Dual Random Utility Maximisation*

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Abstract

Dual Random Utility Maximisation (dRUM) is Random Utility Maximisation when utility depends on only two states. This class has many relevant behavioural interpretations and practical applications. We show that dRUM is (generically) the only stochastic choice rule that satisfies Regularity and two new properties: Constant Expansion (if the choice probability of an alternative is the same across two menus, then it is the same in the merged menu), and Negative Expansion (if the choice probability of an alternative is less than one and differs across two menus, then it vanishes in the merged menu). We extend the theory to menu-dependent state probabilities. This accommodates prominent violations of Regularity such as the attraction, similarity and compromise effects.

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1 Introduction

In Random Utility Maximisation (RUM) choices result from the maximisation of a utility that depends on random states, where the probabilities of the states are independent of the menu the decision maker is facing. In this paper we fully characterise and extend RUM for the case in which the states are exactly two (the dual RUM or dRUM in short) and the number of alternatives is finite. We give a simple picture of the choice behaviour of a dual random utility maximiser, and we also extend the theory to account for the state probabilities to depend on the menu, a feature that -we argue- is compelling in many circumstances.

Dual RUM has many interesting applications, both when stochastic choices emanate from an individual and when they describe instead random draws from a population. We list a few:

- *Dual-self.* For individual behaviour, a leading interpretation is in terms of a 'dual-self' process: one state is a 'cool' state, in which a long-run utility is maximised, while the other is a 'hot' state, in which a myopic self, subject to short-run impulses (such as temptation), takes control. Indeed, the menu-dependent version of dRUM describes the second-stage choices of the dual-self model characterised by Chatterjee and Krishna [5], and the menu-independent version is the one used earlier by Eliaz and Spiegler [10]. Correspondingly, our results can be seen as a direct characterisation of this model in terms of choice *from* menus rather than in terms of preferences *over* menus as in [5].¹
- *Hidden conditions of choice.* In general, the states may correspond to other choice-relevant conditions that are hidden to the observer. For example, the effect of time pressure on decision making has recently attracted attention.² The manager

¹Eliaz and Spiegler [10] do not focus on characterisation but rather explore the rich implications of this probabilistic notion of naivete in a contract-theoretic framework, where a principal chooses the optimal menu of contracs to offer to partially naive agents.

²See e.g. Reutskaya et al. [21] for a recent study of this effect on search behaviour in a shoppping environment.

of a store may wonder whether the variability in buying patterns fits a 'dual' form behaviour due to presence or absence of time pressure at the time of purchase. It is likely that scanner data on the repeated purchases of consumers are available to the manager. The same applies for which products were available at the time of each purchase, so that a stochastic choice function can be constructed. It is, however, unlikely that the manager can learn whether or not the consumer was in a hurry at the time of purchase. Our results allow a direct test of the dual behaviour hypothesis.

- *Hidden population heterogeneity.* The dRUM may represent the variability of choice due to a binary form of hidden population heterogeneity. The states correspond to two types of individuals in the population. Interesting examples of this kind of heterogeneity are when the population is split, in unknown proportions, into high and low cognitive ability, strategic and non-strategic, 'fast and slow' or 'instinctive and contemplative' types as suggested by Rubinstein ([22],[23]). Such binary classifications are natural and widespread.³ Our results clarify what type of consistency the choices of a population split in this way will exhibit.
- *Household decisions*. Another main application to a multi-person situation concerns decisions units composed of two agents, notably a household, for which there is uncertainty about the exact identity of the decision maker. Household purchases can be observed through standard consumption data, but typically it is not known which of the two partners made the actual purchase decision on any given occasion. Dual RUM constitutes the basis for a 'random dictatorship' model of household decisions that could complement the 'collective' model (Chiappori [7]; Cherchye, De Rock and Vermeulen [6]).⁴
- Normativity vs selfishness. Many situations of choice present a naturally binary

³See also Caplin and Martin [4], who consider agents who can *choose* whether to make automatic (fast) or considered (slow) choices, depending on an attentional cost.

⁴de Clippel and Eliaz [8] propose a dual-self model that can be thought of as a model of intrahousehold bargaining. Interestingly, this model exhibits the attraction and compromise effects that we study later in the paper.

conflict between a 'normative' mode and a 'selfish' mode of decision. This conflict is both introspectively clear and is observed experimentally, for example, in dictator games (Frolich, Oppenheimer and Moore [15]). This dichotomy can pertain both to a typology of individuals and to a single conflicted individual.

The models we study, in which only two alternatives receive positive probability, are of course a theoretical idealisation. In practice, an empirical distribution roughly conforming to the theory will be *bimodal*, expressing broadly polarised preferences in society or within an individual. Evidence of bimodal distributions in individual and collective choice is found in disparate contexts.⁵

We now discuss the key analytical facts in the 'base case' of the theory. Suppose that choices are compatible with dRUM, but let us exclude for simplicity the special case in which the two states have exactly the same probabilities, which presents some peculiarities. Suppose you observe that *a* is chosen with the same probability α when the menu is *A* and when the menu is *B*. Then, while you do not know in which state the choices were made, you do know that *a* was chosen (had the highest utility) in the *same* state, say *s*, in *A* and *B*. So, when the menu is $A \cup B$, *a* continues to have the highest utility in state *s*, while in the other state some alternative in *A* or *B* has a higher utility than *a*. Therefore *a* must be chosen with probability α also in $A \cup B$. This *Constant Expansion* property is a first necessary property of dRUM.

Suppose instead that a is chosen with different probabilities from A and B, and that these probabilities are less than one. Then it must be the case that a has the highest utility in different states in A and B (this would not be the case if a was chosen with probability one from A or B, for then a would have the highest utility in both states in one or both of the two menus). As a consequence, in each state there is some alternative

⁵E.g. Frolich, Oppenheimer and Moore [15] and Dufwenberg and Muren [9] (choices in a dictator games concentrated on giving nothing or 50/50), Sura, Shmuelib, Bosec and Dubeyc [24] (bimodal distributions in ratings, such as Amazon), Plerou, Gopikrishnan, and Stanley [20] (phase transition to bimodal demand -"bulls and bears"- in financial markets), Engelmann and Normann [12] (bimodality on maximum and minimum effort levels in minimum effort games), McClelland, Schulze and Coursey [19] (bimodal beliefs for unlikely events and willingness to insure).

that has a higher utility than *a* in $A \cup B$, and therefore *a* must be chosen with probability zero from $A \cup B$. This *Negative Expansion property* is a second necessary property of dRUM.

Theorem 1 says that, in the presence of the standard axiom of Regularity (adding new alternatives to a menu cannot increase the choice probability of an existing alternative), Constant and Negative Expansion are in fact *all* the behavioural implications of dRUM: the three axioms completely characterise dRUM when state probabilities are asymmetric.

The case which also allows for symmetric state probabilities is more complex and requires a separate treatment. The Constant Expansion axiom no longer holds, since now it could happen that *a* is chosen in different states from *A* and from *B*. However, a weakening that maintains the same flavour continues to apply: the proviso that *a* is chosen with positive probability from $A \cup B$ must be added. Furthermore, a *contraction consistency* property, which is implied by the other axioms in the generic case, must be assumed explicitly here. It says that the *impact* of an alternative *a* on another alternative *b* (removing *a* affects the probability of *b*) is preserved when moving from larger to smaller menus. Theorem 2 shows that Regularity, the modified expansion axioms, and contraction consistency characterise general dRUMs.

The identification properties of the preferences are very different in the two cases: there is uniqueness only in the asymmetric one, and otherwise uniqueness holds only up to a certain class of ordinal transformation of the rankings.

While the symmetric state case is non generic, it is not only of technical interest, because it serves as a basis to analyse the extension of the model that we describe below.

dRUM is a restriction of the classical RUM. In the final part of the paper we explore a new type of random utility maximisation that explains behaviours outside the RUM family, by allowing state probabilities to depend on the menu. This is an extremely natural feature in many contexts: for example, it is so in the dual self interpretations (Chatterjee and Krishna [5]). This extended model is also interesting because it can -unlike dRUM- represent as random utility maximisation some prominent violations of Regularity, such as the attraction, similarity and compromise effects. At first sight, allowing total freedom in the way probabilities can depend on menus seems to spell a total lack of structure, but we show in Theorem 3 that in fact the menu-dependent probability model, too, can be neatly characterised by consistency properties, which mirror those characterising the fixed-probability model (while obviously being weaker). These properties are 'modal' in the sense that they make assertions on the relationship between the *certainty, impossibility and possibility*, rather than the numerical probability, of selection across menus.

The behaviour corresponding to a general RUM is not very well-understood: RUM is characterised by a fairly complex set of conditions whose behavioural interpretation is not transparent.⁶ For this reason other scholars have examined interesting special cases where additional structure is added, obtaining transparent behavioural characterisations of RUM. We recall in particular Gul and Pesendorfer [13], who assume alternatives to be lotteries and preferences to be von Neumann Morgenstern; and Apesteguia and Ballester [1], who examine, in an abstract context, a family of utility functions that satisfies a single-crossing condition. The two state case of RUM of this paper is another restriction that is useful in this sense.

2 Preliminaries

Let *X* be a set of $n \ge 2$ alternatives. The nonempty subsets of *X* are called *menus*. Let \mathcal{D} be the set of all nonempty menus.

A *stochastic choice rule* is a map $p : \mathcal{D} \to [0,1]$ such that: $\sum_{a \in A} p(a, A) = 1$ for all $A \in \mathcal{D}$; p(a, A) = 0 for all $a \notin A$; and $p(a, A) \in [0,1]$ for all $a \in A$, for all $A \in \mathcal{D}$. The value p(a, A) may be interpreted as expressing the probability with which an individual agent, on whose behaviour we focus, chooses *a* from the menu *A*, or as

⁶The RUM is classically characterised by the Block-Marschak-Falmagne conditions (Block and Marschak [3]; Falmagne [14]; Barbera and Pattanaik [2]). These conditions constitute a very useful *algorithm*, but they are complex and can be interpreted only in terms of the representation (i.e. as indicating certain features of the probabilities of rankings). Thus they offer little insight on the behavioural restrictions imposed by the model which is the main theme of the present paper.

the fraction of a population choosing *a* from *A*.

In Random Utility Maximisation (RUM) a stochastic choice rule is constructed by assuming that there is probabilistic state space and a state dependent utility which is maximised. But provided that the event that two alternatives have the same realised utility has probability zero, RUM can be conveniently described in terms of random rankings (see e.g. Block and Marschak [3]), which we henceforth do. A *ranking* of *X* is a bijection $r : \{1, ..., n\} \rightarrow X$. Let \mathcal{R} be the set of all rankings. We interpret a $r \in \mathcal{R}$ as describing alternatives in decreasing order of preference: r(1) is the most preferred alternative, r(2) the second most preferred, and so on.

Let $\mathcal{R}(a, A)$ denote the set of rankings for which *a* is the top alternative in *A*, that is,

$$\mathcal{R}\left(a,A\right) = \left\{r \in \mathcal{R}: r^{-1}\left(b\right) < r^{-1}\left(a\right) \Rightarrow b \notin A\right\}.$$

Let μ be a probability distribution on \mathcal{R} . A *RUM* is a Stochastic choice rule p such that

$$p(a, A) = \sum_{r \in \mathcal{R}(a, A)} \mu(r)$$

A *dual RUM* (*dRUM*) is a RUM that uses only two rankings, that is one for which the following condition holds: $\mu(r) \mu(r') > 0$, $r \neq r' \Rightarrow \mu(r'') = 0$ for all rankings $r'' \neq r, r'$. A dRUM *p* is thus identified by a triple (r_1, r_2, α) of two rankings and a number $\alpha \in [0, 1]$. We say that (r_1, r_2, α) generates *p*.

To ease the notation for ranking relations, from now on given rankings r_i , i = 1, 2, we shall write $a \succ_i b$ instead of $r_i^{-1}(a) < r_i^{-1}(b)$.

A dRUM is a RUM and thus satisfies, for all menus *A*, *B*:

Regularity: If $A \subset B$ then $p(a, A) \ge p(a, B)$.

The two key additional properties are the following, for all menus *A*, *B*:

Constant Expansion: If $p(a, A) = p(a, B) = \alpha$ then $p(a, A \cup B) = \alpha$.

Negative Expansion: If p(a, A) < p(a, B) < 1 then $p(a, A \cup B) = 0$.

Constant Expansion extends to stochastic choice the classical expansion property of rational deterministic choice: if an alternative is chosen (resp., rejected) from two menus, then it is chosen (resp., rejected) from the union of the menus. In the stochastic setting the axiom remains a menu-independence condition. For example, in the population interpretation with instinctive and contemplative agents, the same frequencies of choice for *a* in two menus reveal that *a* got support either from the instinctive group in both menus or from the contemplative group in both menus. Therefore, if support is menu-independent it will persist when the menus are merged.

Negative Expansion is peculiar to stochastic choice since the antecedent of the axiom is impossible in deterministic choice, and is a 'dual' form of menu-independence condition. For example, in the population interpretation, different frequencies of choice for *a* in two menus reveal (given that they are less than unity) that in one menu *a* was not supported by one group, and in the other menu *a* was not supported by the other group. Therefore, if support is menu-independent no group will support *a* when the menus are merged.

Remark 1 We note in passing that Regularity and Negative Expansion imply almost immediately a well-known property satisfied by any RUM, Weak Stochastic Transitivity, defined as follows: if $p(a, \{a, b\}) \ge \frac{1}{2}$ and $p(b, \{b, c\}) \ge \frac{1}{2}$ then $p(a, \{a, c\}) \ge \frac{1}{2}$. To see this, suppose that the premise of Weak Stochastic Transitivity holds but that $p(a, \{a, c\}) < \frac{1}{2}$. Then (noting that it must be $p(c, \{b, c\}) \le \frac{1}{2}$ and $p(c, \{a, c\}) > \frac{1}{2}$) Negative Expansion implies $p(a, \{a, b, c\}) = 0 = p(c, \{a, b, c\})$, and hence $p(b, \{a, b, c\}) = 1$, contradicting Regularity and $p(a, \{a, b\}) > 0$.

Since dRUMs satisfy Regularity and Negative Expansion, they satisfy Weak Stochastic Transitivity, whereas any RUM with more than two states may violate it.

3 Asymmetric dRUMs

The case in which all non-degenerate choice probabilities are exactly equal to $\frac{1}{2}$ requires a separate method of proof and a different axiomatisation from all other cases $\alpha \in$ (0,1). Thus we begin our analysis by excluding for the moment this special case. The characterisation of the general case will be based on the results in this section.

A stochastic choice rule is *asymmetric* if $p(a, A) = \alpha \neq \frac{1}{2}$ for some *A* and $a \in A$. It is

symmetric if $p(a, A) \in (0, 1)$ for some A and $a \in A$ and $p(a, A) = \frac{1}{2}$ for all A and $a \in A$ for which this is the case.

Our main result in this section is:

Theorem 1 An asymmetric stochastic choice rule *p* is a dRUM if and only if it satisfies Regularity, Constant Expansion and Negative Expansion.

In order to prove the theorem we establish two key preliminary results. The first lemma shows that in any menu there are at most two alternatives receiving positive choice probability

Lemma 1 Let p be a stochastic choice rule that satisfies Regularity, Constant Expansion and Negative Expansion. Then for any menu A, if $p(a, A) \in (0, 1)$ for some $a \in A$ there exists $b \in A$ for which p(b, A) = 1 - p(a, A).

Proof: Suppose by contradiction that for some menu *A* there exist $b_1, ..., b_n \in A$ such that n > 2, $p(b_i, A) > 0$ for all *i* and $\sum_{i=1}^n p(b_i) = 1$. Since n > 2, we have $p(b_i, A) + p(b_j, A) < 1$ for all distinct *i*, *j*, and therefore (up to relabeling) $p(b_i, \{b_i, b_j\}) > p(b_i, A)$. Fixing *i*, *j*, if it were $p(b_i, \{b_i, b_j\}) = p(b_i, \{b_i, b_k\})$ for all $k \neq i, j$, then by Constant Expansion $p(b_i, \{b_i, b_j\}) = p(b_i, \{b_1, ..., b_n\}) = p(b_i, A)$ (the last equality holding by Regularity), a contradiction. Therefore $p(b_i, \{b_i, b_j\}) \neq p(b_i, \{b_i, b_k\})$ for some *k*. But then by Negative Expansion $p(b_i, \{b_i, b_j, b_k\}) = 0$, contradicting Regularity and $p(b_i, A) > 0$.

The next lemma shows that the two alternatives receiving positive probability do not change across menus.

Lemma 2 Let *p* be a stochastic choice rule that satisfies Regularity, Constant Expansion and Negative Expansion. Then there exists $\alpha \in [0, 1]$ such that, for any menu *A* and all $a \in A$, $p(a, A) \in \{0, \alpha, 1 - \alpha, 1\}$.

Proof: If for all *A* we have p(a, A) = 1 for some *a* there is nothing to prove. Then take *A* for which $p(a, A) = \alpha \in (0, 1)$ for some *a*. By Lemma 1 there exists exactly one alternative $b \in A$ for which $p(b, A) = 1 - \alpha$. Note that by Regularity $p(a, \{a, c\}) \ge \alpha$ for all $c \in A \setminus \{a\}$ and $p(b, \{b, c\}) \ge 1 - \alpha$ for all $c \in A \setminus \{b\}$.

Now suppose by contradiction that there exists a menu *B* with $c \in B$ for which $p(c, B) = \beta \notin \{0, \alpha, 1 - \alpha, 1\}$. By

Lemma 1 there exists exactly one alternative $d \in B$ for which $p(d, B) = 1 - \beta$.

Consider the menu $A \cup B$. By Regularity $p(e, A \cup B) = 0$ for all $e \in (A \cup B) \setminus \{a, b, c, d\}$. Moreover, by Lemma 1 $p(e, A \cup B) = 0$ for some $e \in \{a, b, c, d\}$, and w.l.o.g., let $p(a, A \cup B) = 0$.

If $p(b, A \cup B) > 0$ then by Regularity and Negative Expansion $p(b, A \cup B) = 1 - \alpha$ (if it were $p(b, A \cup B) < 1 - \alpha$, then we would have $p(b, A \cup B) \neq p(b, A)$ with $p(b, A \cup B)$, p(b, A) < 1 so that Negative Expansion would imply the contradiction $p(b, A \cup B) = 0$). It follows from Lemma 1 that either $p(c, A \cup B) = \alpha$ or $p(d, A \cup B) = \alpha$. In the former case, since $\alpha > 0$ a reasoning analogous to the one followed so far implies that $p(c, A \cup B) = p(c, B) = \beta$, contradicting $\alpha \neq \beta$; and in the latter case we obtain the contradiction $p(d, A \cup B) = p(d, B) = 1 - \beta$.

The proof is concluded by arriving at mirror contradictions starting from $p(c, A \cup B) > 0$ or $p(b, A \cup B) > 0$.

Proof of Theorem 1: Necessity, described informally in the introduction, is easily checked. For sufficiency, we proceed by induction on the cardinality of *X*. As a preliminary observation, note that if *p* satisfies the axioms on *X*, then its restriction to $X \setminus \{a\}$ also satisfies the axioms on $X \setminus \{a\}$.⁷

Moving to the inductive argument, if |X| = 2 the result is easily shown. For |X| = n > 2, we consider two cases.

<u>CASE A</u>: p(a, X) = 1 for some $a \in X$.

Denote p' the restriction of p to $X \setminus \{a\}$. By the preliminary observation and the inductive hypothesis there exist two rankings r'_1 and r'_2 on $X \setminus \{a\}$ and $\alpha \in [0,1]$ such that (r'_1, r'_2, α) generates p' on $X \setminus \{a\}$. Extend r'_1 and r'_2 to r_1 and r_2 on X by letting $r_1(1) = r_2(1) = a$ and letting r_i agree with r'_i , i = 1, 2, for all the other alternatives,

⁷If $A \subseteq B \subseteq X \setminus \{a\}$ then by Regularity (defined on the entire *X*) $p(b, A) \ge p(b, B)$. If $A, B \subseteq X \setminus \{a\}$ and $p(b, A) = p(b, B) = \alpha$ then $A \cup B \subseteq X \setminus \{a\}$ and by Constant Expansion (on *X*) $p(b, A \cup B) = \alpha$. If $A, B \subseteq X \setminus \{a\}$, $p(b, A) \neq p(b, B)$ and p(b, A), p(b, B) < 1 then $A \cup B \subseteq X \setminus \{a\}$ and by Negative Expansion (on *X*) $p(b, A \cup B) = 0$.

that is $c \succ_i d \Leftrightarrow c \succ'_i d$ for all $c, d \neq a$ (where for all x, y we write $x \succ'_i y$ to denote $(r'_i)^{-1}(x) < (r'_i)^{-1}(y)$). Let \tilde{p} be the dRUM generated by (r_1, r_2, α) . Take any menu A such that $a \in A$. Then $\tilde{p}(a, A) = 1$. Since by Regularity p(a, A) = 1, we have $p = \tilde{p}$ as desired.

<u>CASE B</u>: $p(a, X) = \alpha \neq \frac{1}{2}$ and $p(b, X) = 1 - \alpha$ for distinct $a, b \in X$ and $\alpha \in (0, 1)$.

Note that by Lemma 1 and 2 this is the only remaining possibility (in particular, if it were $p(a, X) = \frac{1}{2}$ then $p(c, A) \in \{0, \frac{1}{2}, 1\}$ for all menus A and $c \in A$, in violation of p being asymmetric). Assume w.l.o.g. that $\alpha > \frac{1}{2}$. As in Case A, by the inductive step there are two rankings r'_1 and r'_2 on $X \setminus \{a\}$ and $\beta \in [0, 1]$ such that (r'_1, r'_2, β) generates the restriction p' of p to $X \setminus \{a\}$. If for all $A \subseteq X \setminus \{a\}$ there is some $c_A \in A$ such that $p(c_A, A) = 1$, then the two rankings must agree on $X \setminus \{a\}$. In this case, for any $\gamma \in [0, 1]$, (r'_1, r'_2, γ) also generates p' on $X \setminus \{a\}$ for which $p(c, A) \neq 1$ for all $c \in A$, we know from Lemmas 1 and 2 that then there exist exactly two alternatives $c, d \in A$ such that $p(d, A) = \alpha$ and $p(c, A) = 1 - \alpha$. Therefore in this case it must be $\beta = \alpha$ (otherwise (r'_1, r'_2, β) would not generate the choice probabilities $p(d, A) = \alpha$ and $p(c, A) = 1 - \alpha$). Note that we must have $r'_2(1) = b$ since by Regularity and Lemma 2 either $p(b, X \setminus \{a\}) = 1$ (in which case also $r'_1(1) = b$) or $p(b, X \setminus \{a\}) = 1 - \alpha$.

Now extend r'_1 to r_1 on X by setting $r_1(1) = a$ and letting the rest of r_1 agree with r'_1 on $X \setminus \{a\}$; and extend r'_2 to r_2 on X by setting

 $a \succ _{2}c$ for all c such that $p(a, \{a, c\}) = 1$ $c \succ _{2}a$ for all c such that $p(a, \{a, c\}) = \alpha$.

and letting the rest of r_2 agree with r'_2 on $X \setminus \{a\}$. Note that by Regularity and Lemma 2 the two possibilities in the display are exhaustive (recall the assumption $\alpha > \frac{1}{2}$).

We need to check that this extension is consistent, that is that r_2 does not rank a above some c but below some d that is itself below c. Suppose to the contrary that there existed c and d for which $c \succ_2 d$ but $a \succ_2 c$ and $d \succ_2 a$. These three inequalities imply, respectively, that $p(c, \{c, d\}) \ge 1 - \alpha$ (since r'_1 and r'_2 generate p' on $X \setminus \{a\}$), $p(a, \{a, c\}) = 1$ and $p(a, \{a, d\}) = \alpha$ (by the construction of r_2).

If $c \succ'_1 d$ then $p(d, \{c, d\}) = 0$, hence $p(d, \{a, c, d\}) = 0$ by Regularity. Since $p(c, \{a, c\}) = 0$ it is also $p(c, \{a, c, d\}) = 0$ by Regularity, and thus $p(a, \{a, c, d\}) = 1$ contradicting Regularity and $p(a, \{a, d\}) = \alpha < 1$.

If instead $c \succ'_1 d$ then $p(d, \{c, d\}) = \alpha$. Since $p(d, \{a, d\}) = 1 - \alpha$, by Negative Expansion $p(d, \{a, c, d\}) = 0$ and we can argue as above. This shows that the extension r_2 is consistent.

Denote \tilde{p} the choice probabilities associated with the dRUM (r_1, r_2, α) . We now show that $p = \tilde{p}$.

Take any menu *A* for which $a \in A$. By Regularity (recalling the assumption $\alpha > \frac{1}{2}$) either (1) p(a, A) = 1 or (2) $p(a, A) = \alpha$ and $p(c, A) = 1 - \alpha$ for some $c \in A \setminus \{a\}$. In case (1), $p(a, \{a, d\}) = 1$ for all $d \in A \setminus \{a\}$ by Regularity. Thus $a \succ_1 d$ and $a \succ_2 d$ for all $d \in A \setminus \{a\}$, and therefore $\tilde{p}(a, A) = 1 = p(a, A)$ as desired.

Consider then case (2). Since $p(c, A) = 1 - \alpha$ we have $p(c, \{a, c\}) \ge 1 - \alpha$ by Regularity. Then by Regularity again we have $p(a, \{a, c\}) = \alpha$ and therefore $c \succ_2 a$, so that $\tilde{p}(a, A) = \alpha$ as desired. It remains to be checked that $\tilde{p}(c, A) = 1 - \alpha$. This means showing that $c \succ_2 d$ also for all $d \in A \setminus \{a, c\}$, which is equivalent to showing that $c \succ'_2 d$ for all $d \in A \setminus \{a\}$. If $p(c, A \setminus \{a\}) = 1$, this is certainly the case since (r'_1, r'_2, α) generates the restriction of p to $A \setminus \{a\}$. If instead $p(c, A \setminus \{a\}) = 1 - \alpha$ (the only other possibility by Regularity and Lemmas 1 and 2), noting that $\alpha \neq 1 - \alpha$, again this must be the case given that (r'_1, r'_2, α) generates the restriction of p to $A \setminus \{a\}$ (observe in passing that this conclusion might not hold in the case $\alpha = \frac{1}{2}$), and this concludes that proof.

We derived from primitive properties the fact that in an asymmetric dRUM there can be only two fixed non-degenerate choice probabilities $\alpha \neq \frac{1}{2}$ and $1 - \alpha$ in any menu. Since this fact, if it occurs, should be easily detectable from the data, one may wonder if a characterisation can be obtained by simply postulating it directly as an axiom together with Regularity:

Range: There exists $\alpha \in (0,1)$, $\alpha \neq \frac{1}{2}$, such that for all $A \subseteq X$ and $a \in A$, $p(a, A) \in \{0, \alpha, 1 - \alpha, 1\}$.

But it turns out that the Range and Regularity assumptions are not sufficient to ensure that the data are generated by a dRUM: they must still be disciplined by further across-menu consistency properties. Here is a stochastic choice rule p that satisfies Range and Regularity (and even Constant Expansion) but is not a dRUM:

	$\{a,b,c,d\}$	$\{a,b,c\}$	$\{a,b,d\}$	$\{a, c, d\}$	$\{b,c,d\}$	$\{a,b\}$	$\{a,c\}$	$\{a,d\}$	$\{b,c\}$	$\{b,d\}$	$\{c,d\}$
а	0	$\frac{2}{3}$	$\frac{1}{3}$	0	_	$\frac{2}{3}$	$\frac{2}{3}$	$\frac{1}{3}$	_	_	_
b	0	0	0	_	0	$\frac{1}{2}$	_	_	$\frac{2}{3}$	$\frac{1}{3}$	_
С	$\frac{1}{3}$	$\frac{1}{3}$	-	$\frac{1}{3}$	$\frac{1}{3}$	_	$\frac{1}{3}$	_	$\frac{1}{3}$	_	$\frac{1}{3}$
d	$\frac{2}{3}$	_	$\frac{2}{3}$	$\frac{2}{3}$	$\frac{2}{3}$	_	—	$\frac{2}{3}$	—	$\frac{2}{3}$	$\frac{2}{3}$

Table 1: A Stochastic Choice Function that fails Negative Expansion

That *p* is not a dRUM follows by Theorem 1 and the fact that *p* fails Negative Expansion (e.g. $p(a, \{a, b\}) = \frac{2}{3} \neq \frac{1}{3} = p(a, \{a, d\})$ while $p(a, \{a, b, d\}) > 0$).

4 General dRUMs

The easy inductive method we used in the proof of Theorem 1 breaks down for general dRUMs. We have to engage more directly with reconstructing the rankings from choice probabilities. The difficulty here is that, while it is clear that an alternative *a* being ranked above another alternative *b* in *some* ranking is revealed by the fact that removing *a* from a menu increases the choice probability of *b*, it is not self-evident in *which* ranking *a* is above *b* (whereas in the asymmetric case the two rankings are distinguished by the distinct probabilities α and $1 - \alpha$).

If the probabilities are allowed to take on the value $\frac{1}{2}$, then Constant Expansion is no longer a necessary property. In fact, if $p(a, A) = p(a, B) = \frac{1}{2}$ it could happen that ais top in A according to the ranking r_1 (but not according to r_2) and top in B according to r_2 (but not according to r_1). Then in $A \cup B$ there will be alternatives that are above ain both rankings, so that $p(a, A \cup B) = 0$.

The following weakening of Constant Expansion is an expansion property that circumvents the above situation and is necessary: Weak Constant Expansion: If $p(a, A) = p(a, B) = \alpha$ and $p(a, A \cup B) > 0$ then $p(a, A \cup B) = \alpha$.

It is easy to check that Weak Constant Expansion is necessary for a dRUM. In addition, general dRUMs are also easily checked to satisfy Negative Expansion. The following example indicates however that Regularity, Weak Constant Expansion and Negative Expansion are not tight enough to characterise general dRUMs:

	$\{a,b,c,d\}$	$\{a,b,c\}$	$\{a, b, d\}$	$\{a, c, d\}$	$\{b, c, d\}$	$\{a,b\}$	$\{a,c\}$	$\{a,d\}$	$\{b,c\}$	$\{b,d\}$	$\{c,d\}$
а	0	$\frac{1}{2}$	$\frac{1}{2}$	0	_	$\frac{1}{2}$	$\frac{1}{2}$	$\frac{1}{2}$	—	_	_
b	0	0	0	—	0	$\frac{1}{2}$	_	_	$\frac{1}{2}$	$\frac{1}{2}$	_
С	$\frac{1}{2}$	$\frac{1}{2}$	_	$\frac{1}{2}$	$\frac{1}{2}$	_	$\frac{1}{2}$	_	$\frac{1}{2}$	_	$\frac{1}{2}$
d	$\frac{1}{2}$	_	$\frac{1}{2}$	$\frac{1}{2}$	$\frac{1}{2}$	—	—	$\frac{1}{2}$	—	$\frac{1}{2}$	$\frac{1}{2}$

Table 2: A Stochastic Choice Function that fails Contraction Consistency

To see why the example above is not a dRUM, suppose it were. Then $p(a, \{a, b, c, d\}) = 0$ and $p(a, \{a, b, c\}) = \frac{1}{2}$ imply that *d* is an immediate predecessor of *a* in one of the two rankings on $\{a, b, c, d\}$. This must remain the case when the orderings are restricted to $\{a, b, d\}$, yet $p(a, \{a, b, d\}) = \frac{1}{2} = p(a, \{a, b\})$ imply that *d* is not an immediate predecessor of *a* on $\{a, b, d\}$ in either ranking.

The final new property used in the characterisation below is interesting and uses the concept of *impact* (Manzini and Mariotti [18]): *b* impacts *a* in *A* whenever $p(a, A) > p(a, A \cup \{b\})$.

Contraction Consistency: If *b* impacts *a* in *A* and $a \in B \subset A$ then *b* impacts *a* in *B*. Formally: For all $a \in B \subseteq A$ and $b \in X$: $p(a, A) > p(a, A \cup \{b\}) \Rightarrow p(a, B) > p(a, B \cup \{b\})$.

Contraction Consistency simply says that impact is inherited from menus to submenus. This property is implied by the other properties in the statement of Theorem 1 when the dRUM is asymmetric, but it needs to be assumed explicitly otherwise. Let's check the implication first: **Proposition 1** Let *p* be an asymmetric dRUM that satisfies Constant Expansion, Regularity and Negative Expansion. Then it satisfies Contraction Consistency.

Proof. Let *p* satisfy Constant Expansion, Regularity and Negative Expansion, let $a \in$ $B \subseteq A$ and $b \in X$ and suppose by contradiction that $p(a, A) > p(a, A \cup \{b\})$ but $p(a, B) \leq p(a, B \cup \{b\})$. Since by Regularity $p(a, B) \geq p(a, B \cup \{b\})$, it can only be $p(a, B) = p(a, B \cup \{b\})$. Also, by Regularity it must be $p(a, B) \ge p(a, A)$. However it cannot be p(a, B) = p(a, A), for in that case $p(a, B \cup \{b\}) = p(a, A)$ and Constant Expansion would imply the contradiction $p(a, A \cup \{b\}) = p(a, (B \cup \{b\}) \cup A) =$ $p(a, A) > p(a, A \cup \{b\})$. So it must be p(a, B) > p(a, A), and the only possible configuration of choice probabilities is $p(a, B) = p(a, B \cup \{b\}) > p(a, A) > p(a, A \cup \{b\})$. Now w.l.o.g. let $\alpha > 1 - \alpha$. It cannot be that $p(a, B \cup \{b\}) = \alpha$ and p(a, A) = $1 - \alpha$, since then Negative Expansion and $p(a, B) \neq p(a, A)$ imply the contradiction $p(a, A \cup B) = p(a, A) = 0 > p(a, A \cup \{b\})$. Similarly, it cannot be p(a, A) = 0 $\alpha > p(a, A \cup \{b\}) = 1 - \alpha$, for again Negative Expansion implies the contradiction $p(a, A \cup \{b\}) = p(a, A \cup (A \cup \{b\})) = 0 \neq 1 - \alpha = p(a, A \cup \{b\})$. The only remaining possibility is $p(a, B \cup \{b\}) = p(a, B) = 1$, $p(a, A) \in \{\alpha, 1 - \alpha\}$ and $p(a, A \cup \{b\}) = 0$ 0. By Regularity, $p(a, B \cup \{b\}) = 1$ implies $p(a, \{a, b\}) = 1$. Moreover if p(a, A) > 0 $p(a, A \cup \{b\})$ it must be $p(b, A \cup \{b\}) > 0$ (if not, then there exist $x, y \in A$ such that $p(x, A \cup \{b\}), p(y, A \cup \{b\}) > 0$ and $p(x, A \cup \{b\}) + p(y, A \cup \{b\}) = 1$, where possibly x = y while $a \neq x, y$ since $p(a, A \cup \{b\}) = 0$. But then by Regularity $p(x, A) \ge p(x, A \cup \{b\})$ and $p(y, A) \ge p(y, A \cup \{b\})$, implying the contradiction p(a, A) = 0. But if $p(b, A \cup \{b\}) > 0$ the contradiction $p(b, \{a, b\}) > 0$ follows. We conclude that Contraction Consistency holds.

Theorem 2 *A stochastic choice rule is a dRUM if and only if it satisfies Regularity, Weak Constant Expansion, Negative Expansion and Contraction Consistency.*

Proof. Necessity is straightforward, so we show sufficiency.

Preliminaries and outline Let p satisfy Regularity, Weak Constant Expansion, Negative Expansion and Contraction Consistency. The proof consists of three blocks: first we define an algorithm to construct two rankings r_1 and r_2 explicitly (block 1), then we

show that the algorithm is well defined (block 2), and finally we show that r_1 and r_2 so constructed retrieve p (block 3).

Observe that Lemma 1 and Lemma 2 continue to hold when Constant Expansion is replaced by Weak Constant Expansion.⁸ In the proof of Theorem 1 we did not use Constant Expansion explicitly. So in view of Theorem 1 we only need to consider the case $p(a, A) \in \{0, \frac{1}{2}, 1\}$ for all A and $a \in A$, and at most two alternatives are chosen with positive probability in any menu A.

1. The algorithm to construct the two rankings.

Let $a, b \in X$ be such that p(a, X), p(b, X) > 0. Enumerate the elements of the rankings r_1 and r_2 to be constructed by $r_1(i) = x_i$ and $r_2(i) = y_i$ for i = 1, ...n. Set $x_1 = a$ and $y_1 = b$ (were possibly $x_1 = a = b = y_1$).

The rest of the rankings r_1 and r_2 for $i \ge 2$ are defined recursively. Let

$$L_i^1 = X \setminus \bigcup_{j=1}^{i-1} x_j$$

Next, define the set S_i of alternatives that are impacted by x_{i-1} in $L_i^1 \cup \{x_{i-1}\}$:

$$S_i = \left\{ s \in L_i^1 : p\left(s, L_i^1\right) > p\left(s, L_i^1 \cup \{x_{i-1}\}\right) \right\}$$

where observe that by lemmas 1 and 2 we have $0 < |S_i| \le 2$.

If $S_i = \{c\}$ for some $c \in X$, then let $x_i = c$.

If $|S_i| = 2$, then let $S_i = \{c, d\}$ with $c \neq d$ and consider two cases:

(1.*i*) $p(c, L_i^1) > p(c, L_i^1 \cup \{x_j\})$ (resp., $p(d, L_i^1) > p(d, L_i^1 \cup \{x_j\})$) for all j = 1, ...i - 1 and $p(d, L_i^1) = p(d, L_i^1 \cup \{x_j\})$ (resp., $p(c, L_i^1) = p(c, L_i^1 \cup \{x_j\})$) for some $j \in \{1, ..., i - 2\}$ (that is, one alternative is impacted by *all* predecessors x_j in L_i^1 while the other alternative is not).

In this case let $x_i = c$ (resp., $x_i = d$).

(1.*ii*) $p(c, L_i^1) > p(c, L_i^1 \cup \{x_j\})$ and $p(d, L_i^1) > p(d, L_i^1 \cup \{x_j\})$ for all j = 1, ..., i - 1(both alternatives are impacted by all predecessors x_j in L_i^1).

⁸To check, note that Constant Expansion is invoked only in the proof of Lemma 1, and only once, in a step where the additional clause of Weak Constant Expansion is satisfied by hypothesis.

In this case, let $x_i = c$.

Proceed in an analogous way for the construction of r_1 , starting from y_1 , that is for all i = 2, ...n define recursively

$$L_i^2 = X \setminus \bigcup_{j=1}^{i-1} y_j$$
$$T_i = \left\{ t \in L_i^2 : p\left(t, L_i^2\right) > p\left(t, L_i^2 \cup \{y_{i-1}\}\right) \right\}$$

and as before $0 < |S_i| \le 2$.

If $|T_i| = \{t\}$, then let $y_i = t$, while if $|T_i| = 2$, then letting $T_i = \{e, f\}$:

(2.*i*). $p(e, L_i^2) > p(e, L_i^2 \cup \{y_j\})$ (resp. $p(e, L_i^2) > p(e, L_i^2 \cup \{y_j\})$) for all j = 1, ...i - 1, and $p(f, L_i^2) = p(f, L_i^2 \cup \{y_j\})$ (resp. $p(e, L_i^2) = p(e, L_i^2 \cup \{y_{i-g}\})$) for some $j \in \{1, ..., i-2\}$.

In this case let $y_i = e$ (resp., f).

(2.*ii*).
$$p(e, L_i^2) > p(e, L_i^2 \cup \{y_j\})$$
 and $p(f, L_i^2) > p(f, L_i^2 \cup \{y_j\})$ for all $j = 1, ..., i - 1$.

In this case let $y_i = f$ (resp., $y_i = e$) whenever $eP(r_1) f$ (resp., $fP(r_1) e$) (i.e. we require consistency with the construction of the first order).

2. Showing that the algorithm is well-defined

We show that cases (1.*i*) and (1.*ii*) are exhaustive, which means showing that if $|S_i| = 2$ at least one alternative in S_i is impacted by *all* its predecessors according to r_1 . We proceed by induction on the index *i*. If i = 2 there is nothing to prove. Now consider the step i = k + 1. If $|S_{k+1}| = 1$ again there is nothing to prove, so let $|S_{k+1}| = 2$, with $S_{k+1} = \{c, d\}$. By construction $p(c, L_{k+1}^1) > p(c, L_{k+1}^1 \cup \{x_k\})$ and $p(d, L_{k+1}^1) > p(d, L_{k+1}^1 \cup \{x_k\})$, so that by Lemma 1 and Lemma 2 it must be

$$p(c, L_{k+1}^1) = \frac{1}{2} = p(d, L_{k+1}^1)$$
 (1)

By contradiction, suppose that there exist $u, v \in \{x_1, ..., x_{k-1}\}$ for which u does not impact c in L_{k+1}^1 and v does not impact d in L_{k+1}^1 , i.e.

$$p(c, L_{k+1}^1 \cup \{u\}) = p(c, L_{k+1}^1) = \frac{1}{2}$$
 (2)

and

$$p(d, L_{k+1}^1 \cup \{v\}) = p(d, L_{k+1}^1) = \frac{1}{2}$$
 (3)

(note that by construction we have $u, v \neq x_k$). We can rule out the case u = v. For suppose $u = v = x_j$ for some j. Then, recalling (1), $p(d, L_{k+1}^1 \cup \{x_j\}) = p(d, L_{k+1}^1) = \frac{1}{2} = p(c, L_{k+1}^1) = p(c, L_{k+1}^1 \cup \{x_j\})$ and thus $p(x_j, L_{k+1}^1 \cup \{x_j\}) = 0$, which contradicts Regularity and $p(x_j, L_j^1 \cup \{x_j\}) > 0$ with $L_{k+1}^1 \subset L_j^1$.

Since by construction and Regularity $p(v, L_{k+1}^1 \cup \{v\}) > 0$, Lemma 1, Lemma 2 and (3) imply

$$p\left(c, L_{k+1}^1 \cup \{v\}\right) = 0 \tag{4}$$

Similarly, $p(v, L_{k+1}^1 \cup \{v\}) > 0$ (by construction and Regularity), Lemma 1, Lemma 2 and (2) imply

$$p\left(d, L_{k+1}^1 \cup \{u\}\right) = 0 \tag{5}$$

(i.e. *u* impacts *d* in $L_{k+1}^1 \cup \{u\}$ and *v* impacts *c* in L_{k+1}^1). Now consider the menu $L_{k+1}^1 \cup \{u, v\}$. By (4), (5) and Regularity it must be:

$$p(c, L_{k+1}^{1} \cup \{u, v\}) = 0 = p(d, L_{k+1}^{1} \cup \{u, v\})$$
(6)

It cannot be that $p(u, L_{k+1}^1 \cup \{u, v\}) = 1$, for otherwise Regularity would imply $p(u, L_{k+1}^1 \cup \{u\}) = 1$, contradicting $p(c, L_{k+1}^1 \cup \{u\}) = p(c, L_{k+1}^1) = \frac{1}{2}$. Similarly, it cannot be that $p(v, \{u, v\} \cup L_{k+1}^1) = 1$. Finally, if either $p(v, \cup L_{k+1}^1 \cup \{u, v\}) = 0$ or $p(u, L_{k+1}^1 \cup \{u, v\}) = 0$, then in view of (6) it would have to be $p(w, L_{k+1}^1 \cup \{u, v\}) > 0$ for some $w \in L_{k+1}^1$. But this is impossible since $c \neq w \neq d$ by (6), and then by Regularity the contradiction $p(w, L_{k+1}^1) > 0$ would follow. Therefore it must be

$$p\left(u, L_{k+1}^{1} \cup \{u, v\}\right) = \frac{1}{2} = p\left(v, L_{k+1}^{1} \cup \{u, v\}\right)$$

It follows that both

$$p\left(u, L_{k+1}^1 \cup \{u\}\right) = p\left(u, L_{k+1}^1 \cup \{u, v\}\right)$$

$$\tag{7}$$

and

$$p\left(v, L_{k+1}^1 \cup \{v\}\right) = p\left(v, L_{k+1}^1 \cup \{u, v\}\right)$$

$$\tag{8}$$

i.e. neither does u impact v in $L_{k+1}^1 \cup \{v\}$, nor does v impact u in $L_{k+1}^1 \cup \{u\}$. Suppose w.l.o.g. that v is a predecessor of u, and let $u = x_j$. By the inductive hypothesis v impacts u in L_j^1 , so that by Contraction Consistency it also impacts u in $L_{k+1}^1 \cup \{u\} \subset L_j^1$, i.e.

$$p\left(u, L_{k+1}^1 \cup \{u\}\right) > p\left(u, L_{k+1}^1 \cup \{u, v\}\right)$$

a contradiction with (7). A symmetric argument applies if u is a predecessor of v using (8).

A straightforward adaptation of the argument above shows that cases (2.*i*) and (2.*ii*) are exhaustive.

3. Showing that the algorithm retrieves the observed choice.

Let $p_{\frac{1}{2}}$ be the dRUM generated by $(r_1, r_2, \frac{1}{2})$. For any alternative *x* denote L_x^i its (weak) lower contour set in ranking r_i , that is

$$L_x^i = \{x\} \cup \{s \in X : x \succ_i s\}$$

and note that by construction $p(x, L_x^i) > 0$. We examine the possible cases of failures of the algorithm in succession.

<u>3.1</u>: p(a, A) = 0 and $p_{\frac{1}{2}}(a, A) > 0$.

Then $a \succ_i a'$ for some *i*, for all $a' \in A \setminus \{a\}$, hence $A \subseteq L_a^i$, and thus by Regularity and $p(a, L_a^i) > 0$ we have p(a, A) > 0, a contradiction.

<u>3.2</u>: p(a, A) = 1 and $p_{\frac{1}{2}}(a, A) < 1$.

Then there exists $b \in B$ such that $b \succ_i a'$ for some i, for all $a' \in A \setminus \{b\}$, hence $A \subseteq L_b^i$, and thus by Regularity and $p(b, L_a^i) > 0$ we have p(b, A) > 0, a contradiction. <u>3.3:</u> $p(a, A) = \frac{1}{2}$ and $p_{\frac{1}{2}}(a, A) = 1$.

Let $\mathcal{I}\left(p_{\frac{1}{2}}\right)$ be the set of all pairs (a, A) satisfying the conditions of this case. Fix an $(a, A) \in \mathcal{I}\left(p_{\frac{1}{2}}\right)$ that is maximal in $\mathcal{I}\left(p_{\frac{1}{2}}\right)$ in the sense that $(b, B) \in \mathcal{I}\left(p_{\frac{1}{2}}\right) \Rightarrow a \succ_i b$ for some *i*.

Because $p_{\frac{1}{2}}(a, A) = 1$ we have $A \subseteq L_a^1 \cap L_a^2$. If $p(a, L_a^1 \cap L_a^2) = 1$ we have an immediate contradiction, since by Regularity p(a, A) = 1. So it must be $p(a, L_a^1 \cap L_a^2) = \frac{1}{2}$. We show that this also leads to a contradiction.

Suppose first that $L_a^1 \setminus L_a^2 \neq \emptyset$, and take $z \in L_a^1 \setminus L_a^2$ such that $a \succ_1 z$ and $z \succ_2 a$. By construction this implies $p(a, L_a^2) > p(a, \{z\} \cup L_a^2)$, so that by Contraction Consistency $p(a, L_a^1 \cap L_a^2) > p(a, \{z\} \cup (L_a^1 \cap L_a^2))$. Since by Regularity $p(a, \{z\} \cup (L_a^1 \cap L_a^2)) \ge p(a, L_a^2) = \frac{1}{2}$ (where recall $z \in L_a^1$, so that $z \cup L_a^1 = L_a^1$), we have $p(a, L_a^1 \cap L_a^2) = 1$, a contradiction. A symmetric argument applies if $L_a^2 \setminus L_a^1 \neq \emptyset$.

Finally, consider the case $L_a^1 = L_a^2 = L_a$ for some $L_a \subset X$. Then $X \setminus L_a^1 = X \setminus L_a^2 = U_a$ for some $U_a \subset X$ (where observe that $u \in U_a \Rightarrow u \succ_i a, i = 1, 2$). Since we assumed that $p(a, L_a^1 \cap L_a^2) = \frac{1}{2}$, there exists a $w \in (L_a^1 \cap L_a^2)$ for which $p(w, L_a^1 \cap L_a^2) = \frac{1}{2}$.

Let $x_1, ..., x_n$ be the predecessors of a in the ranking r_1 and let $y_1, ..., y_n$ be the predecessors of a in the ranking r_2 (where obviously $\{x_1, ..., x_n\} = U_a = \{y_1, ..., y_n\}$). So $a = x_{n+1} = y_{n+1}$ (note that since $L_a^1 = L_a^2$, a must have the same position in both rankings).

By the construction of the algorithm x_n and y_n impact a in L_a , that is $p(a, L_a) > p(a, L_a \cup \{x_n\})$ and $p(a, L_a) > p(a, L_a \cup \{y_n\})$.

We claim that there exists a $z \in U_a$ that does not impact w in L_a , that is $p(w, L_a) = p(w, L_a \cup \{z\})$. To see this, if $p(w, L_a) = p(w, L_a \cup \{x_n\})$ (resp. $p(w, L_a) = p(w, L_a \cup \{y_n\})$) simply set $z = x_n$ (resp. $z = y_n$). If instead $p(w, L_a) > p(w, L_a \cup \{x_n\})$ and $p(w, L_a) > p(w, L_a \cup \{y_n\})$ then, since $a = x_{n+1} = y_{n+1}$, $p(w, L_a) = p(w, L_a \cup \{z\})$ for some $z \in U_a$ by the construction of the algorithm (otherwise, given that x_n and y_n impact w in L_a , it should be $w = x_{n+1}$ or $w = y_{n+1}$). Fix such a z.

Regularity and the construction imply that $p(z, L_a \cup \{z\}) = \frac{1}{2}$. But since $z \in U_a$ we have $z \succ_i a'$, i = 1, 2, for all $a' \in L_a$, so that $p_{\frac{1}{2}}(z, L_a \cup \{z\}) = 1$ and therefore $z \in \mathcal{I}(p_{\frac{1}{2}})$. And since in particular $z \succ_i a$, i = 1, 2, the initial hypothesis that (a, A) is maximal in $\mathcal{I}(p_{\frac{1}{2}})$ is contradicted. <u>3.4</u>: $p(a, A) = \frac{1}{2}$ and $p_{\frac{1}{2}}(a, A) = 0$.

By construction there exist alternatives *b* and *c* (where possibly b = c) such that $b \succ_1 a'$ for all $a' \in A \setminus \{b\}$ and $c \succ_2 a'$ for all $a' \in A \setminus \{c\}$, so that $A \subseteq L_b^1$ and $A \subseteq L_c^2$. If $b \neq c$, then by Regularity the contradiction $p(b, A) = \frac{1}{2} = p(c, A)$ follows. Thus let b = c, so that $b \succ_i a'$ for all $a' \in A \setminus \{b\}$ for i = 1, 2, and by construction $p(b, L_b^1 \cap L_b^2) > 0$. If $p(b, L_b^1 \cap L_b^2) = 1$ then by Regularity we have a contradiction, since $A \subseteq L_b^1 \cap L_b^2$, so that $p(b, L_b^1 \cap L_b^2) = p(b, A) = \frac{1}{2}$, and otherwise we are back in case 3.3.

We note that a subtle relation for the case $p(a, A) \in \{0, \frac{1}{2}, 1\}$ exists with the "topand-the-top" (tat) model studied in Eliaz, Richter and Rubinstein [11] (ERR). A tat describes a *deterministic* choice procedure, in which the agent uses two ordering and picks all the alternatives in a menu that are top in at least one of the two orderings. While we are dealing with stochastic orderings, when these orderings have the same probabilities the resulting probabilities of choice are uninformative in distinguishing the orderings: the only information they give concerns their *support* (whereas in the asymmetric case the orderings can be told apart as the α -ordering and the $(1 - \alpha)$ -ordering). So, the stochastic choice function contains in fact the same information on the rankings as a tat. Note, however, that of course our axioms characterise simultaneously both the symmetric and the asymmetric case, and the latter has no correspondence with the deterministic procedure.⁹

5 Menu-dependent state probability and violations of Regularity

So far we have assumed that the probabilities with which the two orders are applied are fixed across menus. This is quite natural, for example, in the population interpretation. Yet in several other situations it makes sense to allow the probabilities of the two rankings in a dRUM to depend on the menu. We highlight some leading such cases.

• In the dual-self interpretation, if the duality of the self is due to temptation, then the presence of tempting alternatives may increase the probability that the shortterm self is in control, and possibly the more so the more numerous the tempting alternatives are.

⁹The key property used by ERR to characterise tat says that if an *x* is chosen from two menus *A* and *B* and also from $A \cap B$, and if the choice from $A \cap B$ consists of two elements, then *x* is chosen from $A \cup B$.

- In the household interpretation, husband and wife may have different 'spheres of control', so that menus containing certain items are more likely to be under the control of one of the two partners.
- If choices are subject to unobserved time pressure, the effect of time pressure is likely to be different in large and small menus. Since the latter are arguably easier to analyse, time pressure may be activated with a lower probability in simple menus.
- The logic of Luce and Raiffa's [17] well-known 'frog legs example' is that the presence of a specific item *a** (frog legs in the example) in a menu triggers the maximisation of a different preference order because *a** conveys information on the nature of the available alternatives, so that the choice is 'dual': in all menus containing *a** one preference order is maximised while in all those not containing *a** a different order is maximised. Making the probabilities in a dRUM menudependent further generalises this idea to a probabilistic context, to allow for example: *p* (steak, {steak, chicken, frog legs}) > *p* (steak, {steak, chicken, frog legs}) without making the extreme assumption that steak is chosen for sure when frog legs are also available.
- The 'similarity' and the 'attraction' effects are very prominent in the psychology and behavioural economics literature. According to these effects, the probability of choice of two alternatives that are chosen approximately with equal frequency in a binary context can be shifted in favour of one or the other through the addition of an appropriate 'decoy' alternative that is in a relation of similarity or dominance (in a space of characteristics) with one of the alternatives. Such effects can be also described as dual RUM by making the probabilities of the two rankings dependent on the presence and nature of the decoy alternative. Analogous considerations apply for the 'compromise effect', according to which the probability of choosing an alternative increases when it is located at an intermediate position in the space of characteristics compared to other more extreme alternatives.

A *Menu-dependent dRUM (mdRUM)* is a stochastic choice rule p for which the following is true: there exists a triple $(r_1, r_2, \tilde{\alpha})$ where r_1 and r_2 are rankings and $\tilde{\alpha}$: $2^X \setminus \emptyset \rightarrow (0, 1)$ is a function that for each menu A assigns a probability $\tilde{\alpha}(A)$ to r_1 (and $1 - \tilde{\alpha}(A)$ to r_2), such that, for all A, p(a, A) = p'(a, A) where p' is the dRUM generated by $(r_1, r_2, \tilde{\alpha}(A))$.

An mdRUM obviously loses, compared to a dRUM, all the properties that pertain to the specific magnitudes of choice probabilities, which depend on the menu in an unrestricted way. Hence in particular an mdRUM fails Regularity. But it is precisely this feature that gives the model its descriptive power.

It is easy to check, however, that discarding alternatives from a menu still preserves the *possibility* and the *certainty* of the event in which an alternative is chosen, when moving to sub-menus: in this sense it preserves the 'mode' of choice:¹⁰

Modal Regularity: Let $a \in B \subset A$. (i) If p(a, A) > 0 then p(a, B) > 0. (ii) If p(a, A) = 1 then p(a, B) = 1.

For the second key property we consider a notion of impact related to modality. Let's say that *b* modally impacts *a* in *A* if

$$p(a, A) > 0 \text{ and } p(a, A \cup \{b\}) = 0$$

or $p(a, A) = 1 \text{ and } p(a, A \cup \{b\}) \in (0, 1)$

That is, *b* modally impacts *a* if adding *b* transforms the choice of *a* from possible (including certain) to impossible or from certain to merely possible.

Modal Impact Consistency: Let $b \notin A$. If *b* does not modally impact *a* for all $a \in A$ then $p(b, A \cup \{b\}) = 0$.

Modal Impact Consistency is the modal version of a property that is implied automatically in the menu-independent model, and is in fact satisfied by *any* RUM (if *b* is chosen with positive probability in a menu, then it obviously impacts some alternative in any sub-menu).

¹⁰'Mode' and 'modal' are meant here and elsewhere in ther logical and not statistical meaning.

To prove that the properties are sufficient as well as necessary we exploit a relationship that exists between mdRUMs and dRUMs. The idea is to associate with each p another stochastic choice rule \hat{p} that is a symmetric dRUM if and only if p is an mdRUM. This trick allows us to rely on the previous results, as we then just need to 'translate' the properties that make \hat{p} an mdRUM into properties of p, which is an easier task.

As a preliminary, we prove a lemma analogous to Lemma 1. Say that a stochastic choice rule *p* is *binary* whenever it assigns positive probability to at most two elements in any menu (i.e. for any menu *A*, if $p(a, A) \in (0, 1)$ for some $a \in A$ then there exists a $b \in A$ for which p(b, A) = 1 - p(a, A)).

Lemma 3 Let *p* be a stochastic choice rule that satisfies Modal Regularity and Modal Impact Consistency. Then *p* is binary.

Proof: Suppose by contradiction that for some menu *A* there exist $b_1, ..., b_n \in A$ such that n > 2, $p(b_i, A) > 0$ for all i = 1, ..., n and $\sum_{i=1}^{n} p(b_i) = 1$. By Modal Regularity both $p(b_i, \{b_1, ..., b_{n-1}\}) > 0$ for all i = 1, ..., n - 1 and $p(b_i, \{b_1, ..., b_n\}) > 0$ for all i = 1, ..., n. Then b_n does not modally impact any alternative in $\{b_1, ..., b_{n-1}\}$, contradicting Modal Impact Consistency.

Second, we prove an auxiliary result that is of interest in itself, that is, a modal form of Contraction Consistency is implied by Modal Regularity and Modal Impact Consistency.

Lemma 4 Let *p* be a stochastic choice rule that satisfies Modal Regularity and Modal Impact Consistency. Suppose that b modally impacts a in A. Then b modally impacts a in any $B \subset A$.

Proof:

Step 1. For all *a*, *b* and *A*, if *b* modally impacts *a* in *A*, then it must be that $p(b, A \cup \{b\}) > 0$. To see this, suppose by contradiction that $p(b, A \cup \{b\}) = 0$. Then by Lemma 3 there are $c, d \neq b$ for which $p(c, A \cup \{b\}) > 0$ and $p(d, A \cup \{b\}) > 0$. If $p(c, A \cup \{b\}) = 1$ (i.e. c = d) then by Modal Regularity p(c, A) = 1, so that p(a, A) = 0 and we contradict the hypothesis that *b* modally impacts *a* in *A*. Otherwise $c \neq d$, $0 < p(c, A \cup \{b\}) < 1$ and $0 < p(d, A \cup \{b\}) < 1$, so that by Modal Regularity p(c, A) > 0

and p(d, A) > 0 (recall that $c, d \neq b$). If a = c or a = d, then since b modally impacts a in A it must be p(a, A) = 1, a contradiction with p(c, A) > 0. Therefore p(a, A) = 0, which is impossible if b modally impacts a in A.

<u>Step 2</u>. Let *a* ∈ *B* ⊂ *A*, and suppose that *b* modally impacts *a* in *A* in the sense that p(a, A) > 0 and $p(a, A \cup \{b\}) = 0$. We show the statement to be true in this case. By Modal Regularity p(a, B) > 0. By step 1 $p(b, A \cup \{b\}) > 0$, and then by Modal Regularity $p(b, B \cup \{b\}) > 0$, so that $p(a, B \cup \{b\}) < 1$. Therefore if p(a, B) = 1 we conclude immediately that *b* impacts *a* in *B* as desired, and likewise if p(a, B) < 1 and $p(a, B \cup \{b\}) = 0$. Suppose finally that p(a, B) < 1 and $p(a, B \cup \{b\}) = 0$. Suppose finally that p(a, B) < 1 and $p(a, B \cup \{b\}) = 0$. Suppose finally that p(a, B) < 1 and $p(a, B \cup \{b\}) > 0$. It cannot be $p(b, A \cup \{b\}) = 1$, for then Modal Regularity would imply the contradiction $p(b, B \cup \{b\}) = 1$, so there must exist some $c \in A$ such that $p(c, A \cup \{b\}) = 1 - p(b, A \cup \{b\}) > 0$ (where $c \neq a$ by the assumption $p(a, A \cup \{b\}) = 0$). But this cannot be the case, since then by Modal Regularity $p(c, B \cup \{b\}) > 0$, contradicting Lemma 3. Step 3. Let $a \in B \subset A$, and suppose that *b* modally impacts *a* in *A* in the sense that p(a, A) = 1 and $p(a, A \cup \{b\}) \in (0, 1)$. By Modal Regularity p(a, B) = 1. By step 1 $p(b, A \cup \{b\}) > 0$ and then by Modal Regularity $p(b, B \cup \{b\}) > 0$, so that $p(a, B \cup \{b\}) > 1$ and we conclude that *b* modally impacts *a* in *B* in this final case too.

Theorem 3 *A stochastic choice rule is an mdRUM if and only if it satisfies Modal Regularity and Modal Impact Consistency.*

Proof. Necessity is obvious. For sufficiency, given a stochastic choice rule *p*, let us say that \hat{p} is the *conjugate of p* if for all menus *A* and $a \in A$:

$$\hat{p}(a, A) = 1 \Leftrightarrow p(a, A) = 1$$
$$\hat{p}(a, A) = \frac{1}{2} \Leftrightarrow 1 > p(a, A) > 0$$

If *p* satisfies Modal Regularity and Modal Impact Consistency, by Lemma 3 it is binary and so the conjugate of a *p* always exists and is defined uniquely, but it is not necessarily a dRUM. When it is a dRUM, it is a symmetric dRUM. We now investigate when this is the case. Suppose that *p* is an mdRUM generated by $(r_1, r_2, \tilde{\alpha})$. Then, in view of the relationship p(a, A) = p'(a, A) where p' is the dRUM generated by $(r_1, r_2, \tilde{\alpha}(A))$, its conjugate \hat{p} is a dRUM generated by $(r_1, r_2, \frac{1}{2})$, and therefore by Theorem 2 \hat{p} satisfies Regularity, Weak Constant Expansion, Negative Expansion and Contraction Consistency. Conversely, if the conjugate \hat{p} of a p satisfies these axioms then \hat{p} is a dRUM generated by some $(r_1, r_2, \frac{1}{2})$ and therefore p is an mdRUM generated by $(r_1, r_2, \tilde{\alpha})$ where $\tilde{\alpha}$ is defined so that for any A such that p(a, A) > 0 and p(b, A) > 0 for distinct a and b, $\tilde{\alpha}(A) = p(a, A)$ and $1 - \tilde{\alpha}(A) = p(b, B)$ (or vice-versa). This reasoning shows that p is an mdRUM if and only if it is binary and its conjugate satisfies Regularity, Weak Constant Expansion, Negative Expansion and Contraction Consistency. Therefore we just need to verify that \hat{p} satisfies these axioms if p satisfies the axioms in the statement.

Step 1. \hat{p} satisfies Regularity. In fact, let $A \subset B$. If $\hat{p}(a, B) = 0$ then Regularity cannot be violated. If $0 < \hat{p}(a, B) < 1$ then it must be $\hat{p}(a, B) = \frac{1}{2}$, which implies p(a, B) > 0and hence by Modal Regularity part (i) p(a, A) > 0. Then $\hat{p}(a, A) \in \left\{\frac{1}{2}, 1\right\}$, satisfying Regularity. Finally if $\hat{p}(a, B) = 1$ then p(a, B) = 1, so that by part (ii) of Modal Regularity p(a, A) = 1, and then $\hat{p}(a, A) = 1$ as desired.

Step 2. \hat{p} satisfies Weak Constant Expansion. In fact, if $\hat{p}(a, A) = \hat{p}(a, B) = 1$ then it must be p(a, A) = 1 = p(a, B). Suppose $p(a, A \cup B) < 1$. Then there exists a $b \in A \cup B$ for which $p(b, A \cup B) > 0$ and either p(b, A) = 0 or p(b, A) = 0, contradicting Modal Regularity part (i). Therefore $p(a, A \cup B) = 1$ and then $\hat{p}(a, A \cup B) = 1$. If instead $\hat{p}(a, A) = \frac{1}{2} = \hat{p}(a, B)$ and $\hat{p}(a, A \cup B) > 0$ it must be p(a, A) < 1, p(a, B) < 1 and $p(a, A \cup B) > 0$ so that by Modal Regularity part (ii) $p(a, A \cup B) < 1$. Therefore $0 < p(a, A \cup B) < 1$ and so $\hat{p}(a, A \cup B) = \frac{1}{2}$ as desired.

Step 3. \hat{p} satisfies Negative Expansion. In fact, if $\hat{p}(a, A) < \hat{p}(a, B) < 1$ it must be $\hat{p}(a, A) = 0$ and $\hat{p}(a, B) = \frac{1}{2}$, and therefore p(a, A) = 0. If $p(a, A \cup B) > 0$ then Modal Regularity part (i) is violated. Therefore $p(a, A \cup B) = 0$ and thus $\hat{p}(a, A \cup B) = 0$ as desired.

Step 4. \hat{p} satisfies Contraction Consistency. In fact, suppose that b impacts a in A, that is $\hat{p}(a, A) > \hat{p}(a, A \cup \{b\})$. This means that either $\hat{p}(a, A) \in \{\frac{1}{2}, 1\}$ and $\hat{p}(a, A \cup \{b\}) = 0$, or $\hat{p}(a, A) = 1$ and $\hat{p}(a, A \cup \{b\}) = \frac{1}{2}$. Therefore either $p(a, A) \in (0, 1]$ and $p(a, A \cup \{b\}) = 0$, or p(a, A) = 1 and $p(a, A \cup \{b\}) \in (0, 1)$. Let $B \subset A$. Then by

Lemma 4 either p(a, B) > 0 and $p(a, B \cup \{b\}) = 0$, or p(a, B) = 1 and $p(a, B \cup \{b\}) \in (0, 1)$. This means that either $\hat{p}(a, B) \in \left\{\frac{1}{2}, 1\right\}$ and $\hat{p}(a, B \cup \{b\}) = 0$, or $\hat{p}(a, B) = 1$ and $\hat{p}(a, B \cup \{b\}) = \frac{1}{2}$, so that in either case *b* impacts *a* in *B*.

Observe that the only role performed by Modal Impact Consistency in the proof is that it ensures, together with Modal Regularity, that *p* is binary. So, we have immediately:

Corollary 1 A stochastic choice rule is an mdRUM if and only if it is binary and satisfies Modal Regularity.

It should be easy to check whether a p is binary, and in this sense being binary can serve as a good property for testing the model. However, Modal Impact Consistency is behaviourally a more interesting property because it offers indications on the 'comparative statics' of the model across menus.

We conclude this section with a few remarks.

Remark 2 mdRUMs can accommodate violations of Weak Stochastic Transitivity as well as of Regularity. For example consider *p* given by $p(a, \{a, b\}) = p(a, \{a, b, c\}) = p(b, \{b, c\}) = \frac{2}{3}$, $p(a, \{a, c\}) = \frac{1}{3} = p(b, \{a, b, c\})$. Then *p* violates Weak Stochastic Transitivity but it is an mdRUM generated by the rankings $r_1 = acb$ and $r_2 = bca$ with $\tilde{\alpha}(\{a, b\}) = \tilde{\alpha}(\{a, b, c\}) = \frac{2}{3}$ and $\tilde{\alpha}(\{b, c\}) = \tilde{\alpha}(\{a, c\}) = \frac{1}{3}$.

Remark 3 The ordinal preference information contained in an mdRUM is entirely determined by which choice of alternative, in each menu, is possible, impossible or certain, not by the exact values of the choice probabilities. More precisely suppose that p is an mdRUM generated by $(r_1, r_2, \tilde{\alpha})$ and that p' is an mdRUM generated by $(r'_1, r'_2, \tilde{\alpha}')$. If for all menus A, $p(a, A) > 0 \Leftrightarrow p'(a, A) > 0$, then it must be the case that $r_1 = r'_1$ and $r_2 = r'_2$ or $r_2 = r'_1$ and $r_1 = r'_2$: that is, the rankings coincide up to a relabelling. The exact probabilities of choice in each menu then serve to determine the probabilities with which each ranking is maximised, but not to determine what those rankings look like.

Remark 4 In an mdRUM, while the rankings can have arbitrarily small probability, they always have positive probability. This is crucial to tell apart the case in which the choice probability of an alternative *a* is zero because it is dominated by another alternative in both rankings, from the case in which *a* is dominated say in r_1 and top in r_2 , but r_2 occurs with zero probability in a given menu.

6 Identification

The algorithm we provided in the proof of theorem 3 to construct the rankings would also work for the asymmetric case, but in that case the identification is easier: in fact if *p* is a dRUM generated by some (r_1, r_2, α) , it is easily checked that r_1 and r_2 must satisfy:

$$r_{1}(a) < r_{1}(b) \Leftrightarrow p(a, \{a, b\}) \in \{\alpha, 1\}$$
$$r_{2}(a) < r_{2}(b) \Leftrightarrow p(a, \{a, b\}) \in \{1 - \alpha, 1\}$$

Thus if $\alpha \neq 1 - \alpha$, the rankings can be *uniquely* (up to a relabelling of the rankings) inferred from any dRUM (this definition can be used to provide an alternative, constructive proof of Theorem 1 - such a proof is available from the authors). The uniqueness feature is lost when we drop the requirement of asymmetry. For example consider the dRUM *p* generated by

$$\frac{1}{2} \quad \frac{1}{2}$$

$$a \quad b$$

$$b \quad a$$

$$c \quad d$$

$$d \quad c$$

The same *p* could be alternatively generated by the two different rankings

 $\frac{1}{2} \quad \frac{1}{2}$ $a \quad b$ $b \quad a$ $d \quad c$ $c \quad d$

On the other hand, if these rankings had probabilities α and $1 - \alpha$ with $\alpha \neq \frac{1}{2}$, one could tell the two possibilities apart by observing whether $p(d, \{d, c\}) = \alpha$ or $p(d, \{d, c\}) = 1 - \alpha$.

More in general, the source of non-uniqueness when rankings have symmetric probabilities is the following. Suppose that, given the rankings r_1 and r_2 , the set of alternatives can be partitioned as $X = A_1 \cup ... \cup A_n$ such that if $a \in A_k$, $b \in A_l$ and k < l, then $a \succ_i b$ for i = 1, 2. Now construct two new rankings r'_1 and r'_2 in the following manner. Fix A_k , and for all $a, b \in A_k$ set $a \succ_i b$ if and only if $a \succ_j b$, $i \neq j$, while for all other cases r'_1 and r'_2 coincide with r_1 and r_2 (i.e. r'_1 and r'_2 swap the subrankings within A_k). Then it is clear that the probabilities generated by $(r'_1, r'_2, \frac{1}{2})$ are the same as those generated by $(r_1, r_2, \frac{1}{2})$. Apart from these "swaps", the identification is unique.

7 Concluding remarks

We have argued that the dual RUM constitutes a parsimonious framework to represent the variability of choice in many contexts of interest, both in the fixed state probability version and in the new menu-dependent probability version.

The properties at the core of our analysis are simple and informative about how choices react to the merging, expansion and contraction of menus and thus may provide the theoretical basis for comparative statics analysis and for inferring behaviour in larger menus from behaviour in smaller ones, or vice-versa.

The menu-dependent version of the model, while still quite restrictive (as demonstrated by our characterisation), significantly extends the range of phenomena that the model can encompass, because it does not need to satisfy the full Regularity property but only a 'modal' version of it. As a consequence, for example the attraction and the compromise effects and Luce and Raiffa's 'frog legs' types of anomalies can be described as manu-dependent dual random utility maximisation even if though they are incompatible with any RUM.

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Appendices

A Independence of the axioms - Theorem 1

None of the stochastic choice functions below is a dRUM since there are more than two values in the support.

A.1: Regularity and Constant Expansion, but not Negative Expansion

	$\{a,b,c\}$	$\{a,b\}$	$\{a,c\}$	$\{b,c\}$
а	$\frac{1}{3}$	$\frac{1}{2}$	$\frac{1}{3}$	_
b	0	$\frac{1}{2}$	—	$\frac{1}{3}$
С	$\frac{2}{3}$	—	$\frac{2}{3}$	$\frac{2}{3}$

Table 3:

A.2: Regularity and Negative Expansion, but not Constant Expansion

	$\{a, b, c\}$	$\{a,b\}$	$\{a, c\}$	$\{b,c\}$
а	$\frac{1}{3}$	$\frac{1}{2}$	$\frac{1}{2}$	_
b	$\frac{1}{3}$	$\frac{1}{2}$	_	$\frac{1}{2}$
С	$\frac{1}{3}$	—	$\frac{1}{2}$	$\frac{1}{2}$

Table 4:

A.3: Constant Expansion and Negative Expansion, but not Regularity

Table 5:

B Independence of the axioms - Theorem 2

B.1: Regularity, Weak Constant Expansion, Negative Expansion, but not Contraction

Consistency

See example in Table 2.

B.2: Regularity, Negative Expansion, Contraction Consistency, but not Weak Constant Expansion

See example in Table 4.

B.3: Regularity, Weak Constant Expansion, Contraction Consistency, but not Negative Expansion

	$\{a, b, c\}$	$\{a,b\}$	$\{a, c\}$	$\{b,c\}$
а	$\frac{1}{3}$	0	$\frac{1}{2}$	_
b	$\frac{1}{3}$	$\frac{1}{3}$	_	$\frac{2}{3}$
С	$\frac{1}{3}$	_	$\frac{1}{2}$	$\frac{1}{3}$

Table 6:

B.4: Negative Expansion, Contraction Consistency, Weak Constant Expansion, but not Regularity

	$\{a, b, c\}$	$\{a,b\}$	$\{a, c\}$	$\{b,c\}$
а	1	0	$\frac{1}{2}$	—
b	0	$\frac{1}{3}$	_	$\frac{2}{3}$
С	0	—	$\frac{1}{2}$	$\frac{1}{3}$
		Table 7		

C Independence of the axioms - Theorem 3

C.1: Modal Regularity (i), Modal Regularity (ii), but not Impact Consistency

	$\{a, b, c\}$	$\{a,b\}$	$\{a, c\}$	$\{b,c\}$		
а	$\frac{1}{3}$	$\frac{1}{2}$	$\frac{1}{2}$	—		
b	$\frac{1}{3}$	$\frac{1}{2}$	_	$\frac{1}{2}$		
С	$\frac{1}{3}$	_	$\frac{1}{2}$	$\frac{1}{2}$		
Table 8:						

This stochastic choice function cannot be a mdRUM since in the set $\{a, b, c\}$ there are three alternatives, which cannot be all at the top of two linear orders.

C.2: Modal Regularity (i), Impact Consistency, but not Modal Regularity (ii)

	$\{a, b, c\}$	$\{a,b\}$	$\{a, c\}$	$\{b,c\}$
а	1	$\frac{1}{2}$	1	_
b	0	$\frac{1}{2}$	_	$\frac{1}{2}$
С	0	_	0	$\frac{1}{2}$
		Table 9	:	

This stochastic choice function cannot be a mdRUM since on the one hand $p(a, \{a, b, c\}) =$ 1 implies that *a* is the highest ranked alternative in both linear orders, while on the other $p(b, \{a, b\}) > 0$ would require *b* to rank above *a* in at least one ranking.

C.3: Modal Regularity (ii), Impact Consistency, but not Modal Regularity (i)

	$\{a,b,c\}$	$\{a,b\}$	$\{a, c\}$	$\{b,c\}$		
а	$\frac{1}{2}$	0	1	—		
b	$\frac{1}{2}$	1	—	$\frac{1}{2}$		
С	0	—	0	$\frac{1}{2}$		
Table 10:						

This stochastic choice function cannot be a mdRUM since $p(b, \{a, b\}) = 1$ requires b to be ranked before a in both rankings, while $p(a, \{a, b, c\}) > 0$ would require a to rank above b in at least one ranking.