentator argued that a Roosevelt presidency would be bad for merger activity, stating that "plans for combinations in the railroad world will, however, have to wait until the news from Buffalo is definitely more reassuring." ¹⁷

On Saturday September 7, the market declined sharply, with stocks falling 6.2 percent on average. Following his surgery, McKinley's condition improved, and his physicians offered optimistic assessments of his chances. When the markets reopened on Monday, September 9, firms' valuations largely recovered, with an average increase of 3.3 percent. Still, financial markets continued to follow the daily reports of McKinley's condition very closely. "The prevailing opinion," according to the *Chicago Tribune*, was that "the stock market will be controlled largely during the coming week by the nature of the bulletins sent from the President's bedside." These remained consistently favorable, and by September 10 physicians declared that McKinley was "practically out of danger." ²⁰

In the early morning hours of Friday September 13, however, McKinley's condition suddenly became grave.²¹ The market opened with heavy declines, and by the time it closed, prices had fallen by an average of

¹⁴ Market to Be Supported," New-York Daily Tribune 7 September 1901.

¹⁵"Think Business Interests Safe," Chicago Tribune, 7 Sept. 1901.

¹⁶Cleveland Plain Dealer, 8 September 1901; Chicago Tribune, 7 Sept. 1901.

¹⁷Henry E. Wallace, in *New-York Daily Tribune* 8 September 1901.

¹⁸For example, the front page of the *New York Times* reported that "Mr. Roosevelt gets reassuring news" on September 8, and that "Physicians say they are certain he will get well. All symptoms favorable" on September 9.

¹⁹Chicago Tribune 9 September 1901.

²⁰"President Past the Crisis," New-York Daily Tribune 11 September 1901.

²¹The physician's bulletin issued at 3 AM stated that "the worst is feared. His death might occur any time," *New York Times*, September 13, 1901.

5.1 percent. This magnitude is similar to the decline experienced on September 7 in response to the shooting. But the decline on the 13th was purely a response to the expected transition from McKinley to Roosevelt; any adverse consequence of the shooting itself, such as expectations of greater social conflict due to the rise of anarchism, should have already been reflected in asset prices following the 7th.

President McKinley died at 2:15 AM on Saturday the 14th, and the NYSE was closed in mourning on that day. After taking the oath of office, Roosevelt announced: "In this hour of deep and terrible national bereavement I wish to state that it shall be my aim to continue absolutely unbroken the policy of President McKinley for the peace and prosperity and honor of our beloved country." Roosevelt also stated that he would retain McKinley's entire cabinet. This dispelled fears that he would break with McKinley on key policy issues and was interpreted as a sign that "the president's death would not be disturbing" to financial markets. When the market reopened on Monday the 16th, prices increased 4.9 percent on average.

Our empirical analysis utilizes the four most relevant dates described above to discern the effects of a change in policy between McKinley and Roosevelt: September 7 and 13, dates when the market received negative news about McKinley's health, and September 9 and 16, when the market was told that McKinley would survive and that Roosevelt would continue with McKinley's agenda, respectively.²³

1.4 Impact of the Assassination in Historical Perspective

One way to gauge the magnitude of the expected change in policy resulting from the sudden transition from McKinley to Roosevelt is to compare its effect on stock prices to those associated with other assassinations and assassination attempts made on U.S. presidents. Table 1 presents the stock market's one-day reaction to assassination attempts in which someone actually fired a gun at the president.²⁴

It should be noted that the variation of the timing of the different shootings relative to the opening hours of the NYSE, and the variation in the institutional response of the NYSE to the news of the shooting, limits comparability across events. For example, trading on the NYSE was halted shortly after President Kennedy was assassinated, which may have curtailed the decline in prices on that day. Nonetheless, some suggestive evidence on the perceived significance of the transition from McKinley to Roosevelt relative to other assassinations can be found in the table.

Excluding McKinley's, the shootings resulted in an average decline in share prices of 1.6 percent, suggesting that the transition from the president to the vice president typically does not signify a dramatic change in policy. Yet the stock market's decline in reaction to the shooting of McKinley was nearly four times as

²²Commercial and Financial Chronicle, 4 January 1902, p. 14.

²³One concern regarding stock market data is that they could reflect other sources of volatility, for example related to seasonal effects. Yet the standard deviation of daily returns of the Dow Jones Industrial Average (DJIA) and Dow Jones Transportation Average (DJTA) over September 6–16, 1901, was 0.030 and 0.025, respectively, nearly four times larger than it was on average over those same days in the years 1898–1904 excluding 1901 (0.009 for the DJIA and 0.007 for the DJTA); the p-value of the difference is less than 0.0000001 in both cases.

²⁴Table 1 excludes an attempt made on the life of President Andrew Jackson on January 30, 1835, as it occurred prior to the invention of the telegraph, making the arrival of news in New York difficult to time. It also excludes an attempt on Harry Truman's life by Puerto Rican nationalists, in which they fired weapons at police officers and Secret Service agents, but never at the president. Finally it also excludes an attempt on Gerald Ford's life in 1975 by Lynette "Squeaky" Fromme, a follower of Charles Manson, in which she pointed a gun at the president but had failed to chamber a round, so that it did not actually fire.

large as the average for the other events, and about twice as large as the decline in prices when President Garfield was shot, the second-largest stock market reaction of record. McKinley's was also the only presidential assassination that produced elevated economic policy uncertainty in Baker et al.'s (2016) study of twentieth century data.

The only similar event that provoked a reaction comparable to the one caused by McKinley's shooting was the heart attack suffered by President Eisenhower on the evening of Saturday, September 24, 1955. On the following Monday, shares on the NYSE fell by an average of 6.6 percent.²⁵ The market reacted strongly to Eisenhower's heart attack because it came so late in his first term that it was believed he would not be able to run for reelection and a Democrat would likely win the presidency in 1956.²⁶ Thus, the stock market's response to the accession of Roosevelt to the presidency was roughly comparable in magnitude to the effect of an expected transition from a Republican to a Democrat at the height of the Cold War.

The strong stock market reaction to McKinley's assassination suggests that Roosevelt was perceived to be quite different from McKinley, and less friendly toward business interests. It contradicts the arguments of some revisionist historians that Roosevelt's administration is best understood as a continuation of trends that developed under McKinley (for example, Phillips, 2003). In particular, Roosevelt's antitrust agenda was expected to differ significantly from McKinley's.

Some evidence that the market's assessments were correct is found in Figure 1, which presents the number of antitrust cases instituted by the federal government under presidents Harrison through Roosevelt. McKinley's administration initiated just three antitrust cases, a historical low, and even though the largest wave of industrial mergers in American history occurred during his time in office, none of McKinley's suits were against manufacturers. In contrast, Roosevelt initiated suits that established important precedents in his first term, and then used those precedents aggressively in a wide array of suits in his second term. The defendants in Roosevelt's suits included Standard Oil, Swift & Co. (the "Beef Trust"), American Tobacco (the "Tobacco Trust"), Du Pont, and several major railroads.

The strong aggregate reaction of the stock market to news relating to Roosevelt becoming president, however, could have reflected investors' views about a variety of other policies. To determine the effects of a sudden change in the expected enforcement of antitrust laws, our analysis compares the variation in market values for firms more and less likely to be directly affected by such changes in policy.

2 The Effect of McKinley's Assassination on Firms

2.1 Construction of the Sample

Our analysis focuses on the variation in market values of publicly traded firms in response to the assassination. In this section, we present a brief description of the sources and methods used to construct a new dataset containing a variety of firm characteristics, while the appendix presents more detailed information.

²⁵Calculated from CRSP; 950 price changes relative to the previous day were observed.

²⁶"Stock market trading, brokers said, appeared to be dominated by the conviction that President Eisenhower would not again be a candidate, and ... that only he could win in 1956 for the Republicans. Traders were credited with believing that a Democratic Administration would not be so friendly to business" (*New York Times*, 27 September 1955).

Our initial sample includes all railroads and industrial firms with shares listed on the NYSE in 1901. To calculate these firms' stock returns, we collect daily closing prices of common shares and information on dividend payouts from the *New York Times* from September 3 to September 21. At the time, the stock market was relatively illiquid; we observe prices on at least one date for only 99 of the 134 companies we identify as having common stock listed on the NYSE.

We focus on abnormal returns to remove the effects of general stock market price movements. We estimate the market-model parameters over the 75 trading days prior to September 3, and exclude from the analysis those companies for which we observe returns on fewer than half of those dates. In addition, we collect whatever accounting data and firm characteristics are available from contemporary sources such as *Moody's Manual*. Our main sample consists of 48 firms (28 railroads and 20 industrial firms) for which we observe abnormal stock returns and basic firm characteristics. Appendix Section A.4.1 shows that our results are robust to instead using unadjusted returns, which allow us to include in the sample a total of 71 firms.

It should be noted that many firms whose values were likely to have been quite sensitive to changes in antitrust enforcement, such as Standard Oil, are not included in our data because they were not listed on the NYSE. Our estimates may therefore understate the true impact of Roosevelt's accession to the presidency.

2.2 Hypotheses and Main Variables

2.2.1 Vulnerability to Antitrust Enforcement

If investors expected Roosevelt to act more aggressively than McKinley toward anticompetitive behavior, the market values of firms that were likely to become subject to antitrust enforcement would have suffered disproportionately on days of bad news concerning McKinley's health. By contrast, we would expect these firms' valuations to gain disproportionately on the days that doctors said McKinley would recover and that Roosevelt stated he would follow McKinley's agenda. Yet identifying the firms that were particularly vulnerable to stronger enforcement of the Sherman Act presents a challenge. It may not have been clear, for example, which strategies Roosevelt's administration would have been most likely to pursue, or how receptive the courts would have been to those strategies.

We argue that firms that had participated in merger activity in the years leading up to the assassination would have been considered likely targets of renewed efforts to enforce the Sherman Act. These are the firms that likely benefited most from the McKinley Administration's lack of effort in enforcing the Sherman Act.

As we discussed in Section 1.1 above, the industrial consolidations undertaken during the great wave of mergers that followed the Supreme Court's 1895 *E.C. Knight* decision were likely to have been conceived with the expectation that the McKinley Administration would not attempt to use the Sherman Act to restrain them, and their design and conduct likely reflected this expectation. These mergers also attracted a great deal of public attention, partly because of their massive scale, but also because in some cases firms that had participated in cartels to restrain competition were consolidating to protect their rents. If Roosevelt had been expected to strengthen enforcement of the Sherman Act, these firms would have been among some of the

most likely targets. Therefore, we construct an indicator variable, *Merger*, for industrial firms that were incorporated in 1895 or later as a product of consolidations, since these firms would have been differentially vulnerable to antitrust enforcement.

Among the railroads in our sample, there was no equivalent post-1895 merger wave, because many of these firms were in receivership in the wake of the Panic of 1893. However, some significant merger activity among railroads did occur in 1900 and 1901, after the industry had recovered, and after the Supreme Court's *Joint Traffic Association* and *Trans Missouri* decisions, which held that cartels among railroads were illegal. These mergers included some very large, high-profile firms, and may have also been undertaken with the expectation that they were unlikely to be subject to antitrust enforcement.²⁷ Thus, we identify railroads that were involved in merger activity in the 1899–1901 period from various editions of the *Moody's Manual* and the *Poor's Manual*, and designate them as differentially vulnerable to antitrust enforcement.

In Appendix Section A.4.4, we present a brief description of the history of antitrust doctrine to further validate our approach to identifying firms that were differentially vulnerable to more aggressive antitrust enforcement, and also perform robustness analyses to evaluate the importance of alternative measures related to antitrust, such as the degree of concentration of business activity in the firm's industry.

2.2.2 Other Firm Characteristics

Firms that engaged in merger activity may have differed from those that did not in other respects. Thus, an important concern for our analysis is that our main variables of interest may reflect the effect on stock market valuations of other firm characteristics correlated with mergers but unrelated to antitrust enforcement. Our analysis, therefore, includes a variety of firm characteristics that may have resulted in differential stock price responses following the assassination.

The first is the amount of accounting data that the sample firms actually disclosed to investors. In 1901, NYSE-listed firms were subject only to weak disclosure requirements, which were not rigorously enforced by the exchange. Many industrial firms published little detail in their income statements, and some produced no income statements at all.

As governor of New York, Roosevelt advocated for stricter disclosure requirements for major corporations, and made them the centerpiece of his discussion of the trust problem in his 1900 annual message. Roosevelt argued that a requirement for more detailed disclosures would address the problem of "misrepresentation or concealment regarding material facts," which he considered to be among the "chief abuses" of the trusts. He also regarded disclosure as a necessary first step for the design of a policy program to restrain the trusts: "there may be other remedies, but what these are we can only find out by publicity" (Roosevelt, 1926: 46–47). Financiers with ties to major trusts were concerned about the possibility that Roosevelt would impose new disclosure requirements, and met with him immediately after he took office to try to persuade

²⁷Among these mergers was the joint acquisition of the Chicago, Burlington and Quincy by the Northern Pacific Railroad and the Great Northern Railroad. This latter merger was subsequently cemented through the formation of the Northern Securities Company, which owned all three railroads. The legal actions of Roosevelt's government against this holding company are the focus of our analysis in Section 4. Contemporary commentary emphasized the role of acquisitions in the industry during this period in bringing competitive railroads together and highlighted their importance for the railroad valuations (see, for example, Sage et al. 1901).

him against doing so.²⁸

Corporations that had chosen to disclose little accounting information, perhaps to evade regulatory scrutiny, may have been expected to suffer under a Roosevelt presidency. In order to address this possibility, we obtained the financial statements of all sample firms published in their most recent annual reports, as reproduced in *Moody's*, and measure the detail of their income statements by counting the number of lines.²⁹ At the time, railroads were subject to standardized disclosure requirements imposed by the Interstate Commerce Commission (ICC), so there was no variation in the number of income statement lines among those firms. To account for the fact that our measure of income statement detail will be correlated with being a railroad, which itself may have been perceived to be differentially affected under Roosevelt, in all of our regressions we include an indicator for railroads interacted with the relevant event dates.

The second set of characteristics capture political connections to McKinley or Roosevelt. A significant body of work has shown that the political connections of firms may affect their market valuations, both in modern economies as well as in the past (see, for example, Fisman, 2001; Ferguson and Voth, 2008; Fisman et al., 2012; and Braggion and Moore, 2013). Firms managed by individuals who had personal ties to McKinley or Hanna, or who had contributed to McKinley's campaign, might have been perceived to lose political influence in the event of McKinley's death, for example. In contrast, firms with ties to Roosevelt may have differentially benefited when the likelihood of the transition increased. As such firms may also have engaged in merger activity, it is important to control for such ties in the analysis.

To measure ties to McKinley and/or Hanna, we construct an indicator variable, *Donor*, for firms owned or managed by major donors to the McKinley campaign. Although no comprehensive list of donors exists, the two largest donors were J.P. Morgan and Standard Oil. We identify firms affiliated with J.P. Morgan, Standard Oil, or the Rockefellers (the major owners of Standard Oil), as indicated in Moody (1904). If those donations were perceived to buy influence with McKinley, but not with Roosevelt, the affiliated firms could have suffered differentially in response to the assassination.

To study the role of personal ties to Roosevelt, we identify whether any director of a firm had a connection to Roosevelt while he was a student at Harvard. Specifically, we identify the names of graduates of Harvard in 1880 (Roosevelt's class) and the names of members of clubs of which Roosevelt was also a member that were in the graduating classes of 1877 to 1883. We create an indicator variable, *Roosevelt*, that takes the value one for companies that had a director or officer (as listed in the *Moody's Manual*) who was in one of these clubs or graduated from Harvard in 1880. Approximately 17% of the firms in our sample had such a connection.³⁰

²⁸In early October 1901, two J.P. Morgan & Company partners, George W. Perkins, who had known Roosevelt since he was governor of New York, and Robert Bacon, one of Roosevelt's Harvard classmates, met with Roosevelt. Of the meeting, Roosevelt wrote: "Perkins wanted me to ... accept the publication of what some particular company chose to publish, as a favor, instead of demanding what we think ought to be published from all companies as a right" (Letter to Douglas Robinson, 4 Oct 1901, in Roosevelt 1951b).

²⁹We focus on income statements because they presented information of the greatest value to antitrust enforcement, such as revenues, costs, and profits. Many industrials published income statements of only a few lines, with no information at all on revenues or costs, and some published no income statements at all. In contrast, most sample firms published at least a rudimentary balance sheet.

³⁰Recent finance papers use a less restrictive definition of educational connections. For example, Cohen et al. (2008) identify all graduates of a particular school as connected regardless of attendance cohort. This approach is less

By contrast, McKinley did not have such personal connections to directors of major corporations because he was from a modest family in Ohio, did not attend prominent educational institutions, and was not a member of elite social organizations.³¹ An attempt to link members of McKinley's Civil War regiment, the 23rd Ohio Infantry, to corporate directors and officers produced no matches. Thus, we cannot assess the role of personal connections to McKinley, and focus solely on ties with campaign donors.

Finally, the assassination of McKinley may have introduced greater uncertainty regarding the outlook for future economic policy and conditions in the market for reasons other than changes in antitrust policy. If the rise in uncertainty differentially affected the valuations of larger or more profitable firms, which may have been more likely to engage in mergers, we may erroneously attribute the effects of the assassination to antitrust enforcement. Unfortunately, the lack of consistent financial statements precludes us from obtaining reliable measures of cash flow, profitability, or total assets for a large fraction of our sample. Therefore, we use the log of the book value of total capital, which we can observe for all firms in the sample, to control for firm size.³²

2.3 Summary Data

Table 2 presents summary statistics for the 48 firms with data on abnormal returns, which are the main focus of our analysis. A detailed tabulation of firm characteristics for the individual firms in this sample is provided in Appendix Tables A.2 and A.3.

An indication that merger activity was quite common during McKinley's presidency is that our sample is evenly split between firms that had recently engaged in a merger, presented in column (1), and those that had not, shown in column (2). These two groups of firms were remarkably well balanced along several important characteristics—for example, each group was composed of 14 railroads and 10 industrial firms, and contained four firms with connections to Roosevelt. Column (3) shows that there were also no noticeable differences in firm size, or in the export share of the firms' products. We also find no statistically-significant differences in the disclosure of information in the firms' income statements, both when we focus on the entire sample and when we restrict the data to industrial firms. However, recent-merger firms were more likely to be connected to major donors to the McKinley campaign and were younger. These characteristics may have made these firms more vulnerable to higher economic uncertainty or to a transition from McKinley to Roosevelt, independent of any changes in antitrust policy. We therefore control for firm characteristics and political connections in our empirical analysis.

informative in our case as nearly all firms in our sample had at least one director who attended Harvard University.

³¹McKinley graduated from Albany Law School with William E. Barnett, future director of the New York, New Haven and Hartford Railroad, and Goodwin Stoddard, future director of the Atlantic Coast Line Railroad Company. While these two companies were listed on the NYSE in 1901, their shares were illiquid and did not trade on days of interest surrounding the assassination of President McKinley.

³²We do observe dividend payouts, a measure imperfectly correlated with profitability, for all firms in the sample. Reassuringly, controlling for the firms' dividend payout rates prior to the assassination does not materially affect our main estimated effects for recent mergers (results not shown).

3 Impact of the Assassination on Firm Values

The daily abnormal return of NYSE-listed stocks on the dates surrounding the assassination are illustrated in Figure 2, with those of recent mergers presented separately from other firms. The recent mergers exhibit sharp declines on September 7 and 13, the dates when the market responded to McKinley's shooting and expected death, which are offset by strong increases on September 9 and 16, the dates when the market was told that McKinley would survive, and that Roosevelt would continue with McKinley's agenda. In contrast, the other firms' returns exhibit much more muted changes on those dates. This is a clear indication that the transition to Roosevelt was expected to reduce the value of the firms we designate as vulnerable to stricter antitrust enforcement. In what follows, we explore these differences in greater depth.

3.1 Baseline Specifications

We estimate regression models in which we interact our variables for sensitivity to antitrust enforcement with indicators for the four dates when the market responded to significant news regarding the transition from McKinley to Roosevelt: September 7, McKinley's shooting; September 9, the prognosis that McKinley would survive; September 13, McKinley's expected death; and September 16, Roosevelt's statement that he would continue "absolutely unbroken" with McKinley's agenda.

Our panel data regressions for the 14 trading days the NYSE was open between September 4 and September 21, 1901, take the following form:

$$AR_{it} = \lambda_1(Merger_i \times Sept 7_t) + \lambda_2(Merger_i \times Sept 9_t)$$

$$+\lambda_3(Merger_i \times Sept 13_t) + \lambda_4(Merger_i \times Sept 16_t) + \gamma_i + \delta_t + \epsilon_{it},$$
(1)

where AR_{it} is firm *i*'s abnormal return on day *t*; Merger_i is our indicator for firms that were differentially vulnerable to stricter antitrust enforcement; γ_i are firm fixed effects; and δ_t are fixed effects for each trading day in the sample. The parameters of interest are λ_1 to λ_4 , which capture the differential effect of negative or positive news on McKinley's condition for firms that were particularly sensitive to a stronger antitrust regime. To control for the effects of ties to the Presidents or other firm characteristics, we interact those variables with the same four event dates. In all specifications, we control for firm size and an indicator for railroads (interacted with event dates), and cluster standard errors by firm.

Table 3 presents the results for specifications that analyze the effect of the recent merger variable on abnormal returns. Consistent with the replacement of McKinley by Roosevelt being harmful to the valuations of firms that were sensitive to greater antitrust enforcement, the estimates in column (1) indicate that the recent mergers in our sample saw their abnormal returns fall by an additional 175 basis points (bps) on September 7, a sizable effect relative to the average decline of 103 bps on that date. When the stock market rose in response to good news on McKinley's health on September 9, the abnormal returns of recent merger firms differentially increased by a similar amount, 197 bps.

When McKinley's health took a turn for the worse, on September 13, the abnormal returns of firms that had engaged in mergers declined relative to others by 141 bps, and they increased by 197 bps on September

16, when Roosevelt reassured markets. Importantly, any impact of the assassination attempt, rather than the change in the identity of the president, would already have been reflected in prices on September 7 and 9. The remarkably similar estimated effects on September 13 and 16, therefore, suggest that the presidential transition itself was the primary concern of investors. The pattern of stock returns is consistent with recent merger firms losing about 1.4 to 1.9 percentage points of their value relative to other firms in the sample in response to an increase in the perceived likelihood of stronger antitrust enforcement.

Next, we investigate whether the effects we ascribe to antitrust enforcement were due to other firm characteristics that may have been correlated with our antitrust variables. In column (2), we allow the returns to vary on the relevant dates by the level of firm disclosure, as measured by total income statement lines. The estimated parameters indicate that, relative to a firm with no income statement (0 lines), a firm reporting 5 lines (the median length for industrial firms in the sample) saw its value rise by about 90-125 basis points on days of increased probability of a change in administration, and decline by 70-110 basis points when McKinley's policies were perceived to be more likely to continue. To the extent that a lack of disclosure of financial information facilitated collusion or abuses by the trusts, we would expect those firms that were more transparent to have been less impacted by a transition to Roosevelt. Our findings are consistent with this interpretation, although the estimated magnitudes are smaller and not quite as consistent as that of the parameters associated with the recent merger variable. More importantly, the estimated effects on recent mergers are unaffected by including these controls, suggesting that our baseline results are unlikely to be driven by characteristics associated with firm transparency, or a perception that Roosevelt would strengthen corporate regulation on issues unrelated to antitrust.

To rule out the possibility that our results are driven by political or personal connections to the presidents, in column (3) we include interactions for firms controlled by major donors to McKinley and add an indicator for firms managed or controlled by individuals with personal ties to Roosevelt. And the specification in column (4) adds log firm age as a measure of more mature or stable firms. The estimated parameters associated with these variables (presented in the appendix) are generally imprecise, but more importantly, the effects of *Merger* are robust to the inclusion of these additional controls, which include those that differed across the two groups of firms in Table 2. The magnitudes of the estimated coefficients are quite stable across the different specifications, and most of them remain statistically significant.

In the appendix, we show that these results are robust to including alternative controls for firm size, profits, and investment rates—see Appendix Table A.7. This suggests that our findings are unlikely to be driven by differential pre-trends in size or profitability across firms.

In interpreting these baseline results, it is important to bear in mind that, as in any event study design, the estimated effects are based on expected values. That is, we measure the effect of a new antitrust regime on firm values, weighted by the probability of such change actually occurring. To the extent that the probabilities were seen as relatively low for the typical firm engaged in recent mergers, our results suggest that a change in antitrust enforcement was expected to have sizable effects on the value of prosecuted firms.

3.2 Alternative Policy Preferences

Roosevelt's policy agenda may have been expected to differ from McKinley's on issues beyond antitrust enforcement. An important source of concern for our empirical approach is that the firms we designated as vulnerable to antitrust enforcement may also have been vulnerable to changes in other policies Roosevelt was expected to enact.

On what was likely the most important economic policy issue of the time, commitment to the gold standard, Roosevelt did not in fact differ from McKinley. However, Roosevelt's preferences differed from McKinley's on another important issue, tariff policy. Whereas McKinley favored tariffs quite strongly, and had been a sponsor of tariff legislation while in Congress, Roosevelt was more of a free-trader. Businesses that had previously benefited from tariff protection may, therefore, have seen their valuations fall in response to Roosevelt unexpectedly becoming president. To address this possibility, we control for the trade exposure of our sample firms, measured as the share of output that was exported at the industry level for the industrial firms in our sample. For railroads, we construct data on the share of revenues accounted for by export products.

In Panel A of Table 4, we include our measure of exposure to exports in regressions otherwise identical to those presented in columns (1) and (4) of Table 3. The signs of the estimated parameters for this variable are generally consistent with the notion that Roosevelt's agenda would benefit exporters. But including these interactions does not meaningfully change the estimated magnitudes associated with the antitrust variables.

McKinley and Roosevelt may also have differed on their views on labor issues, though not all historians agree on this point. In fact, McKinley was not as anti-labor as some of his contemporaries, and unlike President Cleveland before him, he did not pursue any Sherman Act cases against unions. Rove (2015), for example, argues that McKinley was trying to transform the Republican party to seek the support of foreignborn urban workers. Nonetheless, Roosevelt was relatively friendly towards labor interests, and he helped to enact several important pieces of labor legislation as governor of New York. Firms that were differentially vulnerable to labor activism would perhaps have seen their valuations fall in response to Roosevelt unexpectedly becoming president. We use an event study analysis of a major labor strike that occurred during Roosevelt's administration to assess whether our results could be related to changes in expected labor relations policy.

Specifically, in Panel B of Table 4 we analyze the market's reaction to Roosevelt's intervention in an ongoing coal strike among miners in Pennsylvania's anthracite coal fields. Previous presidents had generally intervened on behalf of employers in labor relations disputes, if they intervened at all. As the strike dragged on, mine operators called for Roosevelt to end the strike by deploying federal troops. The operators argued that the president was granted the authority to do so under the Sherman Act of 1890, as they viewed the union as a unlawful cartel. Roosevelt disagreed; only in response to concerns that the coal strike, which began in May 1902, might substantially raise coal prices for households over the winter did Roosevelt decided to intervene, but he did so in a relatively neutral way. On October 1, 1902, Roosevelt announced that he would meet with representatives of government, management, and labor at the White House in order to facilitate a resolution to the dispute. This marked a substantial change from historical precedent.

If recent mergers were correlated with sensitivity to a change in labor relations policy, then those firms

should have performed differentially worse in response to the announcement of the conference. Panel B of Table 4 presents the results of regressions similar to those reported in columns (1) and (4) of Table 3, but using cumulative abnormal returns during an eight day window around the October 1 announcement. The estimated effects are, on average, positive, and not statistically significant, but the coefficients are too imprecisely estimated to confidently rule out a sizable effect in either direction. These results provide some suggestive evidence that, at least as indicated by this event, the effects on firm valuations that we estimate around the presidential assassination are unlikely to be primarily driven by expectations of changes in policies to favor workers.³³

Roosevelt was also more committed to an aggressive, imperialistic foreign policy than McKinley. It is difficult for us to assess this potential confounding factor directly, since our sample lacks clear variation in exposure to greater military spending or foreign conquests.³⁴ However, very early in his presidency, Roosevelt signaled that he was interested in finding a way to get the Panama canal completed, and it is possible that this could have been anticipated when he became president. We show in Appendix Table A.9 that our main results are not driven solely by intercontinental railroads, which would likely have suffered differentially from the opening of the Panama canal.

Finally, we provide additional evidence in support of our claim that expected changes in antitrust enforcement drove the stock market's response to McKinley's assassination by analyzing the returns of firms that would likely have benefited from stricter antitrust enforcement: those that had recently planned, or announced, mergers that ultimately failed to occur. These included competitors of dominant firms in their industry and relatively weak firms that sought to strengthen their positions through mergers, but failed.³⁵ If these firms faced competition from larger or more efficient rivals, they would have benefited from stricter antitrust enforcement. In Appendix Table A.10, we show that the valuations of these firms fell by less than others on the two days of bad news regarding McKinley's health, and rose by less on the days when it seemed that McKinley would survive or that Roosevelt would follow McKinley's agenda.

4 Litigation against Northern Securities

Our interpretation of the events surrounding McKinley's assassination suggests that investors believed that the replacement of McKinley by Roosevelt would result in more aggressive antitrust enforcement. Yet it is possible that investors could have expected other changes in policy or personnel in response to Roosevelt's accession to the presidency. To test our interpretation, and the validity of the recent merger variable as an

³³An alternative strategy for addressing this concern would be to directly control for measures related to the significance of labor relations for firm values in our baseline specifications. Unfortunately, the labor relations data available for our period are limited and problematic. But with this caveat in mind, in Appendix Section A.4.1.2 we show that our results are robust to controlling for industry-level indicators for labor activism.

³⁴For example, at the time, there were no ship builders listed on the NYSE, and the United Fruit Company was not yet listed. And nearly all of the listed manufacturing firms could have benefited at least indirectly from greater military spending.

³⁵For example, Glucose Sugar Refining, a smaller and weaker competitor of the dominant American Sugar Refining, announced in June 1901 that they had begun negotiations with National Starch and other glucose producers to merge into a larger and stronger firm. However, several directors opposed the initial proposal, and the merger initially failed. It was ultimately consummated in 1902 (Dewing, 1914: 87).

indicator of sensitivity to antitrust issues, we study the stock market's reaction to an event that revealed new information on Roosevelt's approach to antitrust: the announcement of Roosevelt's first antitrust suit. Unlike the assassination event, this one was unquestionably an indication that more aggressive antitrust enforcement was in store. Other suits filed subsequently by Roosevelt potentially revealed additional information about the enforcement strategy he chose to pursue, but only the first suit was kept secret before its announcement, and is therefore suitable for an event study.

In November of 1901, just after Roosevelt had become president, J.P. Morgan created an enormous holding company, the Northern Securities Company, to jointly own the stocks of two major competing railroads, the Northern Pacific and Great Northern, as well as third that connected them to Chicago, the Chicago, Burlington and Quincy. Recent Supreme Court decisions, such as *Trans Missouri* and *Joint Traffic Association*, held that the Sherman Act applied to railroads, and that cartels among them violated the Act. However, on the theory that the *E. C. Knight* decision indicated the Sherman Act did not apply to mergers, these firms had been joined together via a holding company.

The formation of the holding company was a provocation to Roosevelt.³⁶ The shear scale of the firm, with \$300 million in capital, along with its potential to monopolize rail transportation in a large area of the country, was of great concern to Roosevelt. He asked Attorney General Philander Knox to quietly study the possibility of pursuing an antitrust action against the company, explaining in his autobiography that he felt "It was necessary to reverse the Knight case" (Roosevelt 1920: 443). Knox studied the matter carefully, ultimately concluding that the Knight case had been poorly argued and a different strategy was possible in a suit against Northern Securities, which had a reasonable chance of success (Morris 2002: 88).

On February 19, 1902, after the stock market closed, Knox made the following announcement regarding the forthcoming legal actions against Northern Securities: "Some time ago the president requested an opinion as to the legality of this merger, and I have recently given him one to the effect that, in my judgment, it violates the provisions of the Sherman Act of 1890, whereupon he directed that suitable action should be taken to have the question judicially determined" (quoted in Meyer, 1906: 258).³⁷ The actual suit was filed on March 10, 1902.

The timing made it quite a shock. The Supreme Court was then considering a case filed by the attorney general of Minnesota against the railway combination, and was expected to announce its decision on Monday, February 24, 1902.³⁸ Contemporary reports suggest that the market expected the Supreme Court to reject

³⁶Even prior to the assassination, Hanna had warned J.P. Morgan that the formation of the holding company would likely force the government to act against it. Yet James J. Hill, who had significant holdings in the railroads being merged, did not believe that McKinley would actually seek to enforce the Sherman Act, since he had almost never spoken about the trusts (Haeg, 2013: 254).

³⁷In response, on February 23 J. P. Morgan himself went to Washington to meet with Roosevelt and Knox, bringing then-Senator Mark Hanna and Senator Chauncey Depew with him. Unaccustomed to policy decisions that concerned his interests being made without his consultation, Morgan said, characteristically: "If we have done anything wrong, send your man to my man and they can fix it up" (quoted in Bishop, 1920: 184). Over the subsequent weeks, Morgan conferred repeatedly with Hanna regarding the case, but the case against Northern Securities followed its course (Strouse 2000: 442).

³⁸On January 7, 1902, the attorney general of Minnesota, on behalf of his state and with the support of several other western states, moved for leave to file a bill of complaint against the company before the U.S. Supreme Court (Meyer, 1906).

Minnesota's suit, and the announcement of a federal suit just before the Supreme Court's decision generally came as a surprise.³⁹

The announcement confirmed that Roosevelt would indeed attempt to pursue a more aggressive antitrust agenda. In its effort to obtain a judgment that a merger of competitors violated the Sherman Act, the suit sought to expand the boundaries of antitrust doctrine, with potentially far-reaching consequences. We therefore study the effect of the unexpected announcement of the federal case against Northern Securities on the market values of firms that we designated as differentially sensitive to more aggressive antitrust enforcement. Yet, on February 20, it could not have been clear whether the suit would be successful, or if it were, how broad and applicable the decision would ultimately be. We are therefore cautious in our interpretation of any estimate effects of the announcement.

Figure 3 shows average cumulative abnormal returns separately for the firms that had engaged in recent mergers and those that had not, for all trading days from February 14 to March 1, 1902. We normalize the returns to zero for February 14, and cumulate them forward from that date. The returns of merger firms were similar to those of other firms until February 19, when they began to decline. Since the announcement that the government was going to challenge the merger happened late in the day, the filing primarily impacted stock prices on February 20.

It is possible that Roosevelt may have unintentionally revealed that a suit was forthcoming, which may explain the decline in abnormal returns among recent mergers on February 19. On the 18th, Roosevelt met with Mark Hanna and asked his opinion regarding the combination. Hanna, who was a shareholder and a close associate of some of the Northern Securities insiders, expressed enthusiastic support for the firm. It is not clear whether Hanna interpreted Roosevelt's question as a sign that an antitrust suit was in the works, and when the suit was actually announced Hanna is said to have been "thunderstruck" (Morris 2002: 89). But if Hanna had mentioned his conversation to others, or if Roosevelt had any other conversations about the firm on that day that are not known to historians, this could explain the fall in returns in anticipation of the announcement.

In any case, following the 20th, the returns of merger firms remained below those of non-merger corporations for at least a week. Importantly, this pattern was not driven by any direct effects on the shares of Northern Securities or the railroads included in the consolidation, since none of them had shares of common stock that traded on the NYSE at the time. Thus, the figure presents suggestive evidence that the revelation of the stronger stance of Roosevelt's administration against trusts differentially affected those firms that we characterize as ex-ante more sensitive to more aggressive antitrust enforcement.⁴¹

To more formally analyze the effect of the Northern Securities litigation on the market values of NYSE-listed firms, we employ an event study methodology. For each firm in the sample, we calculate cumulative abnormal returns from February 20 to March 1, 1902, and relate them to our measure of the threat of being

³⁹The *Washington Post*, for example, stated that the "announcement from Washington was therefore a rude shock to all of this optimistic sentiment that has been carefully nurtured in the financial district" (February 21, 1902).

⁴⁰This may indicate that Hanna did not anticipate a suit, but it may also indicate that he did not expect one so soon.

⁴¹Prager (1992) studies, instead, the effects on railroad valuations of the circuit courts' decisions that sided with the government's case in 1903. He finds that stock prices declined during those dates, and suggests that Northern Securities was an important development in antitrust policy because it established new precedent for railroad mergers.

subjected to antitrust suits. Our empirical strategy consists of estimating:

$$CAR_{i} = \alpha + \beta Merger_{i} + \delta X_{i} + \epsilon_{i}$$
(2)

where CAR_i is firm i's cumulative abnormal return from February 20 to March 1; Merger_i is an indicator for firms that were more likely to be targets of antitrust litigation due to recent merger activity; and X_i includes controls for firm-specific characteristics.

Table 5 presents the results; we modify the specifications presented across the different columns from those of Table 3 above, in response to the different character of the information revealed by the filing of the suit. In column (1), we include minimal controls. The results indicate that recent mergers lost about 1.3 percent of their value relative to other firms, a magnitude that is generally consistent with those associated with days of bad news regarding McKinley's health. In column (2), we include our measure of accounting disclosure, and find that it has little effect on returns. This is consistent with our interpretation of that variable as highlighting companies that would be affected by stronger disclosure requirements, the likelihood of which was not affected by the Northern Securities suit. As the suit was filed against a railroad, and the potential impact of any decision on firms outside the railroad industry could have been seen as ambiguous, in column (3) we add an indicator for railroads. This has little impact, and does not affect the magnitude of the merger variable.

Finally, in column (4) we add an indicator for firms affiliated with J.P. Morgan & Company. Morgan had engineered Northern Securities and was the most influential financier of his era. Roosevelt and Morgan saw each other partly as rivals (Wiebe, 1959), and one concern about the results could be that they reflect the expectation that Morgan firms would receive greater regulatory scrutiny. Although imprecisely estimated, the coefficient on the Morgan variable is negative. More importantly, however, the estimate on the merger variable is unaffected by the inclusion of the Morgan control.

By focusing on returns cumulated from February 20, our results are conservative and exclude the potential impact of anticipation effects. In the appendix, we show that our results are robust to performing the event study over different windows around the announcement of the administration's action against Northern Securities, and that they double in size when we begin the analysis on February 18 or 19. We also perform a placebo date study to rule out the possibility that our results are driven by serial correlation in returns. In Appendix Figure A.5 we show that recent merger firms had no differential returns over mutually exclusive windows in the month around the announcement of the suit, other than on the event date.

The relative decline in abnormal returns for recent merger firms is consistent with the suit being interpreted as a sign that stricter antitrust enforcement was likely, and recent mergers were among those likely to face greater antitrust scrutiny.

5 Robustness of the Results

5.1 Small-Sample Tests

Event-study hypothesis tests suffer from the problem that abnormal returns often have a non-normal distribution; the small number of firms and events in our samples may compound this problem, and produce inaccurate standard errors. We address this concern by implementing the binomial sign test of Dube et al. (2011), which enables us to calculate the p-value of obtaining a given number of events with abnormal returns above the 95th percentile within our pre-period sample of 75 daily returns. We also implement uniform rank tests, which similarly capture the probability that event-day returns lie at the extreme ends of an ordered series of pre-period daily returns, and present two standard Z-scores. The first was shown by Maynes and Rumsey (1993) to converge to the standard normal as the number of events increases. The second, by Joiner and Rosenblatt (1971), approximates the standard normal distribution even for very small numbers of events.

Table 6 presents the results of these tests. Column (1) reports the results for the McKinley assassination event dates. The differences between the abnormal returns of the recent mergers and those of other firms on the four McKinley assassination event dates were all above the 95th percentile of the pre-period distribution of returns, producing a one-sided p-value under the null hypothesis of 0.01%. The rank tests produce similar p-values. In column (2), we report the results for the 4-day cumulative abnormal returns for the Northern Securities suit. The returns for this event were also above the 95th percentile (p = 0.05) and the rank test statistics were of similar levels of significance.

In sum, these tests provide reassurance that the statistical significance of our main effects are not driven by our small sample size or the non-normal distribution of returns.

5.2 Other Robustness Tests

The appendix presents a detailed analysis of the robustness of our results. Here we provide an overview of those tests that we have not already discussed in prior sections.

Given our small sample size, one potential concern regarding our analysis is that our estimates could be driven by the returns of a few salient firms. In the appendix, we show that our findings are similar when we exclude any one or two firms from the sample—see Appendix Figures A.1–A.4. Our analysis also focuses on abnormal returns, which is standard in the literature. However, in the context of the relatively illiquid markets of 1901, this approach reduces the number of observations significantly, and it may also raise concerns related to the fit and precision of the betas we use. Yet in Appendix Tables A.5 and A.12, we show that our results are generally robust to the use of unadjusted returns, rather than abnormal returns.

An additional potential source of concern relates to the way we have defined our treatment variable.

$$\sum_{i=m_p}^{M} {M \choose i} (1-p)^i p^{M-i}, \tag{3}$$

where M is the total number of events.

 $[\]overline{^{42}}$ Dube et al. (2011) show that the probability of obtaining m_p events with abnormal returns above the pth percentile is:

The legal history we review in Sections 1.1 and A.4.4 suggests that, at that time, a strengthening in antitrust enforcement would likely have entailed pursuing actions against companies that had engaged in anticompetitive mergers. In contrast, modern antitrust doctrines have focused more directly on market shares, and one might wonder whether market shares, rather than recent merger activity, would have determined firms' vulnerability to more aggressive antitrust in 1901.

We do not utilize market shares in our main specifications for two main reasons. Most importantly, it would be somewhat anachronistic to do so: the *E.C. Knight* decision held that even a firm with 98% of the manufacturing capacity in its industry would not be considered to be in violation of the Sherman Act. Also, as we discuss in detail in the appendix, measures of market concentration for the early twentieth century are notoriously poor. Yet with these caveats in mind, in Appendix Tables A.13 and A.14 we show that the estimated impacts of the McKinley assassination and Northern Securities suit on recent mergers are similar when we include controls for various available measures of market concentration and coordination through common ownership.

6 Longer-run Outcomes

Our focus on stock market returns enables us to measure the immediate impact of expected changes in policy. These results are by nature short run and based on investors' expectations about future outcomes at the time of Roosevelt's accession to the presidency. To provide some insights into whether these changes actually had any longer-term impacts, we next study the outcomes of firms we designated as vulnerable to stronger antitrust enforcement relative to others in the years following the assassination. By focusing on our measure of vulnerability, our strategy provides insights into the effects of stricter enforcement on firm outcomes independent of whether those firms were themselves subject to direct legal action.

Given its importance for economic growth, we focus first on investment. In principle, the effects of stronger antitrust enforcement on investment are ambiguous. First, heightened uncertainty about the regulatory environment may have a detrimental effect on firms' investment (Pindyck, 1991). Stronger enforcement of antitrust statues may also increase competition. This, in turn, could lead to a decline in investment by eroding rents, although firms may instead increase investment rates in an attempt to escape competition (Aghion et al., 2005). Thus, whether investment increases or declines in response to more aggressive antitrust is ultimately an empirical question. We next study whether more vulnerable firms experienced differential declines in profitability. To the extent that Roosevelt's actions resulted in increased competition, more vulnerable firms may have seen their profits and markups dwindle relative to others. Firms may also have responded to anticipated changes in antitrust enforcement by altering their business practices in ways that reduced profitability, in an attempt to preempt enforcement actions against them.

To implement our tests, we collect annual accounting data for our sample firms from 1895 to 1905, from the E.C. Knight case that helped ignite the Great Merger wave, until it essentially ended. ⁴³ Given the state

⁴³We hand-collect these data from a variety of sources, including various editions of the *Moody's Manual of Railroads* and *Corporation Statistics*, *Poor's Manual of Railroads* manuals, and *Manual of Statistics Stock Exchange Handbook*. Definitions and details of how the variables were collected are presented in the Appendix.

of financial reporting at the time, these data are noisy and incomplete, particularly for industrial firms. It is important to be cognizant of the limitations of these data in the analysis that follows.

With the available data, we estimate regressions of the form:

$$y_{it} = \theta Merger_i \times Post1900_t + \beta X_{i,1901} \times Post1900_t \gamma_i + \delta_t + \epsilon_{it},$$
(4)

where y_{it} is the outcome of interest for firm i in year t; Merger $_{it}$ is our firm-specific measure of vulnerability to stricter antitrust; Post1900 $_t$ is an indicator for the years following 1900, when Roosevelt was president; and γ_i and δ_t are firm and year fixed effects. The coefficient of interest, θ , is the post-1900 difference-in-differences in the outcome for firms that were recent mergers relative to others. Our regressions also control for relevant firm characteristics, such as size (measured as log capital or log assets), age, leverage, and affiliation with J.P. Morgan, measured in 1901 and interacted with a post-1900 indicator. In order to ensure that our results are not the product of ongoing differential trends, in some specifications we add a Merger $_{it}$ × trend interaction, giving θ the interpretation of post-1900 deviation from trend. In the appendix, we further explore whether these results could be affected by ongoing trends by estimating a modified version of these regressions, with year-by-year interactions with the recent merger variable.

The results, with standard errors clustered by firm, are presented in Table 7. For each outcome of interest, we present two specifications, with those in the second column controlling for differential time trends by treatment status. The estimated effects in columns (1) and (2) present the results for investment, measured as the percent change in the firms' value of physical assets listed in their balance sheets.⁴⁵ The point estimates for the post-1900 difference in differences are positive, although the one in column (1) is not statistically distinguishable from zero. One must be cautious in interpreting such imprecisely estimated parameters, especially given that our estimate of θ is sensitive to the inclusion of the trend interaction. But we can safely conclude from these estimates that there is no evidence suggesting that investment contracted much more for firms with recent mergers in response to Roosevelt's policies. The 95 percent confidence interval for the estimate in column (1) is [-0.0097, 0.0591], ranging from a small negative effect to a large positive effect: Roosevelt's more aggressive antitrust stance did not significantly diminish investment rates. Our results therefore echo research on modern data that has associated growing levels of economic concentration and reductions in product market competition with declining investment rates (Gutiérrez and Philippon, 2017b). Like that literature, our findings suggests that efforts to strengthen antitrust and foster competition did not have a significant detrimental effect on investment, and may have actually increased it, early in the twentieth century.

The second outcome of interest is profitability, measured as return on equity (operating profits/total shareholders' equity). We focus primarily on ROE as opposed to other measures of profitability because sufficient data is available to calculate it for a larger number of firm-years (about 45 firms have non-missing data for at least some years). The estimates in columns (3) and (4) indicate that the return on equity of

⁴⁴If we use 1900 rather than 1901 values, the results are essentially the same, although some observations are lost.

⁴⁵To reduce the effect of outliers, we trim this variable at the 2% level. Many of the extreme values may have coincided with restructurings or M&A activity.

⁴⁶Many industrial firms did not disclose a full balance sheet, but they did report the book value of their shares and,

firms engaged in mergers prior to McKinley's assassination declined by 1.4 to 1.6 percentage points more than that of other firms, relative to the difference in profitability between these two groups in prior years. The economic magnitude of this effect is quite significant, as it represents about 18 percent of the mean return on equity in year 1900. This result is also generally insensitive to the inclusion of the trend interaction. Indeed, Appendix Figure A.6 suggests that the contraction in profitability occurred right around the change in administration, though the annual estimates are too noisy given the small sample size to establish this conclusively. Overall, our findings suggest that firms that were vulnerable to stronger antitrust enforcement experienced a larger contraction in profitability.

Yet as this measure of profitability is scaled by shareholders' equity, any changes in firms' capital structures would affect it. This raises the possibility that factors not directly related to antitrust could have been responsible for the patterns we observe. We investigate this issue in columns (5) and (6), in which we restrict the sample to railroads, whose detailed financial statements enable us to construct a measure of profit margins, defined as operating profits divided by total revenues.⁴⁷ The results indicate that the profit margins of railroads vulnerable to stricter antitrust enforcement did indeed fall under Roosevelt. However, when we control for a trend interaction in column (6), the estimated post-1900 effect is significantly reduced, suggesting that caution is warranted in interpreting the estimated magnitudes.

In interpreting the results from this section it is important to keep in mind that long-run difference-in-difference estimation is by nature less well suited for convincing identification than the high-frequency stock market analysis we present above. Nevertheless, our results provide suggestive evidence that a stronger stance towards antitrust enforcement may have reduced the profitability of firms that had engaged in anti-competitive mergers, without leading to significant contractions in their investments.

7 Conclusion

We study the assassination of President William McKinley in September 1901 to estimate the potential scope of political discretion in the aggressiveness of antitrust enforcement. The news of McKinley's shooting provoked a significant fall in stock prices. This decline was reversed when doctors subsequently announced that they expected McKinley to recover, and this reversal was itself reversed when McKinley's condition suddenly became grave. The latter fall in stock prices in response to the expectation that McKinley would die was purely a response to the transition from McKinley to Roosevelt, and not a reaction to the fact that an anarchist shot the president, which had occurred 7 days earlier. These swings in market values were borne differentially by firms that were particularly sensitive to changes in antitrust enforcement, as were the declines in the market associated with the surprise announcement of Roosevelt's first antitrust suit. Overall, our analysis suggests that expectations of stronger antitrust depressed the valuations of firms more likely to be affected by stricter enforcement. Roosevelt's policies may also have contributed to longer-run contractions in profitability, and no decreases in corporate investment.

often, their retained earnings ("surplus"), allowing us to construct a measure of book equity.

⁴⁷We also take advantage of the extensive data available for railroads by including some additional controls in the profit margin regressions, including region-specific trends, and the lagged value of log assets. We also use assets, rather than the firm's capital, as the measure of firm size that is interacted with the post-1900 indicator.

At the time Roosevelt took office, antitrust enforcement was at a historic low point, and the Supreme Court's *E.C. Knight* decision seemed to foreclose any possibility of pursuing a more aggressive approach, even as the largest wave of mergers in American history was unfolding. Yet when Roosevelt unexpectedly became president, asset price changes indicated that there was considerable discretion available to him in antitrust enforcement, and that he would utilize that discretion. As one of the authors of the Sherman Act, Senator George Edmunds, stated, "What is needed is not, so much, more legislation as competent and earnest administration of the laws that exist" (quoted in Letwin, 1965: 141). Perhaps something similar is true today, in the sense that recent administrations may have not been as aggressive as they could have been in enforcing existing antitrust rules. Though we cannot analyze this directly, our paper raises the question of whether more discretion can also create more variance in outcomes. As stated by Lemos (2011: 700), "If enforcement controls the effective meaning of the law, it matters a great deal who controls enforcement." 48

The results of this paper imply that the accession of Roosevelt to the presidency resulted in significant changes in economic policy. Yet the longer-run consequences of the assassination, relative to the counterfactual of McKinley serving out his second term, are not easy to infer. The end of the Great Merger Movement occurred around 1904, with the Supreme Court's *Northern Securities* decision. Roosevelt's efforts to build on that victory likely inhibited at least some additional merger activity. Under a more business-friendly president, the merger wave may have continued. Growing economic concentration may have been accompanied by growing political influence of big business, potentially threatening both economic growth and democracy (Zingales, 2017). Roosevelt was no radical, and indeed the Tillman Act of 1907 prohibiting direct contributions from corporations to political campaigns was enacted partly in response to the funding that his own 1904 campaign had received from plutocratic interests. But Roosevelt's independent, reform-minded presidency may have prevented some developments that could have been quite harmful for American democracy.

⁴⁸On the changes in Presidential discretion on antitrust policies over time, see Rill and Turner (2014).

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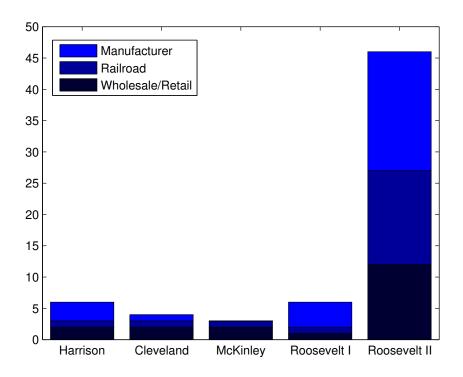


Figure 1: Number of Federal Antitrust Cases Under Different Presidents, by Defendant Type, 1890-1908

Notes: This figure plots the number of antitrust cases instituted by the federal government under different presidents. Suits against labor unions are excluded from the figure. (President Cleveland initiated five such suits). The data are compiled from the summary of cases presented in Commerce Clearing House (1952).

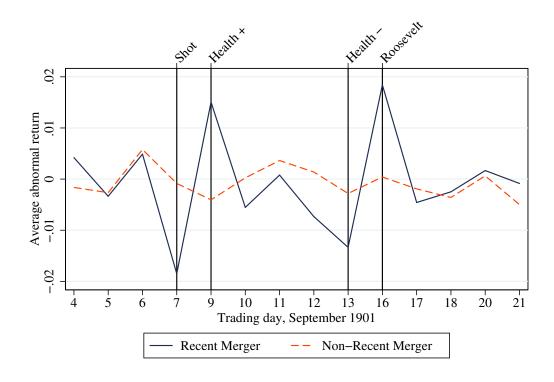


Figure 2: Average Daily Abnormal Returns around McKinley Assassination Event

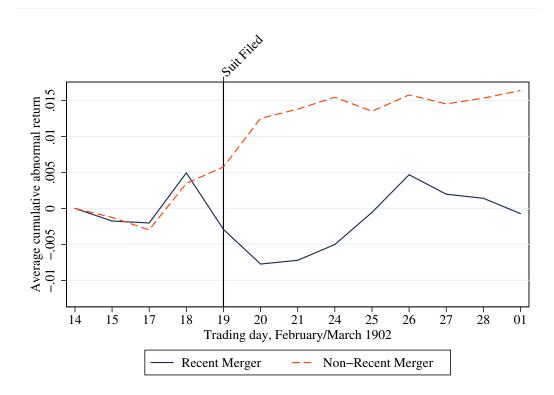


Figure 3: Average Cumulative Abnormal Returns following Northern Securities Suit Announcement

Notes: We set abnormal returns equal to zero for all firms on the 14th and cumulate forward from February 15. For each trading date *t*, the cumulative returns displayed are the group averages of the sum of abnormal returns for each firm from February 15 to date *t*. The attorney general announced his intention to initiate the suit on the 19th after the market was closed.

Table 1: Presidential Assassination Attempts and the Stock Market: Price Response to News of Shootings

| President | Date and Time of Shooting | Date of Trading on News of Shooting | Mean Percent Change, NYSE | Outcome for President |
|-----------|------------------------------------|---|------------------------------|-----------------------------|
| Lincoln | Friday April 14, 1865, 10:25 PM | Monday April 17 (NYSE closed April 15) | -0.7% | Death, Next day, 7:22 AM |
| Garfield | Saturday July 2, 1882, 9:30 AM | Same day | -3.3% | Death, 79 days later |
| McKinley | Friday Sept. 6, 1901, 4:07 PM | Following day | -6.2% | Death, 8 days later |
| Roosevelt | Wed. Feb. 15, 1933, 9:35 PM | Following day | -2.1% | Survived; Was unhurt |
| Kennedy | Friday Nov. 22, 1963, 1:30 PM | Same day (Trading halted 2:07 PM) | -2.8% | Death, Same day, 2:00 PM |
| Ford | Monday Sept. 22, 1975, 6:25 PM | Following day | -0.48% | Survived; Was unhurt |
| Reagan | Mon. March 30, 1981, 2:27 PM | Same day | -0.2% | Survived |

Note: Franklin D. Roosevelt was President-Elect at the time of the shooting attempt on his life. For shootings prior to Roosevelt, the percent change in share prices computed as an equal-weighted index from closing NYSE prices reported in the *New York Times*. For the subsequent shootings, the percent change in share prices is calculated from closing NYSE prices as reported in CRSP. The number of securities for which prices were observed on the day prior to the shooting, and also on the day when trading reflected the news of the shooting, was 16 for the Lincoln assassination, 63 for Garfield, 79 for McKinley, 311 for Roosevelt, 1,144 for Kennedy, 1,468 for Ford and 1,520 for Reagan. All times are reported as EST.

Table 2: Summary Statistics

| | Merger (1) | No Merger (2) | Difference (3) |
|-----------------------------|------------|---------------|----------------|
| | (1) | (2) | (3) |
| Panel A: All Firms | | | |
| I. S. Lines | 20.9583 | 21.6250 | -0.6667 |
| | [14.6420] | [14.4156] | (4.1942) |
| Donor | 0.3333 | 0.0833 | 0.2500** |
| | [0.4815] | [0.2823] | (0.1139) |
| Roosevelt | 0.1667 | 0.1667 | -0.0000 |
| | [0.3807] | [0.3807] | (0.1099) |
| Log(Age) | 1.8477 | 2.8346 | -0.9869*** |
| | [0.9284] | [0.9338] | (0.2688) |
| Log(Capital) | 18.0661 | 17.6861 | 0.3800 |
| | [0.9713] | [0.8244] | (0.2600) |
| Railroad | 0.5833 | 0.5833 | -0.0000 |
| | [0.5036] | [0.5036] | (0.1454) |
| Export/Output | 9.2908 | 9.2235 | 0.0673 |
| - | [4.1793] | [3.2661] | (1.0827) |
| JP Morgan Firm | 0.1250 | 0.0417 | 0.0833 |
| C | [0.3378] | [0.2041] | (0.0806) |
| Observations | 24 | 24 | 48 |
| Donal D. La Lactai al Firma | 01 | | |
| Panel B: Industrial Firms | 2 | 5 7000 | 1 (000 |
| I. S. Lines | 4.1000 | 5.7000 | -1.6000 |
| | [2.5582] | [6.9290] | (2.3357) |
| Observations | 10 | 10 | 20 |

Notes: Columns (1) and (2) report means with standard deviations in brackets. Column (3) reports the difference in means estimated from regressions with a dummy for recent mergers and presents robust standard errors in parentheses.

*** p<0.01, ** p<0.05, * p<0.1.

Table 3: McKinley Assassination Event Analysis using Abnormal Returns, Recent Mergers

| | (1) | (2) | (3) | (4) |
|--------------------------|-----------|-----------|-----------|-----------|
| Merger x Sept. 7 | -0.0175* | -0.0160* | -0.0205** | -0.0205* |
| | (0.0095) | (0.0092) | (0.0094) | (0.0108) |
| Merger x Sept. 9 | 0.0197*** | 0.0191*** | 0.0199*** | 0.0194** |
| - | (0.0061) | (0.0060) | (0.0067) | (0.0084) |
| Merger x Sept. 13 | -0.0141* | -0.0150* | -0.0181** | -0.0114 |
| | (0.0081) | (0.0077) | (0.0082) | (0.0097) |
| Merger x Sept. 16 | 0.0197** | 0.0189** | 0.0216** | 0.0175* |
| | (0.0084) | (0.0082) | (0.0090) | (0.0093) |
| I. S. Lines x Sept. 7 | | 0.0025** | 0.0030*** | 0.0030*** |
| _ | | (0.0010) | (0.0010) | (0.0010) |
| I. S. Lines x Sept. 9 | | -0.0014** | -0.0016** | -0.0016** |
| _ | | (0.0006) | (0.0007) | (0.0007) |
| I. S. Lines x Sept. 13 | | 0.0018 | 0.0011 | 0.0010 |
| | | (0.0020) | (0.0025) | (0.0026) |
| I. S. Lines x Sept. 16 | | -0.0022** | -0.0021* | -0.0021* |
| | | (0.0010) | (0.0011) | (0.0012) |
| Donor x Event Dates | NO | NO | YES | YES |
| Roosevelt x Event Dates | NO | NO | YES | YES |
| Log(Age) x Event Dates | NO | NO | NO | YES |
| Firm, Date Fixed Effects | YES | YES | YES | YES |
| Observations | 603 | 603 | 603 | 603 |
| R-squared | 0.1898 | 0.2181 | 0.2536 | 0.2590 |
| Abnormal Return: | | | | |
| Mean | -0.0007 | -0.0007 | -0.0007 | -0.0007 |
| Std. Dev. | 0.0196 | 0.0196 | 0.0196 | 0.0196 |

Notes: This table presents estimates obtained from different versions of Equation (1). The variable *Merger* indicates differential vulnerability to more aggressive antitrust enforcement. September 7 and 13 were dates with bad news regarding McKinley's health; September 9 and 13 presented good news regarding his health and the likelihood that Roosevelt would follow his agenda, respectively. All regressions include log capital and an indicator for railroads interacted with the event dates (September 7, 9, 13, and 16). Standard errors adjusted for clustering by firm are reported in parentheses.

^{***} p<0.01, ** p<0.05, * p<0.1.

Table 4: Exploring Alternative Policy Differences: Exports and Coal Strike

| Panel A | A: Exports | | Panel B: | Coal Strike Ev | ent |
|--------------------------------|-----------------------|----------------------|-------------------|----------------|-----------------------|
| | (1) | (2) | | (3) | (4) |
| Merger x Sept. 7 | -0.0178* | -0.0208** | Merger | 0.0134 | 0.0082 |
| Merger x Sept. 9 | (0.0094) 0.0201*** | (0.0103) 0.0197** | I. S. Lines | (0.0107) | (0.0110) 0.0042*** |
| Weiger A Sept. 7 | (0.0060) | (0.0076) | I. S. Lines | | (0.0014) |
| Merger x Sept. 13 | -0.0149* | -0.0127 | Donor | | 0.0062 |
| | (0.0078) | (0.0092) | | | (0.0113) |
| Merger x Sept. 16 | 0.0197** | 0.0173* | Roosevelt | | -0.0273** |
| | (0.0084) | (0.0091) | | | (0.0122) |
| Export/Output x Sept. 7 | 0.0010 | 0.0016 | Log(Age) | | -0.0054 |
| | (0.0013) | (0.0013) | _ | | (0.0047) |
| Export/Output x Sept. 9 | -0.0015* | -0.0019** | Constant | -0.0517 | -0.1780* |
| T (0) 10 | (0.0008) | (0.0009) | | (0.0890) | (0.0884) |
| Export/Output x Sept. 13 | 0.0018 | 0.0019* | | | |
| | (0.0011) | (0.0010) | | | |
| Export/Output x Sept. 16 | -0.0001 | -0.0004 | | | |
| I C I in a December December 1 | (0.0015) | (0.0015) | | | |
| I. S. Lines x Event Dates | NO | YES | | | |
| Donor x Event Dates | NO | YES | | | |
| Roosevelt x Event Dates | NO | YES | | | |
| Log(Age) x Event Dates | NO YES | YES YES | | | |
| Firm, Date FEs | IES | YES | | | |
| Observations | 603 | 603 | Observations | 47 | 47 |
| R-squared | 0.2016 | 0.2767 | R-squared | 0.0614 | 0.3145 |
| Abnormal Return: | | | Cum. Abn. Return: | | |
| Mean | -0.0007 | -0.0007 | Mean | 0.0047 | 0.0047 |
| Std. Dev. | 0.0196 | 0.0196 | Std. Dev. | 0.0353 | 0.0353 |

Panel A Notes: This panel presents versions of the regressions presented in Table 3 in which we control for the export share of the sample firms' products interacted with the assassination event dates. The variable *Merger* indicates differential vulnerability to more aggressive antitrust enforcement. All regressions include log capital and an indicator for railroads interacted with the event dates (September 7, 9, 13, and 16). Standard errors adjusted for clustering by firm are reported in parentheses.

Panel B Notes: This table presents results of an event study of the announcement on October 1, 1902, that Roosevelt would intervene in the anthracite coal strike, a significant departure from historical precedent in the approach to labor relations taken by the president. The dependent variable is abnormal returns cumulated from October 1 to October 8, 1902. All regressions also include log capital and an indicator for railroads as controls. Robust standard errors are in parentheses.

^{***} p<0.01, ** p<0.05, * p<0.1.

Table 5: Northern Securities Event Analysis using Cumulative Abnormal Returns

| | (1) | (2) | (3) | (4) |
|-------------------|----------|-----------|-----------|-----------|
| Merger | -0.0130* | -0.0159** | -0.0157** | -0.0159** |
| | (0.0071) | (0.0065) | (0.0066) | (0.0068) |
| I. S. Lines | ` , | 0.0004 | 0.0007 | 0.0006 |
| | | (0.0003) | (0.0005) | (0.0005) |
| Railroad | | , , | -0.0070 | -0.0064 |
| | | | (0.0180) | (0.0181) |
| JP Morgan Firm | | | , , | -0.0074 |
| • | | | | (0.0049) |
| Constant | 0.0968 | 0.1112 | 0.1040 | 0.0911 |
| | (0.0725) | (0.0668) | (0.0761) | (0.0780) |
| Observations | 48 | 48 | 48 | 48 |
| R-squared | 0.1260 | 0.1767 | 0.1779 | 0.1839 |
| Cum. Abn. Return: | | | | |
| Mean | 0.0066 | 0.0066 | 0.0066 | 0.0066 |
| Std. Dev. | 0.0239 | 0.0239 | 0.0239 | 0.0239 |

Notes: This table presents estimates obtained from different versions of Equation (2). We study the effects of the announcement on February 19, 1902, that an antitrust suit would be filed against the Northern Securities Company. The dependent variable is abnormal returns cumulated from February 20 to March 1, 1902. All regressions also include log capital and log age as controls. Robust standard errors are reported in parentheses.

*** p < 0.01, ** p < 0.05, * p < 0.1.

Table 6: Event Analysis Small Sample Tests, Recent Mergers

| | McKinley Assassination (1) | Northern Securities (2) |
|--------------------------------------|----------------------------|-------------------------|
| Return type | 1 day AR | 4 day CAR |
| Number of events | 4 | 1 |
| Binomial Sign Tests | | |
| Number above 95th percentile | 4 | 1 |
| p-value | 0.0001 | 0.0500 |
| Uniform Rank Tests | | |
| Mean rank | 1.25 | 1 |
| Maynes and Rumsey (1993) Z-score | -3.3516 | -1.6846 |
| p-value , | 0.0004 | 0.0460 |
| Joiner and Rosenblatt (1971) Z-score | -4.3156 | -2.2077 |
| p-value | 0.00001 | 0.0136 |

Notes: For this table we first compute the difference in the daily abnormal returns (or 4 day cumulative abnormal returns in the case of Northern Securities) between equal weighted portfolios of recent merger and non–recent merger firms. These differentials for event dates (or windows) are then compared to those of the pre-period to conduct sign and rank tests. A rank of 1 indicates that the differential for the event is greater (less) than the largest (smallest) differential observed in the pre-period when the event is expected to disproportionately benefit (harm) recent merger firms. The cumulation window for the Northern Securities event begins the day prior to the attorney general's announcement.

Table 7: Investment Rates and Profitability under Roosevelt versus McKinley 1896–1905

| | Inves | tment | Return o | n Equity | Profit N (Railroad | |
|---|------------------------------|------------------------------|-----------------------------|------------------------------|------------------------------|------------------------------|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Merger × Post 1900 | 0.025 | 0.066* | -0.016* | -0.014* | -0.050** | -0.031* |
| $Merger \times trend$ | (0.017) | (0.038) -0.009 (0.007) | (0.009) | (0.008) -0.001 (0.001) | (0.024) | (0.018) -0.004 (0.004) |
| Post 1900 × | | (0.007) | | (0.001) | | (0.004) |
| Log size, 1901 | 0.010 (0.015) | 0.010 (0.015) | 0.0001 (0.005) | 0.0001 (0.005) | -0.003 (0.003) | 0.005 (0.023) |
| Log age, 1901 | 0.013 (0.010) | 0.013 (0.010) | -0.007 (0.006) | -0.007 (0.006) | 0.013 (0.012) | 0.014 (0.013) |
| Leverage, 1901 | 0.019 | 0.018 | 0.034 | 0.034 | 0.056 | 0.042 |
| Morgan firm | (0.038) -0.022 (0.019) | (0.038) -0.020 (0.018) | (0.025) 0.007 (0.010) | (0.025) 0.007 (0.010) | (0.051) 0.079* (0.038) | (0.060) 0.079 (0.048) |
| N 01 1 1 100 | | | | | | |
| Mean of dependent variable, 1900 SD of dependent variable, 1900 | |)32)60] | |)88)45] | 0.3 [0.0 | |
| Observations | 315 | 315 | 423 | 423 | 267 | 267 |
| R-squared | 0.269 | 0.277 | 0.799 | 0.799 | 0.700 | 0.740 |
| Firm FE | YES | YES | YES | YES | YES | YES |
| Year FE | YES | YES | YES | YES | YES | YES |

Notes: This table presents estimates for different versions of Equation (3), obtained from a panel of 45 firms over the years 1896-1905. The dependent variables are investment rates, measured as the percentage change in the firm's property account; return on equity, measured as operating income divided by total shareholders' equity; and profit margins, measured as (operating revenues - operating expenses)/(operating revenues). The latter is consistently available for railroads only. Log firm size is defined as log total shareholders' equity in columns (1)-(4), but in the regressions in columns (5) and (6) that use the much more complete data for railroads, we use log total assets. The regressions in columns (5) and (6) also control for region-specific trends, and the lag value of log total assets. For additional details on the construction of these variables, see the Appendix. Standard errors adjusted for clustering by firm are reported in parentheses.

^{***} p<0.01, ** p<0.05, * p<0.1.

Internet Appendix

Table of Contents

| A | Data | Apper | dix | 42 |
|---|------|--------|--|----|
| | A.1 | Variab | le Definitions and Data Sources | 42 |
| | A.2 | Unifor | m Rank Test Z-score Calculations | 4: |
| | A.3 | Full R | esults from Main Assassination Table | 4 |
| | A.4 | Robus | ness Checks | 4 |
| | | A.4.1 | McKinley's Assassination | 4 |
| | | | A.4.1.1 Baseline Results | 4 |
| | | | A.4.1.2 Alternative Policies | 4 |
| | | | A.4.1.3 Alternative Measure–Failed Mergers | 48 |
| | | A.4.2 | Northern Securities | 49 |
| | | | A.4.2.1 Baseline Results | 49 |
| | | A.4.3 | Railroad Accounting Data Regressions | 50 |
| | | A.4.4 | Legal History and Measures of Sensitivity to Antitrust Enforcement | 50 |

A Data Appendix

A.1 Variable Definitions and Data Sources

Stock Returns: The main focus of our analysis are abnormal returns, calculated using daily stock price data and dividend payout information hand-collected from the New York Times on the days surrounding President McKinley's assassination. For each stock, returns (R) are calculated as

$$R_t = \frac{P_t + D_t}{P_{t-1}} - 1,$$

where P represents price, D represents the dividend on days the stock goes ex-dividend, and t indexes trading day. Alpha and beta are estimated for each security using Dimson's (1979) method with three leads and three lags on the market index to address liquidity concerns. The estimating equation follows:

$$R_t = \alpha + \sum_{k=-3}^{3} \beta_{k+4} * I_{t+k} + \epsilon_t,$$

where I represents the relevant market index. For industrial firms we use the Dow Jones Industrial Average and for railroads we use the Dow Jones Transportation Average as the market index. Abnormal returns (AR) are then calculated for each firm as follows:

$$AR_t = R_t - \hat{\alpha} - \sum_{k=-3}^{3} \hat{\beta}_{k+4} * I_t.$$

Pre-period price data was collected for the 75 trading days between May 29 and August 30, 1901, from the *New York Times* in order to estimate $\hat{\alpha}$ and $\hat{\beta}$ for each stock. Due to low liquidity, however, prices are not observed on many days for some firms in the pre-period. Thus, our sample is restricted to stocks for which at least 38 returns are observed in the pre-period. Furthermore, we restrict the sample to firms with an r-squared greater than or equal to 0.32 for railroads and 0.24 for industrial firms in the estimation of $\hat{\alpha}$ and $\hat{\beta}$. These cutoffs represent the 25th percentile of r-squared within each sector after imposing the restriction on the number of returns available in the pre-period. Those firms excluded by this second restriction are missing returns for 22 days of the pre-period on average, while those that pass are missing returns for 8 days on average. These restrictions are made to exclude companies from the analysis for which we cannot obtain reliable estimates of $\hat{\alpha}$ and $\hat{\beta}$.

Merger: This indicator identifies firms that engaged in merger activity in the years prior to the assassination. The variable takes the value one for industrial firms that were incorporated after 1895. We utilize the historical record to verify that all of these firms were the product of consolidations of various smaller firms during the Great Merger Wave. For railroads, this variable takes the value of one for firms with merger activity during the period 1899–1901. This captures mergers occurring after the Supreme Court's 1898 decision on United States v. Joint Traffic Association, which one could argue were more likely to have been enacted to avoid antitrust scrutiny. To identify merger activity we consulted the Moody's Manual and Poor's Manual of the Railroads of the United States and noted railroads that acquired a major stake in another railroad and railroads in which a controlling stake was purchased.

Merger Fail: An indicator variable that takes the value one for firms for which a clear rumor of or plan for a merger with a specific company was announced in *The Commercial and Financial Chronicle* between July 1, 1900, and June 30, 1901, but the merger had not been completed by December 1901, and zero otherwise.

Community of Interest: An indicator variable that takes the value of one for subordinate (or controlled) rail-roads in 'communities of interest' (that is, groups of railroads with common ownership stakes or board seats that could be used to coordinate actions), as identified by Moody (1904). We verify Moody's categorization utilizing *The Moody's Manuals*, and add the Canadian Pacific railroad, which had controlling shares in several American roads that were not part of other groups, as an additional community of interest. We set the variable to zero for industrial firms in the sample.

Donor Firm: An indicator variable that takes the value of one for firms affiliated with J. P. Morgan, the Rockefellers, or Standard Oil (the largest donors to McKinley's presidential campaign), and zero otherwise. The firms designated as affiliated with Morgan are those listed by Moody (1904) as being under "Morgan domination" or "Morgan control," a subset of those with J.P. Morgan & Company partners on their boards. The Standard Oil firms include a number of firms founded by or strongly affiliated with the Rockefellers or Standard Oil as noted by Moody (1904). We include firms connected to the Rockefellers because William Rockefeller, brother of John D. Rockefeller and an executive at Standard Oil, also assisted Hanna in raising funds from other wealthy donors for the McKinley campaign (Rhodes, 1922).

Export/Output: Eysenbach (1976) provides estimates of "exports as a percentage of gross output" for 17 industries for American firms in 1899 in Appendix Table 15. Industrial firms in our sample are matched to these industries using their descriptions in the 1901 edition of the *Moody's Manual*. For railroads, we calculate the percentage of revenues associated with export products among railroads within each of the ICC's ten geographical regions. We first use data from the 1901 Annual Report of the ICC, which details the proportion of different categories of railroad freight by geographical division. We then use tables from the Historical Statistics of the United States to calculate the percentage of each type of product that was exported. For each railroad, the export measure is calculated as: the percentage of revenues accounted for by freight multiplied by the sum of the percentages of freight that was accounted for by different products multiplied by the percentage of each product category that was exported.

Income Statement Lines: A count of the number of substantive lines provided in the income statement of each firm, excluding lines containing intermediate items such as 'subtotals.' Railroads were required to report standardized, detailed income statements to the Interstate Commerce Commission (ICC), which were then published in the ICC's annual reports. Thus, for all railroads this variable takes the value 33, the number of lines on the ICC income statements. For industrial firms we checked the Moody's Manual, the Manual of Statistics Stock Exchange Handbook, and the New York Times guide to investors published in September 1901 for income statements. When income statements were present in more than one source, this variable is a count of the number of lines in the most detailed statement. For firms that did not report an income statement in any of these publications, the income statement lines takes a value of zero. One NYSE company, American Locomotive, was founded in June of 1901, and therefore does not appear in the sources we relied on for financial statements. We have been unable to determine whether investors had access to an income statement at the time of the assassination (or what was known about their dividend rate at that time), and therefore exclude the company from the analysis.

Investment: The percent change in a firm's property account from the previous year. For railroads, this is often denoted "Cost of road and equipment," whereas for industrials, this was typically called "Real estate, plant and equipment" or similar. For some industrials this was listed as several different accounts. This measure is extremely volatile, and to reduce the influence of outliers on our analysis, we trimmed the top 2% of the values of this variable.

JP Morgan Firm: An indicator variable that takes the value of one for firms affiliated with J. P. Morgan, and zero otherwise. These firms are listed by Moody (1904) as being under "Morgan domination" or "Morgan control," a subset of those with J.P. Morgan & Company partners on their boards.

Leverage: Book value of leverage (long-term debt/total assets). Long-term debt includes bonds and other long-term interest-bearing liabilities (e.g. mortgages and equipment trusts), as reported in the firms' balance sheets. For some of our industrial firms, this variable is not available for the year 1901, and we use values from later years for our 1901 leverage × post 1900 interactions. For railroads, this was collected from Poor's Manual of Railroads. For the industrials, the source was the Manual of Statistics Stock Exchange Handbook, and Moody's Manual of Corporation Statistics.

Log(Age): The natural log of one plus 1901 minus the year of incorporation for each firm. The year of incorporation was collected from the 1901 edition of the Moody's Manual.

Log(Assets): The natural log of total assets, as reported in the firms' balance sheets. This is not available for many firm-years for the industrial firms in our sample.

Log(Capital): The natural log of the sum of the preferred and common shares outstanding multiplied by their respective par values. Preferred shares outstanding, common shares outstanding, and the par value of each were collected from the firm's capital stock description from the 1901 edition of the Moody's Manual.

Market Concentration: The relative extent of monopoly by 2 digit SIC codes. Source: Table 39, p 144-147, for Manufacturing and Table 41, p 151, Table 9, p 40, and Table 10, p 41, for Mining, Transportation, Communication, and Public Utilities, Nutter (1951). This variable takes a value of zero for railroads.

Market Share: The percentage of the market controlled by the largest firm in the sector in which the firm operates, where sectors are defined at the 3 and 4 digit SIC codes. Source: Table 37, p 129-141, Nutter (1951). This variable takes a value of zero for railroads.

Profit Margin: Operating profits as a fraction of total revenues. This is calculated as revenues minus operating expenses divided by revenues. This is consistently available only for railroads, and was collected from various editions of *Poor's Manual of Railroads*.

Railroad: An indicator variable that takes the value of one for firms that are listed in the "Eleventh Section: Steam Railroad Securities" of the 1901 Moody's Manual beginning on page 1161, and zero otherwise.

Return on Equity: The ratio of operating profits to total shareholders' equity. For all railroads, this is calculated as revenues minus operating costs, divided by the sum of the book values of common stock, preferred stock, and retained earnings ('surplus'). For industrials, financial reporting was far less standardized. For some firms, rather than operating profits, we have a measure that is closer to net income. And for others, the value of retained earnings was not reported in some years.

Roosevelt: An indicator variable that takes the value of one for firms with directors or executives who attended Harvard with Theodore Roosevelt. Roosevelt matriculated at Harvard in the fall of 1876 and graduated in 1880. While attending Harvard he was a member of the following clubs: Alpha Delta Phi, Delta Kappa Epsilon, Hasty Pudding, Phi Beta Kappa, and Porcellian. We collected the names of the members of these clubs for the graduating classes of 1877–83, who would likely have overlapped with Roosevelt, from club catalogs. We also collected the names of those in the graduating class of 1880 from the Harvard University catalog. Using these lists and a listing of executives and directors of firms collected from the 1901 *Moody's Manual*, we identified firms with connections to Roosevelt.

A.2 Uniform Rank Test Z-score Calculations

In Table 6 we conduct uniform rank tests using two standard calculations of Z-scores. The Maynes and Rumsey (1993) Z-score is given by:

$$Z_U = \sqrt{\frac{12}{m}} \sum_{j=1}^m \left(\frac{h_j}{n+2} - \frac{1}{2} \right), \tag{5}$$

where m is the number of events and h_j is the rank of event j among the n pre-period returns. Meanwhile, the Joiner and Rosenblatt (1971) Z-score is given by:

$$Z_L = \frac{4.91}{\sqrt{m}} \sum_{j=1}^{m} \left[\left(\frac{h_j}{n+2} \right)^{0.14} - \left(1 - \frac{h_j}{n+2} \right)^{0.14} \right].$$
 (6)

The latter more closely approximates the standard normal with small numbers of events.

A.3 Full Results from Main Assassination Table

The tables of results in the paper suppress a number of parameters, to keep the tables from becoming too lengthy. In Appendix Table A.4 we display the coefficients on additional variables of interest for the specifications previously presented in Table 3.

A.4 Robustness Checks

A.4.1 McKinley's Assassination

In this section, we present additional tests to assess the robustness of our main estimated effects for recent mergers discussed in Section 3.1 of the paper.

A.4.1.1 Baseline Results Since our analysis is based on a small number of (48) firms, one potential concern is that the estimated effects are driven by a few salient firms. To address this issue, we re-estimate the interaction effects between Merger_i and the dates with information on McKinley's condition presented in Table 3 by excluding one or two firms at a time, for each of the four specifications in the table. Appendix Figure A.1 presents box plots of the 48 estimated coefficients for each date obtained after dropping one firm at a time. Panel (a), for example, shows that the majority of the estimates are within a 95% confidence interval of the estimated effect reported in Table 3, across all four specifications. Importantly, the estimated effects on the day after McKinley was shot are always negative, and quite similar in magnitude. The results are similar for other relevant dates, as shown in Panels (b)-(d) in Appendix Figure A.1, or when we exclude two firms instead, as shown in Appendix Figure A.2. In no case do the estimated effects flip sign relative to our baseline results, and the magnitudes are quite similar. These results suggest that our main estimates are unlikely to be driven by a small number of firms.

In the paper, we focus the analysis on abnormal stock returns to avoid confounding our estimates with cross-sectional differences in price co-movement with the overall market. This approach reduces the number of observations significantly because it limits the sample to those firms that traded frequently enough to be able to estimate their CAPM betas prior to the assassination, as described in Appendix Section A.1. In Appendix Table A.5 we replicate our main analysis utilizing instead unadjusted returns as the dependent variable. To ensure a minimum level of liquidity, in column (1) we restrict the sample to the firms whose common stocks traded at least 100 shares—the minimum requirement to observe a traded price in *The New York Times*—on each of the four event days. The estimated coefficients get a bit smaller and imprecise, reflecting the fact that low-volume stocks tended to have relatively small unadjusted returns. However, the patterns across dates remains similar—returns of recent merger firms declined on dates of bad news about McKinley's health, and increased when there was positive news about the continuity of his policies. In columns (2) to (5), we restrict the sample to the 48 firms that we include in our baseline specification, but we focus on unadjusted returns instead. The estimated effects are similar to those obtained when using instead abnormal returns, as shown in Table 3.

Our main analysis takes into account differences in firm stability by allowing for differential effects of log capital, a proxy for firm size, on each of the four relevant dates around McKinley's assassination. Here, we replicate this analysis by also including in the regressions other relevant firm characteristics. In Appendix Table A.6 we include log assets. This measure is arguably a more precise proxy for firm size than the log capital variable that we use in all our specifications, but unfortunately we do not observe it for those firms that do not disclose balance sheets. Despite the loss of observations, the estimated effects when we include controls for the value of assets are similar to when we do not.

Our analysis in Section 6 suggests that stronger enforcement of antitrust policies led to a contraction in the profit margins of most affected railroads, and had no substantial differential effects on their investment rates. Though we do not observe strong pre-trends across more and less affected railroads, one could perhaps be concerned that our event study results surrounding McKinley's assassination are driven by firms more likely to have engaged in mergers also experiencing differential profit or investment rates. In Appendix Table A.7 we address this concern by including interactions of both margins and investment, as measured in 1901, with the four relevant dates. Since we only have information on profit and investments for railroads, in columns (1) and (2) of the table, we reproduce the baseline specifications presented in columns (1) and (4) in Table 3, but exclude industrial firms from the sample. Though the sign and magnitude of the estimated coefficients remain largely the same, the estimated effects for September 13 and 16 are no longer statistically significant. But importantly, including controls for the profit margin in columns (3) and (4) and for investment rates in columns (5) and (6) leads to, if anything, stronger effects of recent mergers on abnormal stock returns.

A.4.1.2 Alternative Policies In Panel b of Table 4 we analyze the possibility that our results are driven by a different stance regarding labor relations between McKinley and Roosevelt. Specifically, we find that the returns of firms engaged in recent merger activity did not respond differentially to Roosevelt's intervention in an ongoing coal strike. This analysis suggests that our baseline findings are unlikely to be confounded by differential preferences on labor relations across the administrations. An important limitation of our analysis, however, is that it takes place more than a year after the assassination.

An alternative strategy to address concerns regarding differences in labor policies would be to control for differential intensity in labor activity across firms or across industries—this methodology is akin to our analysis of differential attitudes towards free trade, presented in Panel a of Table 4. Unfortunately, data on unionization rates are not available at the industry level for this period. Instead, we obtain information on strikes and lockouts for various industries from Tables VII and XVIII, respectively, of the Sixteenth Annual Report of the Commissioner of Labor (U.S. Bureau of Labor, 1901).

Appendix Table A.8 presents the results. For ease of comparison, column (1) replicates the baseline results for column (2) of Table 3. These results are primarily robust to including interactions between McKinley's heath news dates and a dummy for whether the firm's industry experienced above the median number of strikes (column [2]) or number of strikers (column [3]). We obtain similar results when we use instead the log number of lockouts (column [4]) or lockoutees (column [5]).

Interestingly, firms that had high labor strife experience an increase in abnormal returns on September 7 and a decline in September 9. These effects perhaps suggest that the market expected a Roosevelt presidency to be beneficial for firms in industries with more organized workers. However, the estimated effects for September 13 and 16 have opposite (albeit insignificant) signs, making any strong conclusions in this regard challenging.

Though these results seem reassuring, there are several important caveats about these data. First, the industry-level data presented in the report was not disaggregated by year. The variables we construct contain data spanning, for most industries, 1881 to 1901. In unreported results, we verify that our results are robust to focusing instead on 1894-1901 data, utilizing earlier labor reports to net out earlier strikes and lockouts. Second, the industry definitions utilized in the labor reports are quite coarse and outdated. While

we do our best to match it to the industries in our sample, the labor strife metrics in our analysis may not be an accurate representation of labor activity in the industries and firms in our analysis. Perhaps more importantly, the Commissioner of Labor does not provide any evidence on the total number of workers, firms or establishments for those industries included in the reports, and this information is extremely difficult to come by from alternative sources for that era. Since the data are simply in levels, one cannot get any sense of the actual intensity of labor strife, which would require estimating a ratio such as, for example, number of strikers relative to total number of employees in the industry. These reasons lead us to believe that these measures of labor activity are at best a very coarse proxy, and we therefore relegate the analysis to the appendix.

McKinley and Roosevelt also differed in their attitudes towards imperialism and naval expansion. Our ability to test this hypothesis directly is limited because there is not obvious variation in exposure across the firms in our sample. One international cause related to imperialism that differed greatly across the administrations is their interest in building the Panama canal. The construction of the canal had been primarily a French endeavor. Shortly after becoming President, Roosevelt acted to buy the rights to the French property and equipment, and to negotiate a Panama treaty with Colombia. This was a prolonged and complex undertaking—the Canal did not open until 1914. Yet if investors anticipated Roosevelt's ideas towards the canal, we would expect those firms more likely to loose business due increased competition from the canal to suffer differentially. To assess this possibility, we create an indicator for the six railroads in our sample that had significant transcontinental freight business—the Northern Pacific; Great Northern; Union Pacific; Chicago Milwaukee, and Saint Paul; Southern Pacific; and Atchison, Topeka, and Santa Fe. In Appendix Table A.9, we include this indicator, interacted with the dates with news on McKinley's health. Interestingly, we find that the market anticipated that these railroads would *benefit* from a change in administration. But more importantly, the main estimated effects on Merger, are unaffected by the inclusion of Panama canal interactions.

A.4.1.3 Alternative Measure–Failed Mergers Our main measure *Merger* captures firms that would have likely suffered differentially from stronger enforcement of antitrust. Alternatively, we can identify a group of firms that would likely have benefited from stricter antitrust enforcement: those that had recently planned, or announced, mergers that ultimately failed to occur. These included competitors of dominant firms in their industry and relatively weak firms that sought to strengthen their positions through mergers, but failed. If these firms faced competition from larger or more efficient rivals, they would have benefited from stricter antitrust enforcement. We create an indicator variable, *Merger Fail*, that identifies any firm in our sample that was mentioned in the year preceding the assassination in the *Commercial and Financial Chronicle*, one of the major business news outlets at that time, as considering, announcing, or being rumored to engage in a merger that failed to occur. While we show that our main results are robust to this alternative measure, it is important to note the limitations of this strategy: we can only observe those failed mergers that progressed enough to be reflected in the press, and there is only a small number of them.

In Appendix Table A.10, we focus on the *Merger Fail* variable as an indicator for firms that would *benefit* from stronger antitrust enforcement. The specifications replicate those in Table 3. Consistent with these firms benefiting from a Roosevelt presidency, their valuations were more resilient on the two days of bad news regarding McKinley's health, when the stock market tanked, but rose by less on the days when it seemed that McKinley would survive or that Roosevelt would follow McKinley's agenda. Although the estimated magnitudes vary somewhat more across the four days, they are all statistically significant, and

⁴⁹For example, Glucose Sugar Refining, a smaller and weaker competitor of the dominant American Sugar Refining, announced in June 1901 that they had begun negotiations with National Starch and other glucose producers to merge into a larger and stronger firm. However, several directors opposed the initial proposal, and the merger initially failed. It was ultimately consummated in 1902 (Dewing, 1914: 87).

robust to the inclusion of our various controls for firm characteristics.

A.4.2 Northern Securities

In this section, we present additional tests to assess the robustness of our main estimated effects for recent mergers discussed in Section 4 of the paper.

A.4.2.1 Baseline Results In Section 4 of the paper, our event-study analysis of the Northern Securities litigation, presented in Table 5, focuses on abnormal returns cumulated from February 20 to March 1, 1902. Here, we show that the results are robust to defining the event over alternative windows. For ease of comparison, column (1) of Appendix Table A.11 reproduces column (4) of Table 5. In columns (2) and (3), we allow for anticipation effects, starting to cumulate returns on February 19 or February 18, respectively. The estimated effects become larger and more precise when we allow for the potential leakage of information to the market prior to Attorney General Knox's announcement.

A potential concern with our main analysis in Table 5 is that other news may have differentially affected firms over a period of nine trading days following the announcement of the suit. To address this problem, we again define the start of the event period on February 20, but instead cumulate returns over shorter windows—eight trading days in column (4) and five trading days in column (5). In both cases, the estimated effects are remarkably similar to those obtained in the baseline estimation, revealing that our main results are not highly sensitive to the window of the analysis.⁵⁰

As we did for the McKinley assassination in A.4.1.1, we next consider the robustness of the estimates for the Northern Securities event to excluding individual firms from the analysis. Appendix Figure A.3 excludes one firm from the sample at a time, whereas Appendix Figure A.4 excludes two. In both cases, the range of estimates remain negative, and the full range of estimates are quite similar in magnitude to the main effects presented for all specifications in Table 5.

To provide further evidence that our results for the Northern Securities event are not artificial outcomes of the data, we next provide a placebo analysis by reestimating our main specifications on a set of placebo dates, similar to the strategy followed by Dube et al. (2011). Specifically, we study the effects of Merger_i on four-data cumulative abnormal returns over mutually exclusive windows during the month surrounding the event, from February 2 to March 1, 1902. Appendix Figure A.5 shows that our four shifted estimates are significant (at the 5% level). The estimated effects are also different from the effects on the cumulative return when the suit filing was announced, on February 19. Though we can only perform this analysis for a short time period due to the constraints hand-collecting stock price data, this placebo analysis suggests that our baseline findings are unlikely to be solely the result of serial correlation in returns.

Finally, in Appendix Table A.12 we replicate the baseline event study for the Northern Securities event utilizing instead cumulative unadjusted returns as the dependent variable, which allows for a larger number of firms to be included in the analysis. To ensure a minimum level of liquidity, in column (1) we restrict the sample to those NYSE-traded firms for which we observe common stock prices on at least three of the nine trading days following, and at least one of the eight trading days prior to, February 19, 1902, (inclusive). The estimated coefficient is a bit larger in magnitude and slightly less precise; still, the results are remarkably similar to the cumulative abnormal returns results presented in Table 5. In columns (2) through (5) we restrict the sample to the 48 firms that we include in our baseline specification, using cumulative unadjusted returns. The estimated effects for firms that engaged in recent mergers are negative and statistically significant, consistent with the main results presented in Table 5.

⁵⁰One firm, American Cotton Oil, is excluded when cumulating from February 20 to February 25 (as shown in Appendix Table A.11 column [5]) because no shares of this firm's common stock are traded during the cumulation period.

A.4.3 Railroad Accounting Data Regressions

Table 7 in the paper presents estimates of the effect of Roosevelt's presidency on railroad accounting outcomes, using a differences-in-differences design. The summary statistics for the data used in these regressions are as follows:

Table A.1: Summary Statistics: Accounting Data

| | Full Sample | Recent Merger | No Recent Merger |
|------------------------|-------------|---------------|------------------|
| | (1) | (2) | (3) |
| | | | |
| Log(Capital), 1901 | 17.533 | 17.294 | 17.804 |
| | [2.718] | [3.659] | [0.859] |
| Log(Age), 1901 | 2.738 | 2.384 | 3.092 |
| | [1.006] | [0.810] | [1.084] |
| Leverage ratio, 1901 | 0.290 | 0.260 | 0.327 |
| | [0.217] | [0.217] | [0.217] |
| Morgan firm | 0.087 | 0.166 | 0 |
| | [0.284] | [0.381] | [0] |
| Investment, 1900 | 0.032 | 0.016 | 0.047 |
| | [0.060] | [0.018] | [0.079] |
| Return on equity, 1900 | 0.088 | 0.084 | 0.092 |
| | [0.045] | [0.035] | [0.053] |
| Profit margin, 1900 | 0.369 | 0.367 | 0.371 |
| - ' | [0.071] | [0.093] | [0.043] |

Notes: Standard deviations in brackets.

In what follows we explore whether those results could be the result of ongoing trends in the industry.

Figure A.6 below presents the 95 percent confidence intervals for estimates of regressions similar to those of Equation (3), only with the recent merger times post-1900 variable replaced by year-by-year interactions, with 1901 excluded. Panel (a) shows the results for investment. There is no strong pattern in the differences in the Roosevelt era relative to the McKinley years, which is consistent with the statistically insignificant results in the table. But the patterns in panel (a) provide clear evidence against any negative effect of Roosevelt's more aggressive antitrust enforcement on firm investment.

Panel (b) shows the results for return on equity. Here the differences in the McKinley years between the recent mergers and the other firms are consistently positive and stable, although of course the 95 percent confidence interval always includes 0. In the Roosevelt years, these differences clearly fall. There is no visual evidence that this change represents the continuation of any sort of ongoing trend.

Finally, panel (c) shows there results for railroads' profit margins. The differences in the McKinley years are consistently higher than the differences in the Roosevelt years. Yet In this case, the differences seem to decline somewhat prior to Roosevelt becoming president. We conclude from this that profit margins did indeed fall, but that one should be cautious in attributing the fall entirely to Roosevelt's policies. Other factors which began prior to Roosevelt's presidency may also have been contributing factors.

A.4.4 Legal History and Measures of Sensitivity to Antitrust Enforcement

Our hypothesis is that McKinley and Roosevelt differed in their stance toward antitrust enforcement. Thus, we would expect changes in the probability of a transition from President McKinley to Roosevelt to differentially affect the market performance of firms that could be subject to antitrust suits, or that could be differentially affected by new antitrust precedents. A significant challenge for our analysis is to identify

these firms. In this section, we expand the paper's description of the history of antitrust doctrine to provide further justification for our variable, which is based on recent merger activity. We also provide evidence that our analysis is robust to considering alternative indicators based on more modern views of anticompetitive behavior, such as market concentration.

Antitrust doctrines were changing rapidly at the turn of the twentieth century. In the 1880s, as industrial interests organized in large combinations, or "trusts," state attorneys first used *quo warranto* suits—legal actions against a corporation incorporated in the state for violating their charters or engaging in illegal acts—but trusts sometimes evaded state courts by reincorporating in friendlier states or adopting different organizational forms. Between 1888 and July 1890, thirteen states passed antitrust statues. These laws typically included stronger penalties than those later introduced by the Sherman Act, and gave most state courts the power to effectively terminate a trust in that state by revoking its charter (Troesken, 2000).

The states' ability to efficiently restrain "bad" combinations by applying common law and their own antitrust laws was limited because the constitution gives Congress the authority to regulate interstate commerce. In July 1890 Congress passed the first federal antitrust law, the Sherman Act. The act banned any contract in restraint of trade, but the broad terms of the law allowed sufficient leeway to distinguish between beneficial forms of cooperation that promoted economic growth and those that suppressed competition (Kovacic and Shapiro, 2000). Which specific practices (such as predatory pricing, price fixing, and many others) were considered violations of the law was evolving over time as the courts were confronted with interpreting the Act. The Addyston Pipe decisions by the Court of Appeals (in 1898) and the Supreme Court (in 1899) helped establish that the Sherman Act would be governed by a rule of reason. In this case, six pipe makers engaged in bid rigging to guarantee that one of the members in the agreement would win projects adjudicated by municipalities using an auction system. The Court of Appeals stated that if the primary purpose of an agreement was to restrain trade, then the agreement was invalid under the law, even if the combination charged reasonable prices. In his opinion, future President Taft argued that the association had acquired the power to charge unreasonable prices, even if they had yet to do so. In Addyston Pipe and Steel Company v. United States (175 U.S. 211 [1899]), the Supreme Court further argued that purely private contracts that directly restrain commerce were in violation of the law. By the time of McKinley's assassination, the courts' interpretation of Section I of the Sherman Act made clear that anticompetitive practices that could restrain trade would be found to be in violation of the law.

The Act, however, did not seem to prohibit large market shares per se. In *United States v. E. C. Knight Co.* (156 U.S. 1 [1895]), the Supreme Court ruled that a series of mergers that gave the American Sugar Refining Company (the *Sugar Trust*) 98% of the nation's sugar refining capacity did not constitute interstate commerce, and was therefore not a violation of the Act. It was not until the Supreme Court's 1911 *Standard Oil* decision that it recognized a high market share as evidence of monopoly in violation of the Sherman Act (Kovacic and Shapiro, 2000: 45).

Given this precedent, it is unlikely that a policy maker concerned with anticompetitive behavior would attempt to prosecute firms simply because they enjoyed large market shares—though clearly these firms would have been at risk if they had employed any illegal tactics in protecting their market shares or exploiting them in other markets.

Another reason that market shares were unlikely to have been utilized as the basis for antitrust enforcement during our period is that systematic market share data were not available. Although rough estimates of manufacturing capacity could be made from firms' disclosures of their capital stock, many of the manufacturing firms in our sample did not disclose their revenues to the public. (Roosevelt in fact called for stronger disclosure by large firms precisely so that regulations could be better designed.) In fact, as we discuss below, most contemporaries described concentration in terms of the level of output.

Our main strategy is, therefore, to focus on firms that had engaged in mergers in the years prior to the assassination, as firms that could no longer collude given the interpretation of Section I often chose instead to merge to be able to preserve monopoly power. The legal history, therefore, validates our choice of focus on merger activity to identify those firms that may have been perceived to be more likely to suffer disproportionately from stronger enforcement of antitrust rules.

Yet one might be concerned that market shares may nonetheless have mattered for investor expectations, particularly if some market participants had access to market share data. In what follows, we further validate our analysis by showing that our results are robust to including controls for available measures of market concentration for our period.

Unfortunately, there is no systematic information on the market shares for the individual firms in our sample at the time of the assassination. Instead, we rely on industry-level measures constructed by Nutter (1951).⁵¹ Appendix Table A.13 replicates the results from Table 3 in the paper but adds as well controls for market concentration interacted with the relevant event dates. In columns (1) and (2), our measure of concentration is the percentage of U.S. output controlled by the largest firm, where markets are defined at the 3 and 4 digit SIC codes as in Nutter (1951: Table 37, p 129-141). In columns (3) and (4), we use instead Nutter's measure of the 'relative extent of monopoly' (Nutter, 1951; Table 39, p 144-147, Table 41, p 151, Table 9, p 40, and Table 10, p 41). Relative extent of monopoly is defined at the two-digit SIC level as the percentage of total value-added accounted for by monopolistic industries at the four-digit SIC level, where monopolistic industries are those in which the four-firm concentration ratio is greater than one half. Importantly, value-added measures can only be obtained for manufacturing industries, as they are constructed from information collected by the Census of Manufactures. In both cases, including our measures of industry concentration does not affect the estimated coefficients for recent mergers.

The estimated coefficients of market share and market concentration are primarily noisily estimated and do not have a clear pattern across McKinley's health news dates. Yet before concluding much from this analysis, a deeper investigation into Nutter's sources reveals severe limitations of these variables. The output measures on which market share is based are collected from a variety of sources. For some industries, such as mining, the data are based on more reliable government reports. In most cases, however, they are obtained from various books written by market analysts or economists, such as Moody (1904) or Dewing (1914). The sources for their numbers, in turn, are often not reported, and most numbers are multiples of five. The sources for their numbers of output concentration are, describing the sources upon which they are based as unreliable (1951: 36). In his view, given the antimonopoly sentiments of the time, the estimates of market shares are likely to overestimate the degree of concentration. The sources upon which Nutter's data are based rarely provide a date or year to which the data corresponds to. This makes it challenging to use for any specific year. Moreover, as Nutter (1951: 46) describes, estimates of the extent of monopoly in a given year may not be representative for years immediately adjacent to it. Given the large data limitations, we urge readers caution when interpreting the estimated effects on these earlier measures of concentration.

Finally, in columns (5) and (6) we take into account that the inability to collude on rates in the aftermath of the Sherman Act led some railroads to form 'communities of interest,' by which a group of investors either individually or through a major railroad would take ownership shares or board seats in other railroads to help coordinate policies. We create an indicator variable for those firms in such a community that were not the dominant firm. These subordinated firms had likely the most to lose from stronger antitrust enforcement, as they would otherwise have to compete with the dominant firms in the group. Again our results show that the

⁵¹To establish the industrial code for the manufacturing firms in our sample, we use the code assigned to each of these firms for the year 1917 by Chandler (1990). When a firm in our sample is not listed by Chandler, we match the industry description provided in the *Moody's Manuals* to the definitions of standardized codes provided by the U.S. Department of Labor (http://www.osha.gov/pls/imis/sic_manual.html). For firms that Chandler scored, the two methods produce the same codes.

⁵²For example, Dewing (1914: 259) states that, during its receivership process, the officers of American Bicycle Company emphasized that the company controlled 70% of the bicycle trade. He gives no specific source for this information, nor is it clear how the managers of the firm arrived to this estimate.

inclusion of this variable does not affect our main estimated effects for recent mergers.

In Appendix Table A.14, we use a similar strategy to investigate the robustness of the estimated effects for the Northern Securities event study to the inclusion of the three alternative measures of market concentration and firm coordination. None of these variables have on their own a significant or sizable differential effect on the cumulative abnormal returns of railroads following the announcement of the suit. Moreover, their inclusion does not affect our main estimated effects for recent mergers. In combination with the results presented in Appendix Table A.13, these results suggest that operating in industries where the dominant firms had large market shares, or utilized common ownership to coordinate policies, was not per se sufficient for being affected differentially by the changes in policies between the two administrations.

The case against Northern Securities Corporation that we analyze in the paper indicates the Act was initially used by Roosevelt's government to forestall mergers that conferred monopoly power, and that investors anticipated so at the time of McKinley's assassination. To provide further validation that our measure of recent mergers is a reasonable indicator of the investors' perception of firms' differential vulnerability to changes in antitrust enforcement at the time, we next relate the our measure to firms that were actually prosecuted for antitrust violations. Specifically, we identify the firms in our sample that were defendants in antitrust cases initiated under Roosevelt. It is important to note that only a small number of firms were subject to legal action, and that some prominent ones, such as Standard Oil, are not included in our sample due to the lack of stock price data. For the sample of 134 railroads and industrial firms that traded on the NYSE after June 1, 1901, and were still listed by the NYT as either active or inactive in September, the likelihood of actual antitrust prosecution was significantly higher (16.7%) for those firms that we categorized as having engaged in merger activity, relative to a much lower 5.8% for those firms with no recent mergers (p-value = 0.072 from univariate regressions with clustered standard errors). The pattern is relatively similar (20.8% versus 8.3%) for the smaller sample of 48 firms for which we can calculate abnormal returns, though the difference is no longer statistically significant.

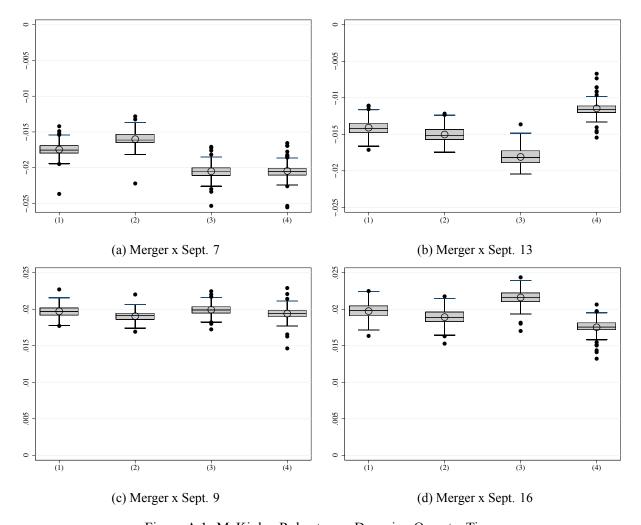


Figure A.1: McKinley Robustness: Dropping One at a Time

Note: This figure shows box plots of the coefficients, on the interaction terms given by the subfigure labels, estimated by dropping each of the 48 firms of our abnormal returns sample in turn and following the specifications shown in Table 3. The labels on the x-axis correspond to the columns of Table 3 and thus indicate the regression specification. The open circles represent the coefficients reported in Table 3 for comparison.

Source: Authors' calculations.

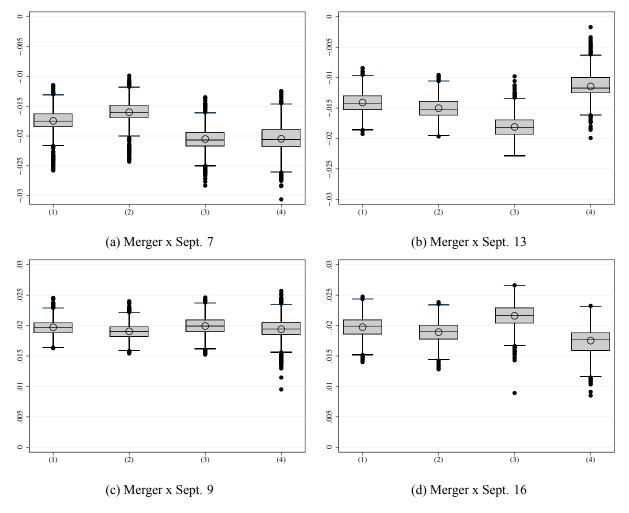


Figure A.2: McKinley Robustness: Dropping Two at a Time

Note: This figure shows box plots of the coefficients, on the interaction terms given by the subfigure labels, estimated by dropping all possible pairwise combinations of the 48 firms of our abnormal returns sample in turn and following the specifications shown in Table 3. The labels on the x-axis correspond to the columns of Table 3 and thus indicate the regression specification. The open circles represent the coefficients reported in Table 3 for comparison. *Source:* Authors' calculations.

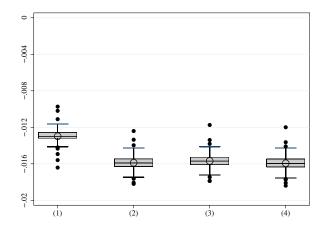


Figure A.3: Northern Securities Robustness: Dropping One at a Time

Note: This figure shows box plots of the coefficients on the recent merger indicator estimated by dropping each of the 48 firms of our cumulative abnormal returns sample in turn and following the specifications shown in Table 5. The labels on the x-axis correspond to the columns of Table 5 and thus indicate the regression specification. The open circles represent the coefficients reported in Table 5 for comparison.

Source: Authors' calculations.

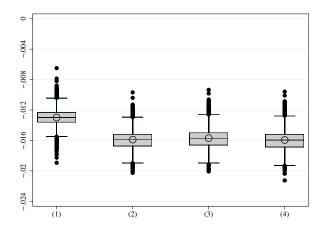


Figure A.4: Northern Securities Robustness: Dropping Two at a Time

Note: This figure shows box plots of the coefficients on the recent merger indicator estimated by dropping all possible pairwise combinations of the 48 firms of our cumulative abnormal returns sample in turn and following the specifications shown in Table 5. The labels on the x-axis correspond to the columns of Table 5 and thus indicate the regression specification. The open circles represent the coefficients reported in Table 5 for comparison. *Source:* Authors' calculations.

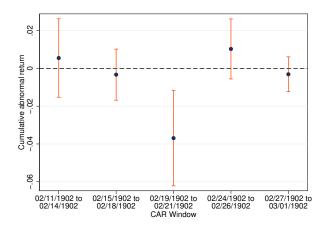


Figure A.5: Northern Securities Robustness: Rolling Windows

Note: This figure shows estimates, with 95 percent confidence intervals, of the differential cumulative abnormal return of recent merger firms over mutually exclusive windows surrounding the announcement on February 19, 1902, that an antitrust suit would be filed against the Northern Securities Company. Estimates are obtained from Equation (2). The dependent variable is abnormal returns cumulated over the range of dates specified on the x-axis. All regressions also include log capital, log age, income statement lines, an indicator for affiliation with J. P. Morgan, and an indicator for railroads as controls.

Source: Authors' calculations.

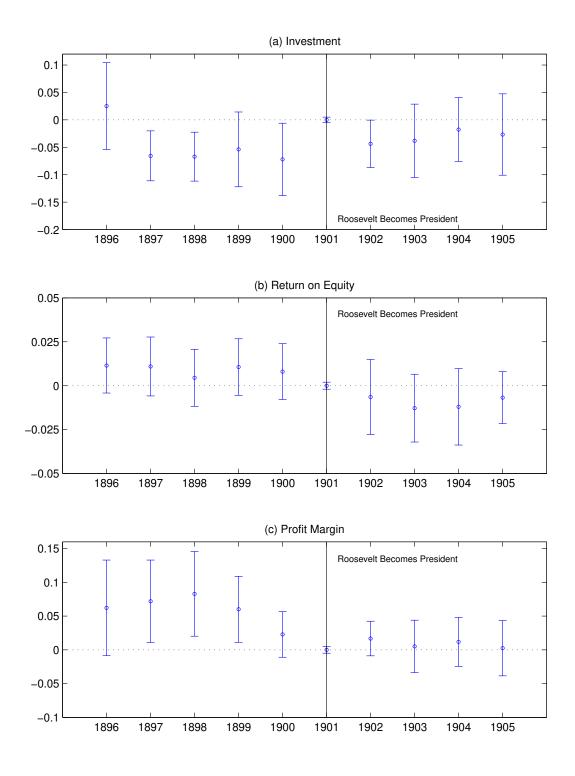


Figure A.6: Year-by-Year Estimates of the Effect of Roosevelt's Presidency on Investment and Profitability

Notes: This figure plots the point estimates and 95 percent confidence intervals from regressions of the same specification as those of columns (1), (3) and (5) of Table 7, but with year-by-year interactions.

Table A.2: Firm Attributes

| Elma Nama | Recent | Failed | D | D 14 | JP Morgan | D - 11 1 |
|------------------------------------|--------|--------|-------|-----------|-----------|----------|
| Firm Name | Merger | Merger | Donor | Roosevelt | Firm | Railroad |
| Amalgamated Copper | Yes | No | Yes | Yes | No | No |
| American Car & Foundry | Yes | No | No | No | No | No |
| American Cotton Oil | No | No | No | No | No | No |
| American Smelting & Refining | Yes | No | Yes | No | No | No |
| American Sugar Refining | No | No | No | Yes | No | No |
| Anaconda Copper Mining | No | No | No | No | No | No |
| Colorado Fuel and Iron | No | Yes | No | No | No | No |
| Columbus & Hocking Coal & Iron | No | No | No | No | No | No |
| Glucose Sugar Refining | Yes | Yes | No | No | No | No |
| International Paper | Yes | No | No | No | No | No |
| International Power | Yes | No | No | No | No | No |
| National Biscuit | Yes | No | No | No | No | No |
| National Lead | No | No | No | No | No | No |
| Pressed Steel Car | Yes | No | No | No | No | No |
| Republic Iron & Steel | Yes | No | Yes | Yes | No | No |
| Tennessee Coal, Iron & Railroad | No | Yes | Yes | No | No | No |
| United States Leather | No | No | No | No | No | No |
| United States Rubber | No | Yes | No | No | No | No |
| United States Steel | Yes | No | Yes | Yes | Yes | No |
| Virginia-Carolina Chemical | No | No | Yes | No | Yes | No |
| Atchison, Topeka & Santa Fe | No | No | No | No | No | Yes |
| Baltimore & Ohio | No | No | No | No | No | Yes |
| Canadian Pacific | No | No | No | No | No | Yes |
| Chesapeake & Ohio | No | No | No | No | No | Yes |
| Chicago Great Western | Yes | No | No | No | No | Yes |
| Chicago, Rock Island & Pacific | No | No | No | No | No | Yes |
| Chicago & Alton | No | No | No | No | No | Yes |
| Chicago, Indianapolis & Louisville | No | Yes | No | Yes | No | Yes |
| Chicago, Milwaukee & St. Paul | Yes | Yes | Yes | No | No | Yes |
| Chicago Terminal Transfer | Yes | No | No | No | No | Yes |
| Delaware & Hudson | No | Yes | No | Yes | No | Yes |
| Erie | Yes | No | Yes | Yes | Yes | Yes |
| Illinois Central | No | No | No | No | No | Yes |
| Louisville & Nashville | No | No | No | Yes | No | Yes |
| Missouri, Kansas & Texas | Yes | Yes | Yes | No | No | Yes |
| Missouri Pacific | Yes | Yes | No | No | No | Yes |
| New York Central & Hudson River | No | No | No | No | No | Yes |
| New York, Ontario & Western | No | No | No | No | No | Yes |
| Pennsylvania | No | No | No | No | No | Yes |
| Reading | Yes | No | No | No | No | Yes |
| Southern | Yes | No | Yes | No | Yes | Yes |
| Southern Pacific | Yes | No | No | No | No | Yes |
| St. Louis & San Francisco | Yes | No | No | No | No | Yes |
| St. Louis Southwestern | No | No | No | No | No | Yes |
| Texas & Pacific | Yes | No | No | No | No | Yes |
| Union Pacific | Yes | Yes | No | No | No | Yes |
| Wabash | Yes | No | No | No | No | Yes |
| Wisconsin Central | Yes | No | No | No | No | Yes |

Notes: See appendix text for a description of the variables.

Table A.3: Firm Attributes (Continued)

| | | Income | Year of | |
|------------------------------------|---------------------------|-----------------|---------------|------------------|
| Firm Name | Capital | Statement Lines | Incorporation | Export/Output |
| Amalgamated Copper | 155,000,000 | 0 | 1899 | 7.121 |
| American Car & Foundry | 60,000,000 | 8 | 1899 | 2.840 |
| American Cotton Oil | 30,435,700 | 7 | 1889 | 13.361 |
| American Smelting & Refining | 100,000,000 | 5 | 1899 | 7.121 |
| American Sugar Refining | 73,936,000 | 0 | 1891 | 13.361 |
| Anaconda Copper Mining | 30,000,000 | 0 | 1895 | 7.121 |
| Colorado Fuel and Iron | 25,000,000 | 23 | 1892 | 5.046 |
| Columbus & Hocking Coal & Iron | 7,000,000 | 6 | 1883 | 5.046 |
| Glucose Sugar Refining | 37,665,600 | 3 | 1897 | 13.361 |
| International Paper | 39,849,500 | 6 | 1898 | 3.082 |
| International Power | 8,000,000 | 4 | 1899 | 2.840 |
| National Biscuit | 53,061,100 | 5 | 1898 | 13.361 |
| National Lead | 29,809,400 | 5 | 1891 | 11.383 |
| Pressed Steel Car | 25,000,000 | 4 | 1899 | 2.840 |
| Republic Iron & Steel | 47,497,900 | 6 | 1899 | 5.046 |
| Tennessee Coal, Iron & Railroad | 22,553,060 | 3 | 1860 | 5.046 |
| United States Leather | 125,139,600 | 0 | 1893 | 4.475 |
| United States Rubber | 47,191,500 | 10 | 1892 | 2.505 |
| United States Steel | 1,014,959,700 | 0 | 1901 | 5.046 |
| Virginia-Carolina Chemical | 24,000,000 | 3 | 1895 | 11.383 |
| Atchison, Topeka & Santa Fe | 216,199,530 | 33 | 1895 | 13.012 |
| Baltimore & Ohio | 104,361,217 | 33 | 1827 | 8.500 |
| Canadian Pacific | 96,171,000 | 33 | 1881 | 11.326 |
| Chesapeake & Ohio | 60,543,100 | 33 | 1878 | 8.648 |
| Chicago Great Western | 40,176,490 | 33 | 1892 | 13.893 |
| Chicago, Rock Island & Pacific | 60,000,000 | 33 | 1880 | 12.797 |
| Chicago & Alton | 39,086,800 | 33 | 1900 | 12.028 |
| Chicago, Indianapolis & Louisville | 15,000,000 | 33 | 1897 | 8.039 |
| Chicago, Milwaukee & St. Paul | 96,397,400 | 33 | 1863 | 13.505 |
| Chicago Terminal Transfer | 30,000,000 | 33 | 1897 | 6.252 |
| Delaware & Hudson | 34,658,200 | 33 | 1823 | 8.920 |
| Erie | 176,240,200 | 33 | 1895 | 8.438 |
| Illinois Central | 76,000,000 | 33 | 1850 | 10.014 |
| Louisville & Nashville | 55,000,000 | 33 | 1850 | 10.830 |
| Missouri, Kansas & Texas | 68,280,300 | 33 | 1896 | 14.075 |
| Missouri Pacific | 76,616,873 | 33 | 1876 | 12.981 |
| New York Central & Hudson River | 115,000,000 | 33 | 1869 | 6.927 |
| New York, Ontario & Western | 58,118,382 | 33 | 1879 | 9.006 |
| Pennsylvania | 202,200,800 | 33 | 1846 | 8.257 |
| Reading | 140,000,000 | 33 | 1871 | 9.124 |
| Southern | 180,000,000 | 33 | 1894 | 9.797 |
| Southern Pacific | | 33 | 1884 | |
| St. Louis & San Francisco | 197,832,148 46,747,400 | 33 | 1896 | 12.103 13.009 |
| | | 33 | 1890 | |
| St. Louis Southwestern | 21,650,000 | | | 14.333 |
| Texas & Pacific | 38,706,000 | 33 | 1874 | 13.334 |
| Union Pacific | 204,511,000 | 33 | 1897 | 13.239 |
| Wabash | 52,000,000 | 33 | 1889 | 7.253 |
| Wisconsin Central | 27,436,000 | 33 | 1899 | 13.321 |

Notes: See appendix text for a description of the variables.

Table A.4: McKinley Assassination Event Analysis using Abnormal Returns, Recent Mergers

| | (1) | (2) | (3) | (4) |
|--------------------------|-----------|-----------|-----------|-----------|
| Merger x Sept. 7 | -0.0175* | -0.0160* | -0.0205** | -0.0205* |
| | (0.0095) | (0.0092) | (0.0094) | (0.0108) |
| Merger x Sept. 9 | 0.0197*** | 0.0191*** | 0.0199*** | 0.0194** |
| | (0.0061) | (0.0060) | (0.0067) | (0.0084) |
| Merger x Sept. 13 | -0.0141* | -0.0150* | -0.0181** | -0.0114 |
| | (0.0081) | (0.0077) | (0.0082) | (0.0097) |
| Merger x Sept. 16 | 0.0197** | 0.0189** | 0.0216** | 0.0175* |
| | (0.0084) | (0.0082) | (0.0090) | (0.0093) |
| I. S. Lines x Sept. 7 | | 0.0025** | 0.0030*** | 0.0030*** |
| | | (0.0010) | (0.0010) | (0.0010) |
| I. S. Lines x Sept. 9 | | -0.0014** | -0.0016** | -0.0016** |
| | | (0.0006) | (0.0007) | (0.0007) |
| I. S. Lines x Sept. 13 | | 0.0018 | 0.0011 | 0.0010 |
| | | (0.0020) | (0.0025) | (0.0026) |
| I. S. Lines x Sept. 16 | | -0.0022** | -0.0021* | -0.0021* |
| | | (0.0010) | (0.0011) | (0.0012) |
| Donor x Sept. 7 | | | 0.0209** | 0.0209** |
| | | | (0.0100) | (0.0101) |
| Donor x Sept. 9 | | | -0.0052 | -0.0052 |
| | | | (0.0076) | (0.0076) |
| Donor x Sept. 13 | | | 0.0158 | 0.0149 |
| | | | (0.0105) | (0.0091) |
| Donor x Sept. 16 | | | -0.0118 | -0.0117 |
| | | | (0.0138) | (0.0132) |
| Roosevelt x Sept. 7 | | | 0.0124 | 0.0124 |
| | | | (0.0078) | (0.0080) |
| Roosevelt x Sept. 9 | | | -0.0062 | -0.0064 |
| | | | (0.0052) | (0.0055) |
| Roosevelt x Sept. 13 | | | -0.0213 | -0.0179 |
| | | | (0.0131) | (0.0131) |
| Roosevelt x Sept. 16 | | | 0.0100 | 0.0091 |
| | | | (0.0129) | (0.0127) |
| Log(Age) x Sept. 7 | | | | 0.0000 |
| | | | | (0.0035) |
| Log(Age) x Sept. 9 | | | | -0.0005 |
| | | | | (0.0035) |
| Log(Age) x Sept. 13 | | | | 0.0062 |
| | | | | (0.0056) |
| Log(Age) x Sept. 16 | | | | -0.0042 |
| | | | | (0.0038) |
| Firm, Date Fixed Effects | YES | YES | YES | YES |
| Observations | 603 | 603 | 603 | 603 |
| R-squared | 0.1898 | 0.2181 | 0.2536 | 0.2590 |

Notes: This table presents estimates obtained from different versions of Equation (1). September 7 and 13 were dates with bad news regarding McKinley's health; September 9 and 13 presented good news regarding his health and the likelihood that Roosevelt would follow his agenda, respectively. All regressions include log capital and an indicator for railroads interacted with the event dates (September 7, 9, 13, and 16). Standard errors adjusted for clustering by firm are reported in parentheses.

^{***} p<0.01, ** p<0.05, * p<0.1.

Table A.5: McKinley Assassination Event Analysis using Unadjusted Returns, Recent Mergers

| | (1) | (2) | (3) | (4) | (5) |
|---------------------------|----------|-----------|-----------|----------|----------|
| Merger x Sept. 7 | -0.0161* | -0.0235** | -0.0235** | -0.0211* | -0.0115 |
| | (0.0095) | (0.0103) | (0.0105) | (0.0109) | (0.0112) |
| Merger x Sept. 9 | 0.0120* | 0.0207*** | 0.0207*** | 0.0180** | 0.0134 |
| | (0.0064) | (0.0066) | (0.0067) | (0.0074) | (0.0086) |
| Merger x Sept. 13 | -0.0134 | -0.0176** | -0.0176** | -0.0161* | -0.0031 |
| - | (0.0081) | (0.0087) | (0.0087) | (0.0088) | (0.0100) |
| Merger x Sept. 16 | 0.0121 | 0.0219** | 0.0220** | 0.0200** | 0.0089 |
| - | (0.0089) | (0.0091) | (0.0092) | (0.0097) | (0.0107) |
| I. S. Lines x Event Dates | NO | NO | YES | YES | YES |
| Donor x Event Dates | NO | NO | NO | YES | YES |
| Roosevelt x Event Dates | NO | NO | NO | YES | YES |
| Log(Age) x Event Dates | NO | NO | NO | NO | YES |
| Firm, Date Fixed Effects | YES | YES | YES | YES | YES |
| Observations | 869 | 603 | 603 | 603 | 603 |
| R-squared | 0.6799 | 0.7636 | 0.7636 | 0.7712 | 0.7822 |
| Return: | | | | | |
| Mean | 0.0012 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| Std. Dev. | 0.0377 | 0.0376 | 0.0376 | 0.0376 | 0.0376 |

Notes: This table presents estimates obtained from different versions of Equation (1). The variable *Merger* indicates differential vulnerability to more aggressive antitrust enforcement. September 7 and 13 were dates with bad news regarding McKinley's health; September 9 and 13 presented good news regarding his health and the likelihood that Roosevelt would follow his agenda, respectively. In Column (1) the sample is restricted to firms whose common stocks traded at least 100 shares on all four event dates. In Columns (2) through (6) the sample is limited to the firms that appear in our abnormal returns analysis. All regressions include log capital and an indicator for railroads interacted with the event dates (September 7, 9, 13, and 16). Standard errors adjusted for clustering by firm are reported in parentheses.

^{***} p<0.01, ** p<0.05, * p<0.1.

Table A.6: McKinley Assassination Event Analysis using Abnormal Returns, Robustness to Assets

| | (1) | (2) | (3) | (4) |
|---------------------------|-----------|-----------|-----------|----------|
| Merger x Sept. 7 | -0.0185* | -0.0164* | -0.0200** | -0.0190* |
| 0 1 | (0.0096) | (0.0092) | (0.0097) | (0.0111) |
| Merger x Sept. 9 | 0.0219*** | 0.0212*** | 0.0225*** | 0.0212** |
| | (0.0061) | (0.0060) | (0.0070) | (0.0084) |
| Merger x Sept. 13 | -0.0116 | -0.0127 | -0.0167* | -0.0111 |
| | (0.0085) | (0.0079) | (0.0086) | (0.0101) |
| Merger x Sept. 16 | 0.0204** | 0.0194** | 0.0229** | 0.0181* |
| | (0.0089) | (0.0087) | (0.0100) | (0.0099) |
| Log(Assets) x Sept. 7 | -0.0039 | -0.0062 | -0.0037 | -0.0039 |
| | (0.0039) | (0.0041) | (0.0041) | (0.0045) |
| Log(Assets) x Sept. 9 | 0.0056** | 0.0066** | 0.0063** | 0.0067** |
| | (0.0026) | (0.0027) | (0.0028) | (0.0032) |
| Log(Assets) x Sept. 13 | 0.0085* | 0.0071 | 0.0048 | 0.0025 |
| | (0.0043) | (0.0048) | (0.0053) | (0.0060) |
| Log(Assets) x Sept. 16 | -0.0004 | 0.0013 | 0.0043 | 0.0057 |
| | (0.0031) | (0.0029) | (0.0055) | (0.0056) |
| I. S. Lines x Event Dates | NO | YES | YES | YES |
| Donor x Event Dates | NO | NO | YES | YES |
| Roosevelt x Event Dates | NO | NO | YES | YES |
| Log(Age) x Event Dates | NO | NO | NO | YES |
| Firm, Date Fixed Effects | YES | YES | YES | YES |
| Observations | 575 | 575 | 575 | 575 |
| R-squared | 0.1982 | 0.2321 | 0.2641 | 0.2701 |
| Abnormal Return: | | | | |
| Mean | -0.0005 | -0.0005 | -0.0005 | -0.0005 |
| Std. Dev. | 0.0197 | 0.0197 | 0.0197 | 0.0197 |

Notes: This table presents estimates obtained from different versions of Equation (1). The variable *Merger* indicates differential vulnerability to more aggressive antitrust enforcement. September 7 and 13 were dates with bad news regarding McKinley's health; September 9 and 13 presented good news regarding his health and the likelihood that Roosevelt would follow his agenda, respectively. All regressions include log capital and an indicator for railroads interacted with the event dates (September 7, 9, 13, and 16). Standard errors adjusted for clustering by firm are reported in parentheses.

^{***} p<0.01, ** p<0.05, * p<0.1.

Table A.7: McKinley Assassination Event Analysis using Abnormal Returns, Recent Mergers Robustness, Railroads Only

| | (1) | (2) | (3) | (4) | (5) | (6) |
|---------------------------|-----------------------|---------------------|-----------------------|---------------------|------------------------|----------------------|
| Manager Cant 7 | 0.0221** | 0.0151 | 0.0205** | 0.0154 | 0.02/7*** | 0.0100* |
| Merger x Sept. 7 | -0.0221** (0.0089) | -0.0151 (0.0107) | -0.0205** (0.0088) | -0.0154 (0.0113) | -0.0267*** (0.0091) | -0.0198* (0.0106) |
| Merger x Sept. 9 | 0.0089) | 0.0107) | 0.0204*** | 0.0113) | 0.0255*** | 0.0100) |
| Meiger A Sept. 9 | (0.0064) | (0.0106) | (0.0204) | (0.0111) | (0.0063) | (0.0112) |
| Merger x Sept. 13 | -0.0082 | -0.0008 | -0.0085 | -0.0009 | -0.0103 | -0.0032 |
| Weiger A Sept. 13 | (0.0081) | (0.0105) | (0.0079) | (0.0102) | (0.0085) | (0.0114) |
| Merger x Sept. 16 | 0.0123 | 0.0048 | 0.0103 | 0.0052 | 0.0156** | 0.0077 |
| | (0.0072) | (0.0085) | (0.0068) | (0.0088) | (0.0074) | (0.0093) |
| Margin x Sept. 7 | () | () | 0.0747 | 0.0267 | () | (******) |
| | | | (0.0875) | (0.1101) | | |
| Margin x Sept. 9 | | | -0.0657 | -0.0488 | | |
| | | | (0.0490) | (0.0716) | | |
| Margin x Sept. 13 | | | -0.0215 | -0.0349 | | |
| | | | (0.0908) | (0.1110) | | |
| Margin x Sept. 16 | | | -0.0977 | -0.0589 | | |
| | | | (0.0578) | (0.0619) | | |
| Investment x Sept. 7 | | | | | 0.1547** | 0.0582 |
| | | | | | (0.0742) | (0.0839) |
| Investment x Sept. 9 | | | | | -0.1298 | -0.1360 |
| | | | | | (0.0768) | (0.0887) |
| Investment x Sept. 13 | | | | | 0.0884 | 0.0431 |
| | | | | | (0.0806) | (0.0739) |
| Investment x Sept. 16 | | | | | -0.1269 | -0.0411 |
| | 110 | | 3.70 | | (0.0928) | (0.0868) |
| I. S. Lines x Event Dates | NO | YES | NO | YES | NO | YES |
| Donor x Event Dates | NO | YES | NO | YES | NO | YES |
| Roosevelt x Event Dates | NO | YES | NO | YES | NO | YES |
| Dividend x Event Dates | NO NO | YES | NO NO | YES | NO NO | YES |
| Log(Age) x Event Dates | NO YES | YES | NO YES | YES | NO YES | YES |
| Firm, Date Fixed Effects | YES | YES | YES | YES | YES | YES |
| Observations | 381 | 381 | 381 | 381 | 367 | 367 |
| R-squared | 0.2984 | 0.3685 | 0.3116 | 0.3719 | 0.3294 | 0.3880 |
| Abnormal Return: | | | | | | |
| Mean | -0.0003 | -0.0003 | -0.0003 | -0.0003 | -0.0002 | -0.0002 |
| Std. Dev. | 0.0167 | 0.0167 | 0.0167 | 0.0167 | 0.0170 | 0.0170 |

Notes: This table presents estimates obtained from different versions of Equation (1). The variable *Merger* indicates differential vulnerability to more aggressive antitrust enforcement. September 7 and 13 were dates with bad news regarding McKinley's health; September 9 and 13 presented good news regarding his health and the likelihood that Roosevelt would follow his agenda, respectively. All regressions include log capital interacted with the event dates (September 7, 9, 13, and 16). Standard errors adjusted for clustering by firm are reported in parentheses.

*** p < 0.01, ** p < 0.05, * p < 0.1.

Table A.8: McKinley Assassination Event Analysis using Abnormal Returns, Recent Mergers, Robustness to Labor Strife

| | (1) | (2) | (3) | (4) | (5) |
|---------------------------|-----------|--------------|---------------|---------------|-----------------|
| Merger x Sept. 7 | -0.0160* | -0.0150* | -0.0115 | -0.0180** | -0.0202** |
| | (0.0092) | (0.0089) | (0.0091) | (0.0085) | (0.0079) |
| Merger x Sept. 9 | 0.0191*** | 0.0185*** | 0.0165*** | 0.0170*** | 0.0187*** |
| | (0.0060) | (0.0062) | (0.0057) | (0.0062) | (0.0059) |
| Merger x Sept. 13 | -0.0150* | -0.0113 | -0.0109 | -0.0115 | -0.0121 |
| | (0.0077) | (0.0082) | (0.0082) | (0.0087) | (0.0088) |
| Merger x Sept. 16 | 0.0189** | 0.0158* | 0.0160* | 0.0178* | 0.0183* |
| | (0.0082) | (0.0089) | (0.0086) | (0.0091) | (0.0092) |
| Labor Strife x Sept. 7 | | 0.0369** | 0.0366* | 0.0234** | 0.0563*** |
| - | | (0.0149) | (0.0192) | (0.0103) | (0.0104) |
| Labor Strife x Sept. 9 | | -0.0122 | -0.0280** | -0.048 | -0.0235** |
| - | | (0.0122) | (0.0112) | (0.0069) | (0.0098) |
| Labor Strife x Sept. 13 | | -0.0057 | 0.0090 | -0.0018 | 0.0064 |
| - | | (0.0242) | (0.0178) | (0.0130) | (0.0242) |
| Labor Strife x Sept. 16 | | 0.0036 | 0.0039 | -0.0012 | -0.0073 |
| _ | | (0.0320) | (0.0207) | (0.0176) | (0.0265) |
| I. S. Lines x Event Dates | YES | YES | YES | YES | YES |
| Firm, Date Fixed Effects | YES | YES | YES | YES | YES |
| | | | | | |
| Observations | 603 | 558 | 558 | 512 | 512 |
| R-squared | 0.2181 | 0.2495 | 0.2607 | 0.2478 | 0.2724 |
| Labor Strife Variable | N/A | High Strikes | High Strikers | Log(Lockouts) | Log(Lockoutees) |

Notes: This table presents estimates obtained from different versions of Equation (1). The variable Merger indicates differential vulnerability to more aggressive antitrust enforcement. September 7 and 13 were dates with bad news regarding McKinley's health; September 9 and 13 presented good news regarding his health and the likelihood that Roosevelt would follow his agenda, respectively. We use four different indicators for labor strife (High Strikes, High Strikers, Log(Lockoutes)) based on data for the period 1880–1901. The indicator High Strikes identifies firms in industries with above the median number of strikes, whereas High Strikers identifies firms in industries with above the median number of strikers, over the relevant period of time. The variable Log(Lockoutees) is the natural log of the number of employees locked out of firms in the relevant industry group. All regressions include log capital and an indicator for railroads interacted with the event dates (September 7, 9, 13, and 16). Standard errors adjusted for clustering by firm are reported in parentheses.

^{***} p<0.01, ** p<0.05, * p<0.1.

Table A.9: McKinley Assassination Event Analysis using Abnormal Returns, Recent Mergers, Robustness to Panama Canal

| | (1) | (2) | (3) | (4) |
|---------------------------|------------|------------|------------|-----------|
| Merger x Sept. 7 | -0.0192** | -0.0177* | -0.0227** | -0.0212** |
| | (0.0093) | (0.0091) | (0.0089) | (0.0102) |
| Merger x Sept. 9 | 0.0205*** | 0.0198*** | 0.0209*** | 0.0200** |
| | (0.0061) | (0.0060) | (0.0067) | (0.0084) |
| Merger x Sept. 13 | -0.0152* | -0.0159** | -0.0189** | -0.0115 |
| | (0.0080) | (0.0076) | (0.0081) | (0.0091) |
| Merger x Sept. 16 | 0.0215** | 0.0206** | 0.0231** | 0.0176** |
| | (0.0082) | (0.0081) | (0.0087) | (0.0082) |
| Panama x Sept. 7 | 0.0269*** | 0.0237*** | 0.0304*** | 0.0249** |
| | (0.0074) | (0.0073) | (0.0089) | (0.0104) |
| Panama x Sept. 9 | -0.0136* | -0.0120 | -0.0147* | -0.0121 |
| | (0.0077) | (0.0077) | (0.0083) | (0.0088) |
| Panama x Sept. 13 | 0.0225*** | 0.0215*** | 0.0182** | 0.0137 |
| | (0.0068) | (0.0069) | (0.0068) | (0.0095) |
| Panama x Sept. 16 | -0.0309*** | -0.0284*** | -0.0277*** | -0.0211** |
| | (0.0074) | (0.0075) | (0.0077) | (0.0097) |
| I. S. Lines x Event Dates | NO | YES | YES | YES |
| Donor x Event Dates | NO | NO | YES | YES |
| Roosevelt x Event Dates | NO | NO | YES | YES |
| Dividend x Event Dates | NO | NO | NO | YES |
| Log(Age) x Event Dates | NO | NO | NO | YES |
| Firm, Date Fixed Effects | YES | YES | YES | YES |
| Observations | 603 | 603 | 603 | 603 |
| R-squared | 0.2206 | 0.2436 | 0.2814 | 0.3208 |
| Abnormal Return: | | | | |
| Mean | -0.0007 | -0.0007 | -0.0007 | -0.0007 |
| Std. Dev. | 0.0196 | 0.0196 | 0.0196 | 0.0196 |

Notes: This table presents estimates obtained from different versions of Equation (1). The variable *Merger* indicates differential vulnerability to more aggressive antitrust enforcement. The variable *Panama* is an indicator that takes a value of one for the six railroads in our sample that had significant transcontinental freight business, and zero otherwise. September 7 and 13 were dates with bad news regarding McKinley's health; September 9 and 13 presented good news regarding his health and the likelihood that Roosevelt would follow his agenda, respectively. All regressions include log capital and an indicator for railroads interacted with the event dates (September 7, 9, 13, and 16). Standard errors adjusted for clustering by firm are reported in parentheses.

^{***} p<0.01, ** p<0.05, * p<0.1.

Table A.10: McKinley Assassination Event Analysis using Abnormal Returns, Failed Mergers

| | (1) | (2) | (3) | (4) |
|---------------------------|------------|------------|------------|------------|
| Merger Fail x Sept. 7 | 0.0292*** | 0.0251*** | 0.0223*** | 0.0207*** |
| 2 1 | (0.0081) | (0.0079) | (0.0076) | (0.0071) |
| Merger Fail x Sept. 9 | -0.0168** | -0.0145** | -0.0154** | -0.0133** |
| | (0.0065) | (0.0066) | (0.0060) | (0.0057) |
| Merger Fail x Sept. 13 | 0.0256*** | 0.0253** | 0.0244*** | 0.0234*** |
| | (0.0092) | (0.0095) | (0.0088) | (0.0074) |
| Merger Fail x Sept. 16 | -0.0351*** | -0.0325*** | -0.0346*** | -0.0318*** |
| | (0.0084) | (0.0083) | (0.0096) | (0.0088) |
| I. S. Lines x Event Dates | NO | YES | YES | YES |
| Donor x Event Dates | NO | NO | YES | YES |
| Roosevelt x Event Dates | NO | NO | YES | YES |
| Log(Age) x Event Dates | NO | NO | NO | YES |
| Firm, Date Fixed Effects | YES | YES | YES | YES |
| Observations | 603 | 603 | 603 | 603 |
| R-squared | 0.2252 | 0.2369 | 0.2572 | 0.2803 |
| Abnormal Return: | | | | |
| Mean | -0.0007 | -0.0007 | -0.0007 | -0.0007 |
| Std. Dev. | 0.0196 | 0.0196 | 0.0196 | 0.0196 |

Notes: This table presents estimates obtained from different versions of Equation (1). The variable *Merger Fail* indicates firms that would be expected to benefit from more aggressive antitrust enforcement. September 7 and 13 were dates with bad news regarding McKinley's health; September 9 and 13 presented good news regarding his health and the likelihood that Roosevelt would follow his agenda, respectively. All regressions include log capital and an indicator for railroads interacted with the event dates (September 7, 9, 13, and 16). Standard errors adjusted for clustering by firm are reported in parentheses.

^{***} p<0.01, ** p<0.05, * p<0.1.

Table A.11: Northern Securities Event Analysis using Cumulative Abnormal Returns

| | (1) | (2) | (3) | (4) | (5) |
|-------------------|------------|------------|------------|------------|------------|
| Merger | -0.0159** | -0.0298*** | -0.0314** | -0.0147** | -0.0150* |
| • | (0.0068) | (0.0098) | (0.0117) | (0.0064) | (0.0075) |
| I. S. Lines | 0.0006 | 0.0009 | 0.0012 | 0.0003 | -0.0004 |
| | (0.0005) | (0.0009) | (0.0011) | (0.0004) | (0.0009) |
| Railroad | -0.0064 | 0.0001 | -0.0050 | 0.0043 | 0.0180 |
| | (0.0181) | (0.0301) | (0.0379) | (0.0137) | (0.0277) |
| JP Morgan Firm | -0.0074 | -0.0065 | -0.0085 | -0.0058 | -0.0176** |
| | (0.0049) | (0.0071) | (0.0083) | (0.0053) | (0.0068) |
| Constant | 0.0911 | 0.1174 | 0.2077 | 0.1089 | 0.0411 |
| | (0.0780) | (0.1183) | (0.1601) | (0.0723) | (0.0848) |
| Observations | 48 | 48 | 48 | 48 | 47 |
| R-squared | 0.1839 | 0.2723 | 0.3068 | 0.2760 | 0.1738 |
| Cum. Abn. Return: | | | | | |
| Mean | 0.0066 | 0.0040 | 0.0107 | 0.0071 | 0.0053 |
| Std. Dev. | 0.0239 | 0.0313 | 0.0365 | 0.0206 | 0.0232 |
| CAR Window: | | | | | |
| Begin Date | 02/20/1902 | 02/19/1902 | 02/18/1902 | 02/20/1902 | 02/20/1902 |
| End Date | 03/01/1902 | 03/01/1902 | 03/01/1902 | 02/28/1902 | 02/25/1902 |

Notes: This table presents estimates obtained from different versions of Equation (2). We study the effects of the announcement on February 19, 1902, that an antitrust suit would be filed against the Northern Securities Company. The dependent variable is abnormal returns cumulated over the range of dates specified under CAR Window. All regressions also include log capital and log age as controls. Robust standard errors are reported in parentheses. *** p < 0.01, ** p < 0.05, * p < 0.1.

Table A.12: Northern Securities Event Analysis using Cumulative Unadjusted Returns

| | (1) | (2) | (3) | (4) | (5) |
|----------------|----------|-----------|-----------|-----------|-----------|
| Merger | -0.0180* | -0.0191** | -0.0173** | -0.0166** | -0.0170** |
| C | (0.0106) | (0.0077) | (0.0070) | (0.0071) | (0.0073) |
| I. S. Lines | , , | ` , | -0.0003 | 0.0005 | 0.0005 |
| | | | (0.0003) | (0.0006) | (0.0006) |
| Railroad | | | | -0.0234 | -0.0225 |
| | | | | (0.0191) | (0.0191) |
| JP Morgan Firm | | | | | -0.0108* |
| | | | | | (0.0056) |
| Constant | 0.1893** | 0.0840 | 0.0751 | 0.0510 | 0.0321 |
| | (0.0922) | (0.0670) | (0.0688) | (0.0784) | (0.0813) |
| Observations | 90 | 48 | 48 | 48 | 48 |
| R-squared | 0.1843 | 0.1811 | 0.1977 | 0.2094 | 0.2204 |
| Cum. Return: | | | | | |
| Mean | -0.0019 | -0.0106 | -0.0106 | -0.0106 | -0.0106 |
| Std. Dev. | 0.0357 | 0.0257 | 0.0257 | 0.0257 | 0.0257 |

Notes: This table presents estimates obtained from different versions of Equation (2). We study the effects of the announcement on February 19, 1902, that an antitrust suit would be filed against the Northern Securities Company. The dependent variable is unadjusted returns cumulated from February 20 to March 1, 1902. All regressions also include log capital and log age as controls. Robust standard errors are reported in parentheses. *** p < 0.01, ** p < 0.05, * p < 0.1.

Table A.13: McKinley Assassination Event Analysis using Abnormal Returns, Robustness to Market Share, Market Concentration, and Community of Interest

| | (1) | (2) | (3) | (4) | (5) | (6) |
|----------------------------------|-----------|-----------|-----------|----------|----------|--------------------|
| Margar v Sant 7 | -0.0167* | -0.0199* | -0.0172** | -0.0197* | -0.0174 | -0.0198* |
| Merger x Sept. 7 | (0.0088) | (0.0199) | (0.0080) | (0.0197) | (0.0109) | (0.0115) |
| Merger x Sept. 9 | 0.0199*** | 0.0102) | 0.0199*** | 0.0100) | 0.0109) | 0.0113) |
| Weiger x Sept. 9 | (0.0060) | (0.0084) | (0.0058) | (0.0085) | (0.0071) | (0.0193°) |
| Merger x Sept. 13 | -0.0150* | -0.0121 | -0.0141* | -0.0116 | -0.0134 | -0.0096 |
| Weiger A Sept. 15 | (0.0082) | (0.0121) | (0.0081) | (0.0097) | (0.0090) | (0.0108) |
| Merger x Sept. 16 | 0.0197** | 0.0175* | 0.0201** | 0.0177* | 0.0191** | 0.0169* |
| Wieiger A Sept. 10 | (0.0083) | (0.0093) | (0.0080) | (0.0092) | (0.0092) | (0.0097) |
| Market Share x Sept. 7 | -0.1673* | -0.1295** | (0.0000) | (0.00)2) | (0.0052) | (0.00)1) |
| warket Share A Sept. 7 | (0.0837) | (0.0536) | | | | |
| Market Share x Sept. 9 | 0.0649 | 0.0525 | | | | |
| warnet share a sept. | (0.0460) | (0.0490) | | | | |
| Market Share x Sept. 13 | 0.0836 | 0.0850 | | | | |
| | (0.0658) | (0.0659) | | | | |
| Market Share x Sept. 16 | 0.0367 | 0.0366 | | | | |
| | (0.1051) | (0.0960) | | | | |
| Market Concentration x Sept. 7 | (, , , , | () | 0.1166** | 0.0795* | | |
| 1 | | | (0.0460) | (0.0430) | | |
| Market Concentration x Sept. 9 | | | -0.0326 | -0.0124 | | |
| • | | | (0.0284) | (0.0311) | | |
| Market Concentration x Sept. 13 | | | -0.0038 | -0.0107 | | |
| • | | | (0.0490) | (0.0370) | | |
| Market Concentration x Sept. 16 | | | -0.0297 | -0.0119 | | |
| • | | | (0.0460) | (0.0426) | | |
| Community of Interest x Sept. 7 | | | , | , | -0.0004 | -0.0054 |
| • | | | | | (0.0114) | (0.0118) |
| Community of Interest x Sept. 9 | | | | | 0.00001 | 0.0011 |
| • | | | | | (0.0079) | (0.0080) |
| Community of Interest x Sept. 13 | | | | | -0.0036 | -0.0077 |
| | | | | | (0.0100) | (0.0113) |
| Community of Interest x Sept. 16 | | | | | 0.0033 | 0.0061 |
| | | | | | (0.0090) | (0.0095) |
| I. S. Lines x Event Dates | NO | YES | NO | YES | NO | YES |
| Donor x Event Dates | NO | YES | NO | YES | NO | YES |
| Roosevelt x Event Dates | NO | YES | NO | YES | NO | YES |
| Log(Age) x Event Dates | NO | YES | NO | YES | NO | YES |
| Firm, Date Fixed Effects | YES | YES | YES | YES | YES | YES |
| Observations | 603 | 603 | 603 | 603 | 603 | 603 |
| R-squared | 0.2175 | 0.2768 | 0.2249 | 0.2696 | 0.1904 | 0.2620 |

Notes: This table presents estimates obtained from different versions of Equation (1). September 7 and 13 were dates with bad news regarding McKinley's health; September 9 and 13 presented good news regarding his health and the likelihood that Roosevelt would follow his agenda, respectively. Specifications in columns (1) and (2) include a measure of market share interacted with event dates. Specifications in columns (3) and (4) include a measure of market concentration interacted with event dates. Specifications in columns (5) and (6) include an indicator for being a subordinate member of a railroad community of interest interacted with event dates. All regressions include log capital and an indicator for railroads interacted with the event dates (September 7, 9, 13, and 16). Standard errors adjusted for clustering by firm are reported in parentheses.

^{***} p<0.01, ** p<0.05, * p<0.1.

Table A.14: Northern Securities Event Analysis using Cumulative Abnormal Returns, Robustness to Market Share, Market Concentration, and Community of Interest

| | (1) | (2) | (3) | (4) | (5) | (6) |
|-----------------------|-----------|-----------|-----------|-----------|-----------|-----------|
| Merger | -0.0155** | -0.0159** | -0.0148** | -0.0156** | -0.0160** | -0.0170** |
| · · | (0.0064) | (0.0068) | (0.0070) | (0.0066) | (0.0069) | (0.0071) |
| Market Share | -0.0165 | 0.0064 | , | , | , | , |
| | (0.0150) | (0.0691) | | | | |
| Market Concentration | , | , | -0.0165 | -0.0103 | | |
| | | | (0.0113) | (0.0317) | | |
| Community of Interest | | | ` , | , , | 0.0092 | 0.0056 |
| • | | | | | (0.0073) | (0.0067) |
| Constant | 0.1251* | 0.0889 | 0.1266* | 0.0937 | 0.0971 | 0.0809 |
| | (0.0736) | (0.0804) | (0.0733) | (0.0789) | (0.0687) | (0.0737) |
| I. S. Lines | NO | YES | NO | YES | NO | YES |
| Railroad | NO | YES | NO | YES | NO | YES |
| JP Morgan Firm | NO | YES | NO | YES | NO | YES |
| Observations | 48 | 48 | 48 | 48 | 48 | 48 |
| R-squared | 0.1630 | 0.1843 | 0.1619 | 0.1869 | 0.1553 | 0.1910 |
| Cum. Abn. Return: | | | | | | |
| Mean | 0.0066 | 0.0066 | 0.0066 | 0.0066 | 0.0066 | 0.0066 |
| Std. Dev. | 0.0239 | 0.0239 | 0.0239 | 0.0239 | 0.0239 | 0.0239 |

Notes: This table presents estimates obtained from different versions of Equation (2). We study the effects of the announcement on February 19, 1902, that an antitrust suit would be filed against the Northern Securities Company. The dependent variable is abnormal returns cumulated from February 20 to March 1, 1902. Specifications in columns (1) and (2) include a measure of market share. Specifications in columns (3) and (4) include a measure of market concentration. Specifications in columns (5) and (6) include an indicator for being a subordinate member of a railroad community of interest. All regressions also include log capital and log age as controls. Robust standard errors are reported in parentheses.

^{***} p<0.01, ** p<0.05, * p<0.1.